SOURCE AND ACCURACY STATEMENT FOR THE 1990 PUBLIC USE FILES FROM THE SURVEY OF INCOME AND PROGRAM PARTICIPATION

SOURCE OF DATA

The data were collected in the 1990 panel of the Survey of Income and Program Participation (SIPP). The SIPP universe is the noninstitutionalized resident population living in the United States. The population includes persons living in group quarters, such as dormitories, rooming houses, and religious group dwellings. Crew members of merchant vessels, Armed Forces personnel living in military barracks, and institutionalized persons, such as correctional facility inmates and nursing home residents, were not eligible to be in the survey. Also, United States citizens residing abroad were not eligible to be in the survey. Foreign visitors who work or attend school in this country and their families were eligible; all others were not eligible to be in the survey. With the exceptions noted above, persons who were at least 15 years of age at the time of the interview were eligible to be in the survey.

The 1990 panel of the SIPP sample is located in 230 Primary Sampling Units (PSUs) each consisting of a county or a group of contiguous counties. Within these PSUs, expected clusters of 2 living quarters (LQs) were systematically selected from lists of addresses prepared for the 1980 decennial census to form the bulk of the sample. To account for LQs built within each of the sample areas after the 1980 census, a sample was drawn of permits issued for construction of residential LQs up until shortly before the beginning of the panel. In jurisdictions that do not issue building permits, small land areas were sampled and the LQs within were listed by field personnel and then clusters of 4 LQs were subsampled. In addition, sample LQs were selected from supplemental frames that included LQs identified as missed in the 1980 Census and persons residing in group quarters at the time of the Census.

The 1990 panel differs from the other panels as a result of oversampling for low income. The oversample was constructed by taking a small subsample from the 1989 panel, and combing it with the 1990 panel. Variables such as race, ethnicity, and sex were used for the oversampling since low income data for 1989 panel households were unavailable. The 1989 panel subsample contains all Black Headed Households, all Hispanic Headed Households, all Households with Heads having no spouse present, living with relatives, and a random sample of all the other Household types. The latter random sample was done in an attempt to avoid bias in the sample.

Approximately 28,300 living quarters were designated for the 1990 panel. For Wave 1 of the 1990 panel, interviews were obtained from the occupants of about 21,900 of the 28,300 designated living quarters. Most of the remaining 6,400 living quarters in

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the 1990 panel were found to be vacant, demolished, converted to nonresidential use, or otherwise ineligible for the survey. However, approximately 1,700 of the 6,400 living quarters in the 1990 panel were not interviewed because the occupants refused to be interviewed, could not be found at home, were temporarily absent, or were otherwise unavailable. Thus, occupants of about 93 percent of all eligible living quarters participated in Wave 1 of the Survey for the 1990 panel. Sample loss at Wave 1 of the 1990 Panel was about 7.1% and is expected to increase to roughly 22.0% at the end of Wave 8.

For Waves 2-8, only original sample persons (Those in Wave 1 sample households and interviewed in Wave 1) and persons living with them were eligible to be interviewed. With certain restrictions, original sample persons were to be followed if they moved to a new address. When original sample persons moved without leaving a forwarding address or moved to extremely remote parts of the country and no telephone number was available, additional noninterviews resulted.

Sample households within a given panel are divided into four subsamples of nearly equal size. These subsamples are called rotation groups 1, 2, 3, or 4 and one rotation group is interviewed each month. Each household in the sample was scheduled to be interviewed at 4 month intervals over a period of roughly 2 years beginning in February 1990. The reference period for the questions is the 4-month period preceding the interview month. In general, one cycle of four interviews covering the entire sample, using the same questionnaire, is called a wave.

A unique feature of the SIPP design is overlapping panels. The overlapping design allows panels to be combined and essentially doubles the sample sizes. However, the 1990 panel is designed so that the first three waves do not overlap with other panels. (The 1988 and 1989 panels were prematurely terminated to provide the funding needed to enlarge the 1990 panel and allow oversampling to take place.) After the third wave, the 1990 panel overlaps with the 1991 panel. Selected interviews for the 1990 panel can be combined with interviews from the 1991 panel. Information necessary to do this is included later in this statement.

The public use files include core and supplemental (topical module) data. Core questions are repeated at each interview over the life of the panel. Topical modules include questions which are asked only in certain waves. The 1990 and 1991 panel topical modules are given in tables 1 and 2 respectively.

Tables 3 and 4 indicate the reference months and interview months for the collection of data from each rotation group for the 1990 and 1991 panels respectively. For example, Wave 1 rotation group 2 of the 1990 panel was interviewed in February 1990 and data for

the reference months October 1989 through January 1990 were collected.

Estimation. The estimation procedure used to derive SIPP person weights involved several stages of weight adjustments. Each person received a base weight equal to the inverse of his/her probability of selection. A noninterview adjustment factor was applied to the weight of every occupant of interviewed households to account for households which were eligible for the sample but were not interviewed. (Individual nonresponse within partially interviewed households was treated with imputation. No special adjustment was made for noninterviews in group quarters.) A factor was applied to each interviewed person's weight to account for the SIPP sample areas not having the same population distribution as the strata from which they were selected.

An additional stage of adjustment to persons' weights was performed to reduce the mean square error of the survey estimates by ratio adjusting SIPP sample estimates to monthly Current Population Survey (CPS) estimates[1] of the civilian (and some military) noninstitutional population of the United States by age, race, Spanish origin, sex, type of householder (married, single with relatives, single without relatives), and relationship to householder (spouse or other). The CPS estimates were themselves brought into agreement with estimates from the 1980 decennial census which were adjusted to reflect births, deaths, immigration, emigration, and changes in the Armed Forces since 1980. Also, an adjustment was made so that a husband and wife within the same household were assigned equal weights.

Use of Weights. Users should be forewarned to apply the appropriate weights given on this file before attempting to calculate estimates. The weights vary between units due to the oversampling that took place. If analysis is done for the general population without applying the appropriate weights, the results will be erroneous. Each household and each person within each household on each wave tape has five weights. Four of these weights are reference month specific and therefore can be used only to form reference month estimates. Reference month estimates can be averaged to form estimates of monthly averages over some period of time. For example, using the proper weights, one can estimate the monthly average number of households in a specified income range over November and December 1990. To estimate monthly averages of a given measure (e.g., total, mean) over a number of consecutive months, sum the monthly estimates and divide by the number of months.

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1 These special CPS estimates are slightly different from the published monthly CPS estimates. The differences arise from forcing counts of husbands to agree with counts of wives.

The remaining weight is interview month specific. This weight can be used to form estimates that specifically refer to the interview month (e.g., total persons currently looking for work), as well as estimates referring to the time period including the interview month and all previous months (e.g., total persons who have ever served in the military).

