# Empirical Study of Two Aspects of the Topdown Algorithm Output for Redistricting: <br> Reliability \& Variability <br> (August 5, 2021 Update) 

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# EMPIRICAL STUDY OF TWO ASPECTS OF THE TOPDOWN ALGORITHM OUTPUT FOR REDISTRICTING: RELIABILITY \& VARIABILITY 

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#### Abstract

This two-part study provides an update of empirical results for ongoing research and development that were reported in [6]. In this update, data output from the same version of the TopDown Algorithm that will produce the " 2020 Census Redistricting Data (Public Law 94-171) Summary File" are reported in the tables and figures. Except for wording changes due to changes in the data output, the wording throughout is the same as in [6]. The TopDown Algorithm (TDA) [1] is being used to protect the confidentiality of respondent data collected during the 2020 Census. Following the 2010 Census, the swapping methodology (SWA) [7] was applied to respondent data to protect confidentiality.

In Part I, we propose an empirically based solution to the question: "What is the minimum TOTAL population of a district to have reliable characteristics of various demographic groups?" To answer this question, we use data treated by the 2020 Census redistricting data production settings version ( $\epsilon=17.14$, for the person file) of the $T D A$ for all block groups (proxy for districts) in the United States. We also consider "places and minor civil divisions (MCDs)" as proxies for districts. Empirical results suggest a minimum TOTAL that is between 450 and 499 people in a block group provides reliable characteristics of various demographic groups in a block group based on the TDA. Similarly, a minimum TOTAL that is between 200 and 249 is observed to provide reliable characteristics for places and MCDs. No Congressional or state legislative district failed our test for reliability. It is important to keep in mind that these results are comparisons to the swapped 2010 Census data. They do not evaluate the reliability relative to the actual enumeration in 2010 because the 2010 redistricting data contained statistical uncertainty due to swapping.

Part II is an update of our results reported in [6] where $\epsilon=10.3$ with the difference being that this study uses $\epsilon=17.14$. The objective here is to assess the variability of data results from application of the 2020 Census redistricting data production settings version $T D A$ to the 2010 Census Edited File ( 2010 CEF) for Rhode Island and for three additional jurisdictions. Our approach has two parts: (1) to report observations on variability of results among 25 runs of the $T D A$ and (2) to report observations on variability between the results among the 25 runs of the TDA and the published 2010 Census Public Law 94-171 data. We observe that variability in data results from the TDA increases as we consider smaller pieces of geography and population. Variability with the 2020 Census redistricting data production settings version of the $T D A(\epsilon=17.14)$ tends to be less than what we reported in [6] with the 2021-04-28 version where $\epsilon=10.3$.


[^0]COMMENT: Throughout Parts I and II, we compare TDA counts with published corresponding SWA counts from 2010 rather than with the "as enumerated" 2010 counts, i.e., counts in the 2010 Census Edited File (CEF). For a clean comparison, it would be better to compare TDA counts with the corresponding CEF counts. However, we share a few thoughts that provide some support for the path we take, to use the $S W A$ counts as a reference for assessing the TDA counts. First, the $S W A$ counts from 2010 are official; they have been used widely by the public for ten years; and we assume that they have generally been accepted as credible. The public is familiar with the $S W A$ counts. In this spirit, we see some value in comparing TDA counts with $S W A$ counts. This permits the public the opportunity to compare relatively easily and to possibly reproduce most of our results. This would be impossible if we had used the CEF counts, which are confidential. A primary objective in Part I is to convey a new data-based concept - "what we mean by declaring TDA counts reliable". We don't really need the CEF counts to discuss this concept. It should be noted that the SWA TOTAL counts and the corresponding CEF TOTAL counts at the block level were the same in 2010. The same is true for TOTAL18 counts for the 18 years and over population at the block level. It should also be noted that the "tuning" of the TDA makes use of the CEF counts rather than the $S W A$ counts, and we understand that results are similar to what we share, especially with regard to the main question on reliability in Part I. Furthermore, had we used CEF counts, additional Disclosure Review Board clearance would have slowed the speed in sharing our study results.