To form an estimate for a particular month, use the reference month weight for the month of interest, summing over all persons or households with the characteristic of interest whose reference period includes the month of interest. Multiply the sum by a factor to account for the number of rotations contributing data for the month. This factor equals four divided by the number of rotations contributing data for the month. For example, December 1989 data is only available from rotations 2, 3, and 4 for Wave 1 of the 1990 panel (See table 3), so a factor of 4/3 (See Table 7) must be applied. To form an estimate for an interview month, use the procedure discussed above using the interview month weight provided on the file.

When estimates for months without four rotations worth of data are constructed from a wave file, factors greater than 1 must be applied. However, when core data from consecutive waves are used together, data from all four rotations may be available, in which case the factors are equal to 1.

These tapes contain no weight for characteristics that involve a person's or household's status over two or more months (e.g., number of households with a 50 percent increase in income between November and December 1990).

Producing Estimates for Census Regions and States. The total estimate for a region is the sum of the state estimates in that region. Using this sample, estimates for individual states are subject to very high variance and are not recommended. The state codes on the file are primarily of use for linking respondent characteristics with appropriate contextual variables (e.g., state-specific welfare criteria) and for tabulating data by user-defined groupings of states.

Producing Estimates for the Metropolitan Population. For Washington, DC and 11 states, metropolitan or non-metropolitan residence is identified (variable H*-METRO). In 34 additional states, where the non-metropolitan population in the sample was small enough to present a disclosure risk, a fraction of the metropolitan sample was recoded to be indistinguishable from non-metropolitan cases (H*-METRO=2). In these states, therefore, the cases coded as metropolitan (H*-METRO=1) represent only a subsample of that population.

In producing state estimates for a metropolitan characteristic, multiply the individual, family, or household weights by the

metropolitan inflation factor for that state, presented in table 5. (This inflation factor compensates for the subsampling of the metropolitan population and is 1.0 for the states with complete identification of the metropolitan population.)

The same procedure applies when creating estimates for particular identified MSA's or CMSA's--apply the factor appropriate to the state. For multi-state MSA's, use the factor appropriate to each state part. For example, to tabulate data for the Washington, DC-MD-VA MSA, apply the Virginia factor of 1.0521 to weights for residents of the Virginia part of the MSA; Maryland and DC residents require no modification to the weights (i.e., their factors equal 1.0).

In producing regional or national estimates of the metropolitan population, it is also necessary to compensate for the fact that no metropolitan subsample is identified within two states (Mississippi and West Virginia) and one state-group (North Dakota - South Dakota - Iowa). Thus, factors in the right-hand column of table 5 should be used for regional and national estimates. The results of regional and national tabulations of the metropolitan population will be biased slightly. However, less than one-half of one percent of the metropolitan population is not represented.

Producing Estimates for the Non-Metropolitan Population. State, regional, and national estimates of the non-metropolitan population cannot be computed directly, except for Washington, DC and the 11 states where the factor for state tabulations in table 5 is 1.0. In all other states, the cases identified as not in the metropolitan subsample (METRO=2) are a mixture of non-metropolitan and metropolitan households. Only an indirect method of estimation is available: first compute an estimate for the total population, then subtract the estimate for the metropolitan population. The results of these tabulations will be slightly biased.

Combined Panel Estimates. Both the 1990 and 1991 panels provide data for October 1990-August 1992. Thus, estimates for these time periods may be obtained by combining the corresponding panels. However, since the Wave 1 questionnaire differs from the subsequent waves' questionnaire, we recommend that estimates not be obtained by combining Wave 1 data of the 1991 panel (collected February - May of 1991) with data of the 1990 panel. In this case, use the estimate obtained from either panel. Additionally, even for other waves, care should be taken when combining data from two panels since questionnaires for the two panels differ somewhat and since the length of time in sample for interviews from the two panels differ.

Combined panel estimates may be obtained either (1) by combining estimates derived separately for the two panels or (2) by first combining data from the two files and then producing an estimate.

1. <u>Combining Separate Estimates</u>

Corresponding estimates from two consecutive year panels can be combined to create joint estimates by using the formula

$$\hat{J} = W\hat{J}_1 + (1-W)\hat{J}_2 \qquad \text{(A)}$$

$\hat{J}$ = joint estimate (total, mean, proportion, etc);

$\hat{J}_1$ = estimate from the earlier panel;

$\hat{J}_2$ = estimate from the later panel;

W = weighting factor of the earlier panel.

To combine the 1990 and 1991 panels use a W value of 0.608 unless one of the panels contributes no information to the estimate. In that case, the panel contributing information receives a factor of 1. The other receives a factor of zero.

2. <u>Combining Data from Separate Files</u>

Start by first creating a file containing the data from the two panel files. Apply the weighting factor, W, to the weight of each person from the earlier panel and apply (1-W) to the weight of each person from the later panel. Estimates can then be produced using the same methodology as used to obtain estimates from a single panel.

<u>Illustration for computing combined panel estimate.</u>

Suppose SIPP estimates for Wave 5 of the 1990 panel show that there were 441,000 households with monthly December income above $6000. Also, suppose SIPP estimates for Wave 2 of the 1991 panel show that there were 435,000 households with monthly December income above $6000. Using formula (A), the joint level estimate is

$$\hat{J} = (0.608)(441{,}000) + (0.392)(435{,}000)$$

$$= 438{,}648$$

ACCURACY OF THE ESTIMATES

SIPP estimates obtained from public use files are based on a sample; they may differ somewhat from the figures that would have been obtained if a complete census had been taken using the same questionnaire, instructions, and enumerators. There are two types of errors possible in an estimate based on a sample survey: nonsampling and sampling. The magnitude of SIPP sampling error can be estimated, but this is not true of nonsampling error. Found below are descriptions of sources of SIPP nonsampling error, followed by discussions of sampling error, its estimation, and its use in data analysis. More detailed discussions of the existence and control of nonsampling errors in the SIPP can be found in the Quality Profile for the Survey of Income and Program Participation, May 1990, by Jabine, assisted by King and Petroni.

Nonsampling Variability. Nonsampling errors can be attributed to many sources, e.g., inability to obtain information about all cases in the sample, definitional difficulties, differences in the interpretation of questions, inability or unwillingness on the part of the respondents to provide correct information, inability to recall information, errors made in collection such as in recording or coding the data, errors made in processing the data, errors made in estimating values for missing data, biases resulting from the differing recall periods caused by the rotation pattern used and failure to represent all units within the universe (undercoverage). Quality control and edit procedures were used to reduce errors made by respondents, coders and interviewers.

Undercoverage in SIPP results from missed living quarters and missed persons within sample households. It is known that undercoverage varies with age, race, and sex. Generally, undercoverage is larger for males than for females and larger for blacks than for nonblacks. Ratio estimation to independent age-race-sex population controls partially corrects for the bias due to survey undercoverage. However, biases exist in the estimates to the extent that persons in missed households or missed persons in interviewed households have different characteristics than the interviewed persons in the same age-race-Spanish origin-sex group. Further, the independent population controls used have not be adjusted for undercoverage.

Some respondents do not respond to some of the questions. Therefore, the overall nonresponse rate for some items such as income and other money related items is higher than the nonresponse rates presented on page 2. The Bureau uses complex techniques to adjust the weights for nonresponse, but the success of these techniques in avoiding bias is unknown.

Comparability With Other Statistics. Caution should be exercised when comparing data from these files with data from other SIPP products or with data from other surveys. The comparability problems are caused by sources such as the seasonal patterns for many characteristics, definitional differences, and different nonsampling errors.

Sampling Variability. Standard errors indicate the magnitude of the sampling variability. They also partially measure the effect of some nonsampling errors in response and enumeration, but do not measure any systematic biases in the data. The standard errors for the most part measure the variations that occurred by chance because a sample rather than the entire population was surveyed.

Confidence Intervals. The sample estimate and its standard error enable one to construct confidence intervals, ranges that would include the average result of all possible samples with a known probability. For example, if all possible samples were selected, each of these being surveyed under essentially the same conditions and using the same sample design, and if an estimate and its standard error were calculated from each sample, then:

1. Approximately 68 percent of the intervals from one standard error below the estimate to one standard error above the estimate would include the average result of all possible samples.

2. Approximately 90 percent of the intervals from 1.6 standard errors below the estimate to 1.6 standard errors above the estimate would include the average result of all possible samples.

3. Approximately 95 percent of the intervals from two standard errors below the estimate to two standard errors above the estimate would include the average result of all possible samples.

The average estimate derived from all possible samples is or is not contained in any particular computed interval. However, for a particular sample, one can say with a specified confidence that the average estimate derived from all possible samples is included in the confidence interval.

Hypothesis Testing. Standard errors may also be used for hypothesis testing, a procedure for distinguishing between population parameters using sample estimates. The most common types of hypotheses tested are 1) the population parameters are identical versus 2) they are different. Tests may be performed at various levels of significance, where a level of significance is the probability of concluding that the parameters are different when, in fact, they are identical.

To perform the most common hypothesis test, compute the difference $X_A - X_B$, where X_A and X_B are sample estimates of the parameters of interest. A later section explains how to derive an estimate of the standard error of the difference $X_A - X_B$. Let that standard error be s_{DIFF}. If $X_A - X_B$ is between -1.6 times s_{DIFF} and +1.6 times s_{DIFF}, no conclusion about the parameters is justified at the 10 percent significance level. If on the other hand, $X_A - X_B$ is smaller than -1.6 times s_{DIFF} or larger than +1.6 times s_{DIFF}, the observed difference is significant at the 10 percent level. In this event, it is commonly accepted practice to say that the parameters are different. Of course, sometimes this conclusion will be wrong. When the parameters are, in fact, the same, there is a 10 percent chance of concluding that they are different.

Note when using small estimates. Because of the large standard errors involved, there is little chance that summary measures would reveal useful information when computed on a smaller base than 200,000. Also, care must be taken in the interpretation of small differences. For instance, in case of a borderline difference, even a small amount of nonsampling error can lead to a wrong decision about the hypotheses, thus distorting a seemingly valid hypothesis test.

Standard Error Parameters and Tables and Their Use. Most SIPP estimates have greater standard errors than those obtained through a simple random sample because clusters of living quarters are sampled. To derive standard errors that would be applicable to a wide variety of estimates and could be prepared at a moderate cost, a number of approximations were required. Estimates with similar standard error behavior were grouped together and two parameters (denoted "a" and "b") were developed to approximate the standard error behavior of each group of estimates. These "a" and "b" parameters are used in estimating standard errors and vary by type of estimate and by subgroup to which the estimate applies. Table 6 provides base "a" and "b" parameters to be used for estimates obtained from core data and for some estimates from topical module data. **These parameters are considered preliminary. Revised parameters are soon to follow.**

The factors provided in table 7 when multiplied by the base parameters of table 6 for a given subgroup and type of estimate give the "a" and "b" parameters for that subgroup and estimate type for the specified reference period. For example, the base "a" and "b" parameters for total number of households are -0.0000664 and 6,043, respectively. For Wave 1 the factor for October 1989 is 4.0000 since only 1 rotation month of data is available. So, the "a" and "b" parameters for total household income in October 1989 based on Wave 1 are -0.0002656 and 24,172, respectively. Also for Wave 1, the factor for the first quarter of 1990 is 1.2222 since 9 rotation months of data are available

(rotations 1 and 4 provide 3 rotations months each, while rotations 2 and 3 provide 1 and 2 rotation months, respectively). So, the "a" and "b" parameters for total number of households in the first quarter of 1990 are -0.0000812 and 7,386, respectively for Wave 1.

The "a" and "b" parameters may be used to calculate the standard error for estimated numbers and percentages. Because the actual standard error behavior was not identical for all estimates within a group, the standard errors computed from these parameters provide an indication of the order of magnitude of the standard error for any specific estimate. Methods for using these parameters for computation of approximate standard errors are given in the following sections.

For those users who wish further simplification, we have also provided preliminary general standard errors in tables 8 through 11 for making estimates with the use of data from all four rotations. Note that these standard errors must be adjusted by a factor (f) from table 6. The standard errors resulting from this simplified approach are less accurate. Methods for using these parameters and tables for computation of standard errors are given in the following sections. Standard errors provided in tables 8 through 11 will change when revised parameters are available.

For the 1990, 1991 combined panel parameters, multiply the parameters in table 6 by the forthcoming appropriate factor from table 15. The factors later provided in table 16 adjust parameters for the number of rotation months available for a given estimate. These factors, when multiplied by the combined panel parameters derived from table 6 for a given subgroup and type of estimate, give the "a" and "b" parameters for that subgroup and estimate type for the specified combined reference period.

For calculating 1990 topical module variances, table 12 is designated to later provide base "a" and "b" parameters. Table 13 also in the near future will provide base "a" and "b" parameters for computing the 1990, 1991 combined panel topical module variances. These parameters will also be provided when revised generalized variance parameters are available.

Procedures for calculating standard errors for the types of estimates most commonly used are described below. Note specifically that these procedures apply only to reference month estimates or averages of reference month estimates. Refer to the section "Use of Weights" for a more detailed discussion of the construction of estimates. Stratum codes and half sample codes are included on the tapes to enable the user to compute the variances directly by methods such as balanced repeated replications (BRR). William G. Cochran provides a list of

references discussing the application of this technique. (See Sampling Techniques, 3rd Ed., New York: John Wiley and Sons, 1977, p. 321.)

Standard Errors of estimated numbers. The approximate standard error, s_x, of an estimated number of persons, households, families, unrelated individuals and so forth, can be obtained in two ways. Both apply when data from all four rotations are used to make the estimate. However, only the second method should be used when less than four rotations of data are available for the estimate. Note that neither method should be applied to dollar values.