## TECHNICAL SUMMARY

The Census Bureau Data Stewardship Executive Policy Committee (DSEP) approved production settings for the 2020 Census Redistricting Data (Public Law 94-171) Summary File (hereafter " 2020 Census redistricting data production settings") version of the TopDown Algorithm (TDA) [1] that will be applied to the 2020 Census Edited File (CEF), and the results will be used by jurisdictions in devising redistricting plans for selecting officials ranging from Members of the U.S. House of Representatives to local school boards. We also assume the results will be used for the analysis of such plans for compliance with Federal voting rights laws, including Section 2 of the Voting Rights Act of 1965, 52 U.S.C. 10301.

In Part I of this limited study, we attempt to take a closer look at reliability of characteristics of demographic groups inside smaller districts. For convenience, we consider "Census Block Groups, Minor Civil Divisions (MCDs), and Census Places" as proxies for smaller districts and seek to gain more insights regarding the following question:
"What is the minimum TOTAL (ideal ${ }^{a}$ ) population of a district to have reliable characteristics of various demographic groups?"

For each of the 217,740 block groups and 21,591 MCDs and places in the United States, we desire to compare the closeness between the following two sets of population counts: (a) published $S W A$ counts for twenty demographic groups based on the application of a Swapping Algorithm ( $S W A$ ) to the 2010 CEF and (b) the corresponding TDA counts for the same twenty demographic groups based on application of the 2020 Census redistricting data production settings version of the TDA $(\epsilon=17.14)$ to the 2010 CEF. Our comparisons are facilitated by a measure called the difference of ratios $D R$ (see Section I.1). We analyze data for block groups, MCDs, and places as proxies for districts to make reliability statements about TDA output. We also analyze all Congressional and state legislative districts. For block groups, MCDs, places, and legislative districts:

## The Key Empirical Message on Reliability

"for any block group with a TOTAL count between 450 and 499 people, and for MCDs and places between 200 and 249, the difference between the TDA ratio of the largest demographic group ( $L D G$ ) and the corresponding $S W A$ ratio for the $L D G$ is less than or equal to 5 percentage points at least $95 \%$ of the time". No Congressional or state legislative district fails this test; that is, for these districts, the 5 percentage point criterion holds $100 \%$ of the time.

Part II of this study provides empirical results for ongoing research and development and provides an update of the data and results presented in [6] where $\epsilon=10.3$; throughout this updated study, $\epsilon=17.14$. The objective of this part of our study is to assess the variability of data results from application of the 2020 Census redistricting data production settings version of the TDA to the 2010 Census Edited File ( 2010 CEF) for Rhode Island and for three additional jurisdictions. Because there has been more development of the TDA, a larger $\epsilon$, and additional focus on how to allocate this $\epsilon$, we tend to see less variability throughout.

Part II of our study has two components: (1) report variability among 25 runs and (2) report variability of the 25 runs relative to the official published results from the 2010 Census (i.e., the 2010 Census Redistricting Data (Public Law 94-171) Summary File).

The first component of our study is a follow-up to earlier analyses [5, 6] for Rhode Island. For each of the given redistricting plans we studied for Rhode Island, we observe that counts and percentages put in place from swapping being applied to the 2010 CEF have very similar counts and percentages after the TDA is applied to the same 2010 CEF.

In the second component of our study, we repeat our analyses for three specific jurisdictions provided by the U.S. Department of Justice (DOJ). Our observations for these three smaller geographies and populations show similarities between swapping ( $S W A$ ) and TDA results.

The key data analyses are presented
(i) in Tables 7, 8, 9, 10, 11, and 12 that contain $S W A$ counts and percentages publicly released following the 2010 Census and corresponding released TDA counts and percentages; and
(ii) in Tables $7 \mathrm{~V}, 8 \mathrm{~V}, 9 \mathrm{~V}, 10 \mathrm{~V}, 11 \mathrm{~V}$, and 12 V that contain measures of relative variability for the TDA as described in Section II. 8 (APPENDIX B contains an illustration of the computations).