It may be obtained by the use of the formula

$$s_x = fs \tag{1}$$

where f is the appropriate "f" factor from table 6, and s is the standard error on the estimate obtained by interpolation from table 8 or 9. Alternatively, s_x may be approximated by the formula

$$s_x = \sqrt{ax^2 + bx} \tag{2}$$

from which the standard errors in tables 8 and 9 were calculated. Here x is the size of the estimate and "a" and "b" are the parameters associated with the particular type of characteristic being estimated. Use of formula 2 will provide more accurate results than the use of formula 1.

Illustration.

Suppose SIPP estimates for Wave 1 of the 1990 panel show that there were 472,000 households with monthly household income above $6,000. The appropriate parameters and factor from table 6 and the appropriate general standard error from table 8 are

a = -0.0000664 b = 6,043 f = 1.00 s = 53,300

Using formula 1, the approximate standard error is

s_x = 53,300

Using formula 2, the approximate standard error is

$$\sqrt{(-0.0000664)(472{,}000)^2 + (6{,}043)(472{,}000)} = 53{,}300$$

Using the standard error based on formula 2, the approximate 90-percent confidence interval as shown by the data is from 387,000 to 557,000. Therefore, a conclusion that the average estimate derived from all possible samples lies within a range computed in this way would be correct for roughly 90% of all samples.

<u>Illustration for computing standard errors for combined panel estimates.</u> Will be provided when combining factors are available.

Standard Error of a Mean. A mean is defined here to be the average quantity of some item (other than persons, families, or households) per person, family or household. For example, it could be the average monthly household income of females age 25 to 34. The standard error of a mean can be approximated by formula 3 below. Because of the approximations used in developing formula 3, an estimate of the standard error of the mean obtained from this formula will generally underestimate the true standard error. The formula used to estimate the standard error of a mean $\bar{x}$ is

$$s_{\bar{x}} = \sqrt{\left(\frac{b}{y}\right)s^2} \qquad (3)$$

where y is the size of the base, s^2 is the estimated population variance of the item and b is the parameter associated with the particular type of item.

The population variance s^2 may be estimated by one of two methods. In both methods we assume x_i is the value of the item for unit i. (Unit may be person, family, or household). To use the first method, the range of values for the item is divided into c intervals. The upper and lower boundaries of interval j are Z_{j-1} and Z_j, respectively. Each unit is placed into one of c groups such that $Z_{j-1} < x_i \leq Z_j$.

The estimated population variance, s^2, is given by the formula:

$$s^2 = \sum_{j=1}^{c} p_j m_j^2 - \bar{x}^2, \tag{4}$$

where p_j is the estimated proportion of units in group j, and $m_j = (Z_{j-1} + Z_j)/2$. The most representative value of the item in group j is assumed to be m_j. If group c is open-ended, i.e., no upper interval boundary exists, then an approximate value for m_c is

$$m_c = \frac{3}{2} Z_{c-1}.$$

The mean, $\bar{x}$ can be obtained using the following formula:

$$\bar{x} = \sum_{j=1}^{c} p_j m_j.$$

In the second method, the estimated population variance is given by

$$s^2 = \frac{\sum_{i=1}^{n} w_i x_i^2}{\sum_{i=1}^{n} w_i} - \bar{x}^2 , \tag{5}$$

where there are n units with the item of interest and w_i is the final weight for unit i. The mean, $\bar{x}$, can be obtained from the formula

$$\bar{x} = \frac{\sum_{i=1}^{n} w_i x_i}{\sum_{i=1}^{n} w_i}.$$

When forming combined estimates using formula (A), s^2, given by formula (4), should be calculated by forming a distribution for each panel. The range of values for the item will be divided into intervals. Combined estimates for each interval can be obtained using formula (A). Formula (4) can be applied to the combined distribution. To calculate $\bar{x}$ and s^2 given by formula (5), replace x_i by Wx_i for x_i from the earlier panel and $(1-W)x_i$ for x_i from the later panel.

Illustration.

Suppose that based on Wave 1 data, the distribution of monthly cash income for persons age 25 to 34 during the month of January 1988 is given in table 14.

Using formula 4 and the mean monthly cash income of $2,530 the approximate population variance, s^2, is

$$s^2 = \left(\frac{1,371}{39,851}\right)(150)^2 + \left(\frac{1,651}{39,851}\right)(450)^2 + \ldots\ldots +$$

$$\left(\frac{1,493}{39,851}\right)(9,000)^2 - (2,530)^2 = 3,159,887.$$

Using formula 3, the appropriate base "b" parameter and factor from table 6, the estimated standard error of a mean $\bar{x}$ is

$$s_{\bar{x}} = \sqrt{\left(\frac{4,890}{39,851,000}\right)(3,159,887)} = \$20$$

Standard error of an aggregate. An aggregate is defined to be the total quantity of an item summed over all the units in a group. The standard error of an aggregate can be approximated using formula 6.

As with the estimate of the standard error of a mean, the estimate of the standard error of an aggregate will generally underestimate the true standard error. Let y be the size of the base, s^2 be the estimated population variance of the item obtained using formula (4) or (5) and b be the parameter associated with the particular type of item. The standard error of an aggregate is:

$$s_x = \sqrt{(b)\ (y)\, s^2} \qquad (6)$$

Standard Errors of Estimated Percentages. The reliability of an estimated percentage, computed using sample data for both numerator and denominator, depends upon both the size of the percentage and the size of the total upon which the percentage is based. Estimated percentages are relatively more reliable than the corresponding estimates of the numerators of the percentages, particularly if the percentages are 50 percent or more, e.g., the percent of people employed is more reliable than the estimated number of people employed. When the numerator and denominator of the percentage have different parameters, use the parameter (and appropriate factor) of the numerator. If proportions are presented instead of percentages, note that the standard error of a proportion is equal to the standard error of the corresponding percentage divided by 100.

There are two types of percentages commonly estimated. The first is the percentage of persons, families or households sharing a particular characteristic such as the percent of persons owning their own home. The second type is the percentage of money or some similar concept held by a particular group of persons or held in a particular form. Examples are the percent of total wealth held by persons with high income and the percent of total income received by persons on welfare.

For the percentage of persons, families, or households, the approximate standard error, $s_{(x,p)}$, of the estimated percentage p can be obtained by the formula

$$s_{(x,p)} = fs \qquad (7)$$

when data from all four rotations are used to estimate p. In this formula, f is the appropriate "f" factor from table 6 and s is the standard error of the estimate from table 10 or 11.

Alternatively, it may be approximated by the formula

$$s_{(x,p)} = \sqrt{\frac{b}{x}\ (p)\ (100-p)} \qquad (8)$$

from which the standard errors in tables 10 and 11 were calculated. Here x is the size of the subclass of social units which is the base of the percentage, p is the percentage (0<p<100), and b is the parameter associated with the characteristic in the numerator. Use of this formula will give more accurate results than use of formula 7 above and should be used when data from less than four rotations are used to estimate p.

Illustration.

Suppose that, in the month of January 1990, 6.7 percent of the 16,812,000 persons in nonfarm households with a mean monthly household cash income of $4,000 to $4,999, were black. Using formula 8 and the "b" parameter of 4,755 from table 6 and a factor of 1 for the month of January 1990 from table 7, the approximate standard error is

$$\sqrt{\frac{4,755}{(16,812,000)}\ (6.7)\ (100-6.7)} = 0.42 percent$$

consequently, the 90 percent confidence interval as shown by these data is from 6.0 to 7.4 percent.

For percentages of money, a more complicated formula is required. A percentage of money will usually be estimated in one of two ways. It may be the ratio of two aggregates:

$$p_I = 100\ (X_A\ /\ X_N)$$

or it may be the ratio of two means with an adjustment for different bases:

$$p_I = 100\ (\hat{P}_A\ \overline{X}_A\ /\ \overline{X}_N)$$

where x_A and x_N are aggregate money figures, $\bar{x}_A$ and $\bar{x}_N$ are mean money figures, and $\hat{p}_A$ is the estimated number in group A divided by the estimated number in group N. In either case, we estimate the standard error as

$$s_I = \sqrt{\left(\frac{\hat{p}_A \bar{x}_A}{\bar{x}_N}\right)^2 \left[\left(\frac{s_p}{\hat{p}_A}\right)^2 + \left(\frac{s_A}{\bar{x}_A}\right)^2 + \left(\frac{s_B}{\bar{x}_N}\right)^2\right]} \quad , \tag{9}$$

where s_p is the standard error of $\hat{p}_A$, s_A is the standard error of $\bar{x}_A$ and s_B is the standard error of $\bar{x}_N$. To calculate s_p, use formula 8. The standard errors of $\bar{x}_N$ and $\bar{x}_A$ may be calculated using formula 3.

It should be noted that there is frequently some correlation between $\hat{p}_A$, $\bar{x}_N$, and $\bar{x}_A$. Depending on the magnitude and sign of the correlations, the standard error will be over or underestimated.

<u>Illustration.</u>

Suppose that in January 1990, 9.8% of the households own rental property, the mean value of rental property is $72,121, the mean value of assets is $78,734, and the corresponding standard errors are 0.25%, $4678, and $2287. In total there are 86,790,000 households. Then, the percent of all household assets held in rental property is

$$= 100 \left((0.098)\frac{72121}{78734}\right) = 9.0\%$$

Using formula (9), the appropriate standard error is

$$s_I = \sqrt{\left(\frac{(0.098)(72121)}{78734}\right)^2 \left[\left(\frac{0.0025}{0.098}\right)^2 + \left(\frac{4678}{72121}\right)^2 + \left(\frac{2287}{78734}\right)^2\right]}$$

$$= 0.007$$

$$= 0.7\%$$

Standard Error of a Difference. The standard error of a difference between two sample estimates is approximately equal to

$$s_{(x-y)} = \sqrt{s_x^2 + s_y^2} \qquad (10)$$

where s_x and s_y are the standard errors of the estimates x and y. The estimates can be numbers, percents, ratios, etc. The above formula assumes that the correlation coefficient between the characteristics estimated by x and y is zero. If the correlation is really positive (negative), then this assumption will tend to cause overestimates (underestimates) of the true standard error.

Illustration.

Suppose that SIPP estimates show the number of persons age 35-44 years with monthly cash income of $4,000 to $4,999 was 3,186,000 in the month of January 1990 and the number of persons age 25-34 years with monthly cash income of $4,000 to $4,999 in the same time period was 2,619,000. Then, using parameters from table 6 and formula 2, the standard errors of these numbers are approximately 124,000 and 112,000, respectively. The difference in sample estimates is 567,000 and, using formula 10, the approximate standard error of the difference is

$$\sqrt{(124{,}000)^2 + (112{,}000)^2} = 167{,}000$$

Suppose that it is desired to test at the 10 percent significance level whether the number of persons with monthly cash income of $4,000 to $4,999 was different for persons age 35-44 years than for persons age 25-34 years. To perform the test, compare the difference of 567,000 to the product 1.6 x 167,000 = 267,200. Since the difference is greater than 1.6 times the standard error of the difference, the data show that the two age groups are significantly different at the 10 percent significance level.

Standard Error of a Median. The median quantity of some item such as income for a given group of persons, families, or households is that quantity such that at least half the group have as much or more and at least half the group have as much or less. The sampling variability of an estimated median depends upon the form of the distribution of the item as well as the size of the group. To calculate standard errors on medians, the procedure described below may be used.

An approximate method for measuring the reliability of an estimated median is to determine a confidence interval about it. (See the section on sampling variability for a general discussion of confidence intervals.) The following procedure may be used to estimate the 68-percent confidence limits and hence the standard error of a median based on sample data.

1. Determine, using either formula 7 or formula 8, the standard error of an estimate of 50 percent of the group;

2. Add to and subtract from 50 percent the standard error determined in step 1;

3. Using the distribution of the item within the group, calculate the quantity of the item such that the percent of the group owning more is equal to the smaller percentage found in step 2. This quantity will be the upper limit for the 68-percent confidence interval. In a similar fashion, calculate the quantity of the item such that the percent of the group owning more is equal to the larger percentage found in step 2. This quantity will be the lower limit for the 68-percent confidence interval;

4. Divide the difference between the two quantities determined in step 3 by two to obtain the standard error of the median.

To perform step 3, it will be necessary to interpolate. Different methods of interpolation may be used. The most common are simple linear interpolation and Pareto interpolation. The appropriateness of the method depends on the form of the distribution around the median. If density is declining in the area, then we recommend Pareto interpolation. If density is fairly constant in the area, then we recommend linear interpolation. Note, however, that Pareto interpolation can never be used if the interval contains zero or negative measures of the item of interest. Interpolation is used as follows. The quantity of the item such that "p" percent own more is

$$X_{pN} = \exp\left[\left(Ln\left(\frac{pN}{N_1}\right) \quad Ln\left(\frac{N_2}{N_1}\right)\right) \; Ln\left(\frac{A_2}{A_1}\right)\right]A_1 \qquad (11)$$

if Pareto Interpolation is indicated and

$$X_{pN} = \left[\frac{PN-N_1}{N_2-N_1} \quad (A_2-A_1) \; + A_1\right] \qquad (12)$$

if linear interpolation is indicated, where N is the size of the group,

A_1 and A_2	are the lower and upper bounds, respectively, of the interval in which X_{pN} falls,
N_1 and N_2	are the estimated number of group members owning more than A_1 and A_2, respectively,
exp	refers to the exponential function and
Ln	refers to the natural logarithm function.

Illustration.

To illustrate the calculations for the sampling error on a median, we return to table 14. The median monthly income for this group is $2,158. The size of the group is 39,851,000.