## The Key Empirical Message on Variability

The two measures $A V E R V(\cdot)$ and $M E D R V(\cdot)$, defined in Section II.7, summarize the key single empirical message for Part II of this study $(\epsilon=17.14)$. As we reported in $[5,6]$, relative variability in the TDA increases as we consider smaller pieces of geography and population. To see this empirical evidence, sequentially observe the values for $A V E R V(\cdot)$ and $M E D R V(\cdot)$ on the last two rows of Tables $7 \mathrm{~V} ; 8 \mathrm{~V} ; 9 \mathrm{~V} ; 10 \mathrm{~V} ; 11 \mathrm{~V}$; and 12 V ; also see Figure 1. At a high level, Figure 2 tends to show less relative variability using the 2020 Census redistricting data production settings version of the TDA than the 2021-04-28 version.

## I.1. INTRODUCTION

Our earlier empirical study [6] assessed the variability of data results from application of the 2021-04-28 version of the TopDown Algorithm (TDA) to the 2010 Census Edited File (2010 CEF) for disclosure avoidance and confidentiality protection. It documented that it is the smaller geographic districts with smaller ideal ${ }^{a}$ populations where we observed more variability among twenty-five different runs of the TDA. Indeed, it is the block level where redistricting takes place, where local people have some sense of "ground truth", and where some field checking seems possible to assess the reliability of $T D A$ output. In Part I of this study, we attempt to take a closer look at variability for smaller districts (a level closer to the block level) and the reliability of counts of various demographic groups in these smaller districts based on the TDA. As a proxy for smaller districts, we consider Census block groups, Minor Civil Divisions (MCDs) and Census places and seek insights for the following question:
"What is the minimum TOTAL (ideal ${ }^{a}$ ) population of a district to have reliable characteristics of various demographic groups?"
(A block group is a cluster of blocks and generally contains between 600 and 3,000 people. MCDs and places vary in size, but approximately half have population less than or equal to 2,100 people.)

For each of the 217,740 block groups in the United States and for each of the 21,591 MCDs and places, we desire to compare closeness between the following two sets of population counts: (a) published $S W A$ counts for twenty demographic groups based on the application of a Swapping Algorithm (SWA) to the 2010 CEF and (b) the corresponding TDA counts for the same twenty demographic groups based on application of the DSEP-approved production settings for the 2020 Census Redistricting Data (Public Law 94-171) Summary File (hereafter " 2020 Census redistricting data production settings") version of the TDA to the 2010 CEF. Our comparisons are facilitated by the difference of ratios $(D R)$.

Definition 1: Let $C_{S W A}(g)$ and $C_{T D A}(g)$ be two competing counts of the demographic group $g$ associated with a block group (more generally, geographic district) whose total population counts are $C_{S W A}$ and $C_{T D A}$, respectively. The difference of ratios is the absolute value of the difference between the $S W A$ ratio $\frac{C_{S W A}(g)}{C_{S W A}}$ and the TDA ratio $\frac{C_{T D A}(g)}{C_{T D A}}$, given by:

$$
\begin{equation*}
D R_{g}=\left|\frac{C_{S W A}(g)}{C_{S W A}}-\frac{C_{T D A}(g)}{C_{T D A}}\right| . \tag{1}
\end{equation*}
$$

Small values of the difference of ratios $D R_{g}$ imply that the ratios for a group $g$ due to $S W A$ and TDA in the block group, MCD, or place are close.

Definition 2: When $D R_{g}$ is sufficiently small while comparing a $C_{S W A}(g)$ count and corresponding $C_{T D A}(g)$ count for a demographic group $g$ associated with a given block group, MCD or place, we say that the $C_{T D A}(g)$ count (or ratio) provides a reliable characteristic for the block group, MCD, or place.

[^1]
## I.2. ILLUSTRATION OF COMPUTATIONS FOR TWO BLOCK GROUPS

For a block group in Maryland, Table 1a provides differences of ratios for twenty demographic groups as used in the past for redistricting related analyses [6]. For definition of each demographic group, see APPENDIX A. For the demographic group $g=\operatorname{ASIANNH18,~} C_{S W A}(g)=142$ and $C_{T D A}(g)=146$ with difference of ratios $D R_{g}=0.0003$. That is, the difference between the two ratios for demographic group $g$ is 0.03 percentage points for this block group. (Note using Appendix A that $C_{S W A}(g)=142(=130+12)$ where 130 is the count for all individuals 18 years of age or older who chose Asian singly and chose Not Hispanic; and 12 is the count for all individuals 18 years of age or older who chose Asian in combination with White and chose Not Hispanic.)

Note: When the counts being compared are for individuals of all ages for a block group, we take $C_{S W A}=$ TOTAL count using $S W A$ and $C_{T D A}=$ TOTAL count using TDA; when the counts being compared for individuals 18 years and older for a block group, we take $C_{S W A}=$ TOTAL18 count using $S W A$ and $C_{T D A}=$ TOTAL18 count using TDA.

Table 1a: Block Group 240317044041 ( 564 HUs) Characteristics

| Demographic Group $(g)^{b}$ | $C_{S W A}(\mathrm{~g})$ | $C_{T D A}(g)$ | $D R_{g}=$ | $\left\|\frac{C_{S W A}(g)}{C_{S W A}}-\frac{C_{T D A}(g)}{C_{T D A}}\right\|$ |
| :---: | :---: | :---: | :---: | :---: |
| TOTAL | 1,560 | 1,598 |  | c |
| TOTAL18 | 1,198 | 1,229 |  | c |
| TOTALHISP | 133 | 141 |  | $\left\|\frac{133}{1,560}-\frac{141}{1,598}\right\|=0.0030$ |
| TOTALNH | 1,427 | 1,457 |  | $\left\|\frac{1,427}{1,560}-\frac{1,457}{1,598}\right\|=0.0030$ |
| WHITENH | 1,169 | 1,178 |  | $\left\|\frac{1,169}{1,560}-\frac{1,178}{1,598}\right\|=0.0122$ |
| BLACKNH | 36 | 54 |  | $\left\|\frac{36}{1,560}-\frac{54}{1,598}\right\|=0.0107$ |
| AIANNH | 10 | 8 |  | $\left\|\frac{10}{1.560}-\frac{8}{1.598}\right\|=0.0014$ |
| ASIANNH | 187 | 189 |  | $\left\|\frac{187}{1,560}-\frac{189}{1,598}\right\|=0.0016$ |
| HPINH | 5 | 2 |  | $\left\|\frac{5}{1,560}-\frac{2}{1,598}\right\|=0.0020$ |
| OTHERNH | 11 | 12 |  | $\left\|\frac{11}{1,560}-\frac{12}{1,598}\right\|=0.0005$ |
| MLTMNNH | 9 | 14 |  | $\left\|\frac{1,9}{1,560}-\frac{14}{1,598}\right\|=0.0030$ |
| HISP18 | 93 | 95 |  | $\left\|\frac{93}{1,198}-\frac{95}{1,229}\right\|=0.0003$ |
| NONHISP18 | 1,105 | 1,134 |  | $\left\|\frac{1,105}{1,198}-\frac{1,134}{1,292}\right\|=0.0003$ |
| WHITENH18 | 914 | 923 |  | $\left\|\frac{914}{1,198}-\frac{923}{1,229}\right\|=0.0119$ |
| BLACKNH18 | 29 | 42 |  | $\left\|\frac{29}{1,198}-\frac{42}{1,299}\right\|=0.0100$ |
| AIANNH18 | 8 | 8 |  | $\left\|\frac{8,8}{1,198}-\frac{8}{1,299}\right\|=0.0002$ |
| ASIANNH18 | 142 | 146 |  | $\left\|\frac{142}{1,198}-\frac{146}{1,229}\right\|=0.0003$ |
| HPINH18 | 2 | 2 |  | $\left\|\frac{2}{1,198}-\frac{2}{1,229}\right\|=0.0000$ |
| OTHERNH18 | 6 | 4 |  | $\left\|\frac{6}{1,198}-\frac{4}{1,229}\right\|=0.0018$ |
| MLTMNNH18 | 4 | 9 |  | $\left\|\frac{4}{1,198}-\frac{9}{1,229}\right\|=0.0040$ |

${ }^{b}$ For definitions of the demographic groups, see APPENDIX A.
${ }^{c}$ Because $D R_{g}=0.0000$ when $g=$ TOTAL or $g=$ TOTAL18 in Tables 1a, 1b, and 2, we leave the entries for $D R_{g}$ empty. To see comparisons in these cases, one could take $\left|C_{S W A}(g)-C_{T D A}(g)\right| / C_{S W A}$ which is a special case of $D R_{g}$. (A similar approach could be taken for TOTAL18.)