1. Using formula 8, the standard error of 50 percent on a base of 39,851,000 is about 0.6 percentage points.

2. Following step 2, the two percentages of interest are 49.4 and 50.6.

3. By examining table 14, we see that the percentage 49.4 falls in the income interval from 2000 to 2499. (Since 55.5% receive more than $2,000 per month, the dollar value corresponding to 49.4 must be between $2,000 and $2,500). Thus, A_1 = $2,000, A_2 = $2,500, N_1 = 22,106,000, and N_2 = 16,307,000.

In this case, we decided to use Pareto interpolation. Therefore, the upper bond of a 68% confidence interval for the median is

$$\$2,000 \exp\left[\left(Ln\left(\frac{(.494)(39,851,000)}{22,106,000}\right) \quad Ln\left(\frac{16,307,000}{22,106,000}\right)\right)Ln\left(\frac{2,500}{2,000}\right)\right] = \$2177$$

Also by examining table 14, we see that 50.6 falls in the same income interval. Thus, A_1, A_2, N_1 and N_2 are the same. We also use Pareto interpolation for this case. So the lower bound of a 68% confidence interval for the median is

$$\$2,000 \exp\left[\left(Ln\left(\frac{(.506)(39,851,000)}{22,106,000}\right) \quad Ln\left(\frac{16,307,000}{22,106,000}\right)\right)Ln\left(\frac{2,500}{2,000}\right)\right] = \$2137$$

Thus, the 68-percent confidence interval on the estimated median is from \$2137 to \$2177. An approximate standard error is

$$\frac{\$2177 - \$2137}{2} = \$20$$

Standard Errors of Ratios of Means and Medians. The standard error for a ratio of means or medians is approximated by:

$$s_{\frac{x}{y}} = \sqrt{\left(\frac{x}{y}\right)^2 \left[\left(\frac{s_y}{y}\right)^2 + \left(\frac{s_x}{x}\right)^2\right]} \qquad (13)$$

where x and y are the means or medians, and s_x and s_y are their associated standard errors. Formula 13 assumes that the means are not correlated. If the correlation between the population means estimated by x and y are actually positive (negative), then this procedure will tend to produce overestimates (underestimates) of the true standard error for the ratio of means.

SMD:DButler:sc:DBUT133

Table 1. 1990 Panel Topical Modules

Wave	Topical Module
1	None
2	Recipiency History Employment History Work Disability History Education and Training History Marital History Migration History Fertility History Household Relationships
3	Child Care Arrangements Child Support Agreements Support of Non-household Members Utilization of Health Care Services Functional Limitations & Disability Work Schedule
4	Assets & Liabilities Retirement Expectations & Pensions Plan Coverage Real Estate, Property, and Vehicles
5	Taxes Annual Income and Retirement Accounts School Enrollment and Financing
6	Child Support Agreements Support of Non-household Members Utilization of Health Care Services Functional Limitations & Disability Not in Labor Force Spells
7	Selected Financial Assets Medical Expenses & Work Disability Real Estate Shelter Costs Dependent Care Vehicles
8	Annual Income & Retirement Accounts Taxes School Enrollment & Financing

Table 2. 1991 Panel Topical Modules for Waves 1 through 5[1]

Wave	Topical Module
1	None
2	Welfare History Recipiency History Employment History Work Disability History Education and Training History Marital History Migration History Fertility History Household Relationships
3	Work Schedule Child Care Arrangements Child Support Agreements Support for Non-household Members Functional Limitations & Disability Utilization of Health Care Services
4	Selected Financial Assets Medical Expenses & Work Disability Real Estate Shelter Costs Dependent Care Vehicles
5	Taxes Annual Income and Retirement Accounts School Enrollment and Financing

[1] Topical Modules for waves 6 through 8 are not yet available.

Table 3. Reference Months for Each Interview Month - 1990 Panel

Month of Interview	Wave/ Rotation	4th Quarter (1989) Oct	Nov	Dec	1st Quarter (1990) Jan	Feb	Mar	2nd Quarter (1990) Apr	May	Jun	3rd Quarter (1990) Jul	Aug	Sep	4th Quarter (1990) Oct	Nov	Dec	...	2nd Quarter (1992) Apr	May	Jun	3rd Quarter (1992) Jul	Aug	Sep
Feb 90	1/2	X	X	X	X																		
Mar	1/3		X	X	X	X																	
Apr	1/4			X	X	X	X																
May	1/1				X	X	X	X															
Jun	2/2					X	X	X	X														
Jul	2/3						X	X	X	X													
Aug	2/4							X	X	X	X												
Sept	2/1								X	X	X	X											
Oct	3/2									X	X	X	X										
Nov	3/3										X	X	X	X									
Dec	3/4											X	X	X	X								
.														.	.	.							
.																.	...						
.																	...	.					
Sept 92	8/1																		X	X	X	X	

Table 4. Reference Months for Each Interview Month - 1991 Panel

Month of Interview	Wave/ Rotation	Reference Period: 4th Quarter (1990) Oct	Nov	Dec	1st Quarter (1991) Jan	Feb	Mar	2nd Quarter (1991) Apr	May	Jun	3rd Quarter (1991) Jul	Aug	Sep	4th Quarter (1991) Oct	Nov	Dec	...	2nd Quarter (1993) Apr	May	Jun	3rd Quarter (1993) Jul	Aug	Sep
Feb 91	1/2	X	X	X	X																		
Mar	1/3		X	X	X	X																	
Apr	1/4			X	X	X	X																
May	1/1				X	X	X	X															
Jun	2/2					X	X	X	X														
Jul	2/3						X	X	X	X													
Aug	2/4							X	X	X	X												
Sept	2/1								X	X	X	X											
Oct	3/2									X	X	X	X										
Nov	3/3										X	X	X	X									
Dec	3/4											X	X	X	X								
.														.	.	.							
.																.	...						
.																	...	.					
Sept 93	8/1																		X	X		X	X

Table 5. Metropolitan Subsample Factors to be Applied to Compute National and Subnational Estimates

		Factors for use in State or CMSA (MSA) Tabulations	Factors for use in Regional or National Tabulations
Northeast:	Connecticut	1.0387	1.0387
	Maine	1.2219	1.2219
	Massachusetts	1.0000	1.0000
	New Hampshire	1.2234	1.2234
	New Jersey	1.0000	1.0000
	New York	1.0000	1.0000
	Pennsylvania	1.0096	1.0096
	Rhode Island	1.2506	1.2506
	Vermont	1.2219	1.2219
Midwest:	Illinois	1.0000	1.0110
	Indiana	1.0336	1.0450
	Iowa	---	---
	Kansas	1.2912	1.3055
	Michigan	1.0328	1.0442
	Minnesota	1.0366	1.0480
	Missouri	1.0756	1.0874
	Nebraska	1.6289	1.6468
	North Dakota	---	---
	Ohio	1.0233	1.0346
	South Dakota	---	---
	Wisconsin	1.0188	1.0300
South:	Alabama	1.1574	1.1595
	Arkansas	1.6150	1.6179
	Delaware	1.5593	1.5621
	D.C.	1.0000	1.0018
	Florida	1.0140	1.0158
	Georgia	1.0142	1.0160
	Kentucky	1.2120	1.2142
	Louisana	1.0734	1.0753
	Maryland	1.0000	1.0018
	Mississippi	---	---
	North Carolina	1.0000	1.0018
	Oklahoma	1.0793	1.0812
	South Carolina	1.0185	1.0203
	Tennessee	1.0517	1.0536
	Texas	1.0113	1.0131
	Virginia	1.0521	1.0540
	West Virginia	---	---