Thus from Table 1a and for the difference of ratios for demographic group $g=$ TOTALNH, $D R_{g}=0.0030$; the difference between the two ratios is $0.0030 \times 100 \%=0.30$ percentage points.

Table 1b provides similar characteristics of demographic groups for a block group in Washington D.C. From Table 1b and for the difference of ratios for demographic group $g=$ TOTALNH, the difference between the ratios is $0.0017 \times 100 \%=0.17$ percentage points.

Table 1b: Block Group 110010047012 (1,709 HUs) Characteristics
( $C_{T D A}(g)$ counts result from 2020 Census Redistricting Data Production Settings ( $\epsilon=17.14$ for persons) version of TDA.)

|  |  |  |  |
| :--- | ---: | ---: | ---: |
| Demographic Group $(g)$ | $C_{S W A}(g)$ | $C_{T D A}(g)$ | $D R_{g}=\left\|\frac{C_{S W A}(g)}{C_{S W A}}-\frac{C_{T D A}(g)}{C_{T D A}}\right\|$ |
|  |  |  |  |
|  |  |  | $c$ |
| TOTAL | 2,875 | 2,868 | $c$ |
| TOTAL18 | 2,261 | 2,244 |  |
| TOTALHISP |  |  | 0.0017 |
| TOTALNH | 92 | 87 | 0.0017 |
| WHITENH | 2,783 | 2,781 | 0.0020 |
| BLACKNH | 541 | 534 | 0.0021 |
| AIANNH | 1,686 | 1,688 | 0.0007 |
| ASIANNH | 12 | 10 | 0.0036 |
| HPINH | 515 | 524 | 0.0003 |
| OTHERNH | 1 | 0 | 0.0000 |
| MLTMNNH | 3 | 3 | 0.0010 |
|  | 25 | 22 |  |
| HISP18 |  |  | 0.0051 |
| NONHISP18 | 86 | 74 | 0.0051 |
| WHITENH18 | 2,175 | 2,170 | 0.0004 |
| BLACKNH18 | 529 | 526 | 0.0034 |
| AIANNH18 | 1,151 | 1,150 | 0.0009 |
| ASIANNH18 | 12 | 10 | 0.0020 |
| HPINH18 | 460 | 461 | 0.0004 |
| OTHERNH18 | 1 | 0 | 0.0004 |
| MLTMNNH18 | 3 | 2 | 0.0010 |
|  | 19 | 21 |  |

## I.3. CHARACTERISTICS OF TWELVE MORE BLOCK GROUPS

We extend our overview of block groups beyond those in Tables 1a and 1b by considering counts for the demographic groups for block groups with TOTAL that span from 82 (this block group is actually the complete Loving County, Texas) to 37,452 (this block group is the largest block group in population in the United States). Table 2 presents the characteristics we observe. Our analyses focus more on the larger demographic groups within each block group because they may play a larger role when thinking about reliable characteristics of actual districts. We highlight the counts and $D R_{g}$ 's for the following demographic groups \{TOTAL, TOTAL18\} and for some of the demographic groups \{TOTALHISP, WHITENH, BLACKNH, AIANNH, ASIANNH, HPINH\}. The superscripts ${ }^{1},{ }^{2}$, and ${ }^{3}$ represent, in order, the three largest demographic groups among TOTALHISP, WHITENH, BLACKNH, AIANNH, ASIANNH, and HPINH (based on $C_{T D A}(g)$ counts) for the block group. Clearly, as the count for the TOTAL demographic group increases across the twelve block groups in Table 2, corresponding values of highlighted $D R_{g}$ values tend to decrease.