- indicates no metropolitan subsample is identified for the state

Table 5 cont'd. Metropolitan Subsample Factors to be Applied to Compute National and Subnational Estimates

		Factors for use in State or CMSA (MSA) Tabulations	Factors for use in Regional or National Tabulations
West:	Alaska	1.4339	1.4339
	Arizona	1.0117	1.0117
	California	1.0000	1.0000
	Colorado	1.1306	1.1306
	Hawaii	1.0000	1.0000
	Idaho	1.4339	1.4339
	Montana	1.4339	1.4339
	Nevada	1.0000	1.0000
	New Mexico	1.0000	1.0000
	Oregon	1.1317	1.1317
	Utah	1.0000	1.0000
	Washington	1.0456	1.0456
	Wyoming	1.4339	1.4339

- indicates no metropolitan subsample is identified for the state

Table 6: SIPP Generalized Variance Parameters for 1990 Panel Public Use File -- Preliminary *

Characteristics	Parameters		
TOTAL PERSONS	a	b	f
16+ Program Participation and Benefits, Poverty (3)			
Both Sexes	-0.0000843	14344	0.90
Male	-0.0001772	14344	
Female	-0.0001604	14344	
16+ Income and Labor Force (5)			
Both Sexes	-0.0000287	4890	0.52
Male	-0.0000605	4890	
Female	-0.0000547	4890	
16+ Pension Plan ** (4)			
Both Sexes	-0.0000525	8956	0.71
Male	-0.0001108	8956	
Female	-0.0001001	8956	
All Others *** (6)			
Both Sexes	-0.0000771	17784	1.00
Male	-0.0001595	17784	
Female	-0.0001493	17784	
WHITE PERSONS			
16+ Program Participation and Benefits, Poverty (3)			
Both Sexes	-0.0000934	15898	0.95
Male	-0.0001964	15898	
Female	-0.0001778	15898	
16+ Income and Labor Force (5)			
Both Sexes	-0.0000318	5420	0.55
Male	-0.0000670	5420	
Female	-0.0000606	5420	
16+ Pension Plan ** (4)			
Both Sexes	-0.0000582	9926	0.75
Male	-0.0001228	9926	
Female	-0.0001110	9926	
All Others *** (6)			
Both Sexes	-0.0000855	19710	1.05
Male	-0.0001768	19710	
Female	-0.0001655	19710	

Table 6 cont'd. SIPP Generalized Variance Parameters for 1990 Panel Public Use File -- Preliminary *

Characteristics	Parameters		
BLACK PERSONS	a	b	f
Poverty (1)			
Both Sexes	-0.0003182	8843	0.71
Male	-0.0006793	8843	
Female	-0.0005987	8843	
All Others *** (2)			
Both Sexes	-0.0001723	4755	0.52
Male	-0.0003704	4755	
Female	-0.0003223	4755	
HISPANIC PERSONS			
Poverty (1)			
Both Sexes	-0.0000609	10374	0.76
Male	-0.0001282	10374	
Female	-0.0001160	10374	
All Others *** (2)			
Both Sexes	-0.0002294	4755	0.52
Male	-0.0004589	4755	
Female	-0.0006727	4755	
HOUSEHOLDS			
Total	-0.0000641	6043	1.00
White	-0.0000823	6698	1.05
Black	-0.0002888	3018	0.71
Hispanics	-0.0005290	3018	0.71

* For cross-tabulations, use the parameters of the characteristic with the smaller number within the parentheses.

** Use the "16+ Pension Plan" parameters for pension plan tabulations of persons 16+ in the labor force. Use the "All Others" parameters for retirement tabulations, 0+ program participation, 0+ benefits, 0+ income, and 0+ labor force tabulations, in addition to any other types of tabulations not specifically covered by another characteristic in this table.

*** Use the "All Others" parameter for any type of tabulation not specifically covered by another characteristic in this table.

Table 7. Factors to be Applied to Table 6 Base Parameters to Obtain Parameters for Various Reference Periods

# of available rotation months[1]	factor
Monthly estimate	
1	4.0000
2	2.0000
3	1.3333
4	1.0000
Quarterly estimate	
6	1.8519
8	1.4074
9	1.2222
10	1.0494
11	1.0370
12	1.0000

[1] The number of available rotation months for a given estimate is the sum of the number of rotations available for each month of the estimate.

Table 8. Standard Errors of Estimated Numbers of Households, Families or Unrelated Persons (Numbers in Thousands)

Size of Estimate	Standard Error[1]	Size of Estimate	Standard Error
200	35	15,000	275
300	43	25,000	331
500	55	30,000	349
750	67	40,000	368
1,000	77	50,000	369
2,000	109	60,000	351
3,000	132	70,000	312
5,000	169	80,000	242
7,500	204	90,000	78
10,000	232		

[1] To account for sample attrition, multiply the standard error of the estimate by 1.04 for estimates which include data from Wave 5 and beyond.

Table 9. Standard Errors of Estimated Numbers of Persons (Numbers in Thousands)

Size of Estimate	Standard Error	Size of Estimate	Standard Error
200	60	50,000	835
300	73	80,000	964
600	103	100,000	1005
1,000	133	130,000	1004
2,000	188	135,000	999
5,000	295	150,000	966
8,000	371	160,000	934
11,000	432	180,000	838
13,000	467	200,000	688
15,000	499	210,000	578
17,000	529	220,000	425
22,000	595	230,000	108
26,000	641		
30,000	681		

[1] To account for sample attrition, multiply the standard error of the estimate by 1.04 for estimates which include data from Wave 5 and beyond.

Table 10. Standard Errors of Estimated Percentages of of Households Families or Unrelated Persons

Base of Estimated Percentage (Thousands)	Estimated Percentages[1]					
	≤ 1 or ≥ 99	2 or 98	5 or 95	10 or 90	25 or 75	50
200	1.73	2.43	3.79	5.20	7.50	8.70
300	1.41	1.99	3.09	4.26	6.20	7.10
500	1.09	1.54	2.40	3.30	4.76	5.50
750	0.89	1.26	1.96	2.69	3.89	4.49
1,000	0.77	1.09	1.69	2.33	3.37	3.89
2,000	0.55	0.77	1.20	1.65	2.38	2.75
3,000	0.45	0.63	0.98	1.35	1.94	2.24
5,000	0.35	0.49	0.76	1.04	1.51	1.74
7,500	0.28	0.40	0.62	0.85	1.23	1.42
10,000	0.24	0.34	0.54	0.74	1.06	1.23
15,000	0.20	0.28	0.44	0.60	0.87	1.00
25,000	0.15	0.22	0.34	0.47	0.67	0.78
30,000	0.14	0.20	0.31	0.43	0.61	0.71
40,000	0.12	0.17	0.27	0.37	0.53	0.61
50,000	0.11	0.15	0.24	0.33	0.48	0.55
60,000	0.10	0.14	0.22	0.30	0.43	0.50
80,000	0.09	0.12	0.19	0.26	0.38	0.43
90,000	0.08	0.11	0.18	0.25	0.35	0.41

[1] To account for sample attrition, multiply the standard error of the estimate by 1.04 for estimates which include data from Wave 5 and beyond.

Table 11. Standard Errors of Estimated Percentages of Persons

Base of Estimated Percentage (Thousands)	Estimated Percentages					
	≤ 1 or ≥ 99	2 or 98	5 or 95	10 or 90	25 or 75	50
200	2.97	4.17	6.50	9.00	12.90	14.90
300	2.42	3.41	5.31	7.30	10.50	12.20
600	1.71	2.41	3.75	5.20	7.50	8.60
1,000	1.33	1.87	2.91	4.00	5.80	6.70
2,000	0.94	1.32	2.06	2.83	4.08	4.71
5,000	0.59	0.83	1.30	1.79	2.58	2.98
8,000	0.47	0.66	1.03	1.41	2.04	2.36
11,000	0.40	0.56	0.88	1.21	1.74	2.01
13,000	0.37	0.52	0.81	1.11	1.60	1.85
17,000	0.32	0.45	0.70	0.97	1.40	1.62
22,000	0.28	0.40	0.62	0.85	1.23	1.42
26,000	0.26	0.37	0.57	0.78	1.13	1.31
30,000	0.24	0.34	0.53	0.73	1.05	1.22
50,000	0.19	0.26	0.41	0.57	0.82	0.94
80,000	0.15	0.21	0.32	0.45	0.65	0.75
100,000	0.13	0.19	0.29	0.40	0.58	0.67
130,000	0.12	0.16	0.25	0.35	0.51	0.58
220,000	0.09	0.13	0.20	0.27	0.39	0.45
230,000	0.09	0.12	0.19	0.26	0.38	0.44

[1] To account for sample attrition, multiply the standard error of the estimate by 1.04 for estimates which include data from Wave 5 and beyond.

Table 12. 1990 Topical Module Generalized Variance Parameters

	a	b
Fertility		
# Females (16+)		
Total	-0.0000403	3,982
White	-0.0000526	4,414
Black	-0.0002431	2,878
Hispanic	-0.0006864	4,851
Births (16+ females)		
Total	-0.0000735	7,261
White	-0.0000960	8,048
Black	-0.0004432	5,248
Hispanic	-0.0012518	8,847
Educational Attainment (16+)		
Wave 2		
Total	-0.0000286	5,424
White	-0.0000372	6,012
Black	-0.0001810	3,921
Hispanic	-0.0002797	3,921
Wave 5		
Total	-0.0000312	5,913
White	-0.0000405	6,553
Black	-0.0001972	4,273
Hispanic	-0.0003048	4,273
Marital Status and Person's Family Characteristics		
Some HH members (16+)		
Total	-0.0000433	8,209
White	-0.0000563	9,098
Black	-0.0002738	5,933
Hispanic	-0.0004232	5,933
All HH members (0+)		
Total	-0.0000405	9,975
White	-0.0000534	11,055
Black	-0.0002374	7,209
Hispanic	-0.0003478	7,209

	a	b
Child Support (16+ females)		
Wave 3		
Total	-0.0000612	6,043
White	-0.0000799	6,698
Black	-0.0003698	4,368
Hispanic	-0.0006180	4,368
Wave 6		
Total	-0.0000667	6,587
White	-0.0000871	7,301
Black	-0.0004021	4,761
Hispanic	-0.0006736	4,761
Support for non-household members (16+)		
Wave 3		
Total	-0.0000319	6,043
White	-0.0000414	6,698
Black	-0.0002016	4,368
Hispanic	-0.0003116	4,368
Wave 6		
Total	-0.0000347	6,587
White	-0.0000452	7,301
Black	-0.0002198	4,761
Hispanic	-0.0003396	4,761
Health and Disability (0+)		
Total	-0.0000318	7,818
White	-0.0000419	8,666
Black	-0.0001861	5,651
Hispanic	-0.0002727	5,651
0-15 Child Care		
Wave 3		
Total	-0.0000867	4,890
White	-0.0001195	5,420
Black	-0.0004064	3,535
Hispanic	-0.0008883	5,956
Wave 6		
Total	-0.0000945	5,331
White	-0.0001303	5,908
Black	-0.0004430	3,853
Hispanic	-0.0009682	6,492

	a	b
Welfare History and AFDC		
Both Sexes 18+		
Total	-0.0000783	14,344
White	-0.0001016	15,898
Black	-0.0005025	10,367
Hispanic	-0.0007784	10,367
Males 18+		
Total	-0.0001638	14,344
White	-0.0002112	15,898
Black	-0.0011083	10,367
Hispanic	-0.0015697	10,367
Females 18+		
Total	-0.0001501	14,344
White	-0.0001959	15,898
Black	-0.0009194	10,367
Hispanic	-0.0015441	10,367

Table 13. Distribution of Monthly Cash Income Among Persons 25 to 34 years old

	Total	under $300	$300 to $599	$600 to $899	$900 to $1,199	$1,200 to $1,499	$1,500 to $1,999	$2,000 to $2,499	$2,500 to $2,999	$3,000 to $3,499	$3,500 to $3,999	$4,000 to $4,999	$5,000 to $5,999	$6,000 and over
Thousands in interval	39,851	1371	1651	2259	2734	3452	6278	5799	4730	3723	2519	2619	1223	1493
Percent with at least as much as lower bound of interval	--	100.0	96.6	92.4	86.7	79.9	71.2	55.5	40.9	29.1	19.7	13.4	6.8	3.7

Table 14. **Factors to be Applied to Base Parameters to Obtain Combined Panel Parameters for Estimates[1] from Various Reference Periods.**

# of avaialble rotation months for 2 panels combined[2]	factor
Monthly Estimate	
2	4.0000
3	3.0000
4	2.0000
5	1.6667
6	1.3333
7	1.1667
8	1.0000
Quarterly Estimates	
12	1.8519
15	1.5631
18	1.2222
19	1.1470
24	1.0000
Annual Estimates	
96	1.0000

[1] Estimates are based on monthly averages.

[2] The number of available rotation months for a given estimate is the sum of the number of rotations available for each month of the estimate for the two panels. There must be at least one rotation month available for each month from each panel for monthly and quarterly estimates.