## Motivating Example for Reliable Characteristics

Assume we stratify or partition the 12 block groups in Table 2 into 4 strata; the first three, then the next 3 , the next three, and finally the last three with the following $D R_{g}$ values for each stratum where $g$ is the largest demographic group: $\{0.0494,0.0239,0.0032\} ;\{0.0127,0.0024,0.0056\}$; $\{0.0012,0.0001,0.0010\}$; and $\{0.0004,0.0000,0.0000\}$. Assume the $T D A$ count is considered a reliable characteristic for the largest demographic group if its $D R_{g} \leq 0.0050$. One of the block groups in stratum 1 would be reliable; 1 out of 3 ( 0.3333 ) of the block groups in stratum 2 would be reliable; all 3 (1.0000) of the block groups in stratum 3 would be reliable; and finally, again all $3(1.0000)$ of the block groups in stratum 4 would be reliable. We build on this in Section I.4.

Table 2: Characteristics of Twelve Block Groups
( $C_{T D A}(g)$ counts result from 2020 Census Redistricting Data Production Settings ( $\epsilon=17.14$ for persons) version of TDA.)

| Demographic Group | Block Group $483019501001(\mathrm{TX})^{d}$ |  |  | Block Group 010599729001 (AL) |  |  | Block Group 010059507002 (AL) |  |  | $\begin{aligned} & \text { Block Group } \\ & 040030008001 \text { (AZ) } \end{aligned}$ |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| (g) | $C_{S W A}$ | $C_{T D A}$ | $D R_{g}$ | $C_{S W A}$ | $C_{T D A}$ | $D R_{g}$ | $C_{S W A}$ | $C_{T D A}$ | $D R_{g}$ | $C_{S W A}$ | $C_{T D A}$ | $D R_{g}$ |
| TOTAL | 82 | 85 | c | 500 | 514 | c | 1,000 | 980 | c | 1,500 | 1,543 | c |
| TOTAL18 | 73 | 67 | c | 386 | 389 | c | 745 | 733 | $c$ | 1,035 | 1,047 | c |
| TOTALHISP | 18 | $13^{2}$ | 0.0666 | 18 | $16^{2}$ | 0.0049 | 30 | $24^{3}$ | 0.0055 | 1,237 | 1,292 ${ }^{1}$ | 0.0127 |
| TOTALNH | 64 | 72 | 0.0666 | 482 | 498 | 0.0049 | 970 | 956 | 0.0055 | 263 | 251 | 0.0127 |
| WHITENH | 60 | $58{ }^{1}$ | 0.0494 | 455 | $480^{1}$ | 0.0239 | 306 | $301{ }^{2}$ | 0.0011 | 235 | $215{ }^{2}$ | 0.0173 |
| BLACKNH | 0 | 0 | 0.0000 | 7 | $9^{3}$ | 0.0035 | 659 | $649{ }^{1}$ | 0.0032 | 10 | 9 | 0.0008 |
| AIANNH | 4 | 2 | 0.0253 | 6 | 7 | 0.0016 | 4 | 1 | 0.0030 | 0 | 2 | 0.0013 |
| ASIANNH | 0 | $2^{3}$ | 0.0235 | 11 | 0 | 0.0220 | 0 | 2 | 0.0020 | 18 | $20^{3}$ | 0.0010 |
| HPINH | 0 | 2 | 0.0235 | 0 | 0 | 0.0000 | 0 | 0 | 0.0000 | 0 | 1 | 0.0006 |
| OTHERNH | 0 | 7 | 0.0824 | 1 | 1 | 0.0001 | 0 | 0 | 0.0000 | 0 | 2 | 0.0013 |
| MLTMNNH | 0 | 1 | 0.0118 | 2 | 1 | 0.0021 | 1 | 3 | 0.0021 | 0 | 2 | 0.0013 |
| HISP18 | 14 | 3 | 0.1470 | 10 | 11 | 0.0024 | 21 | 22 | 0.0018 | 807 | 818 | 0.0016 |
| NONHISP18 | 59 | 64 | 0.1470 | 376 | 378 | 0.0024 | 724 | 711 | 0.0018 | 228 | 229 | 0.0016 |
| WHITENH18 | 55 | 53 | 0.0376 | 354 | 366 | 0.0238 | 255 | 250 | 0.0012 | 203 | 198 | 0.0070 |
| BLACKNH18 | 0 | 0 | 0.0000 | 6 | 5 | 0.0027 | 464 | 458 | 0.0020 | 9 | 9 | 0.0001 |
| AIANNH18 | 4 | 2 | 0.0249 | 5 | 7 | 0.0050 | 4 | 1 | 0.0040 | 0 | 2 | 0.0019 |
| ASIANNH18 | 0 | 0 | 0.0000 | 9 | 0 | 0.0233 | 0 | 0 | 0.0000 | 16 | 17 | 0.0008 |
| HPINH18 | 0 | 2 | 0.0299 | 0 | 0 | 0.0000 | 0 | 0 | 0.0000 | 0 | 1 | 0.0010 |
| OTHERNH18 | 0 | 6 | 0.0896 | 0 | 0 | 0.0000 | 0 | 0 | 0.0000 | 0 | 1 | 0.0010 |
| MLTMNNH18 | 0 | 1 | 0.0149 | 2 | 0 | 0.0052 | 1 | 2 | 0.0014 | 0 | 1 | 0.0010 |

${ }^{d}$ This block group is all of Loving County, Texas.

Table 2: Characteristics of Twelve Block Groups (continued)
( $C_{T D A}(g)$ counts result from 2020 Census Redistricting Data Production Settings ( $\epsilon=17.14$ for persons) version of TDA.)

| Demographic Group (g) | $\begin{gathered} \text { Block Group } \\ 040030017032 \text { (AZ) } \end{gathered}$ |  |  | $\begin{gathered} \text { Block Group } \\ 051430110011 \text { (AR) } \end{gathered}$ |  |  | $\begin{aligned} & \text { Block Group } \\ & 120210112023 \text { (FL) } \end{aligned}$ |  |  | $\begin{aligned} & \text { Block Group } \\ & 131350505461 \text { (GA) } \end{aligned}$ |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $C_{S W A}$ | $C_{T D A}$ | $D R_{g}$ | $C_{S W A}$ | $C_{T D A}$ | $D R_{g}$ | $C_{S W A}$ | $C_{T D A}$ | $D R_{g}$ | $C_{\text {SWA }}$ | $C_{T D A}$ | $D R_{g}$ |
| TOTAL | 2,000 | 1,982 | c | 3,000 | 2,976 | c | 5,001 | 4,980 | c | 10,000 | 10,042 | c |
| TOTAL18 | 1,562 | 1,549 | c | 2,153 | 2,125 | c | 3,689 | 3,689 | c | 6,704 | 6,716 | c |
| TOTALHISP | 349 | $323{ }^{2}$ | 0.0115 | 224 | $221{ }^{2}$ | 0.0004 | 1,770 | 1,773 ${ }^{2}$ | 0.0021 | 1,291 | 1,298 ${ }^{3}$ | 0.0002 |
| TOTALNH | 1,651 | 1,659 | 0.0115 | 2,776 | 2,755 | 0.0004 | 3,231 | 3,207 | 0.0021 | 8,709 | 8,744 | 0.0002 |
| WHITENH | 1,308 | 1,301 ${ }^{1}$ | 0.0024 | 2,580 | 2,576 ${ }^{1}$ | 0.0056 | 2,891 | 2,873 ${ }^{1}$ | 0.0012 | 3,565 | 3,569 ${ }^{2}$ | 0.0011 |
| BLACKNH | 181 | $171{ }^{3}$ | 0.0042 | 87 | $80^{3}$ | 0.0021 | 235 | $223{ }^{3}$ | 0.0022 | 4,475 | 4,495 ${ }^{1}$ | 0.0001 |
| AIANNH | 25 | 24 | 0.0004 | 65 | 64 | 0.0002 | 18 | 13 | 0.0010 | 30 | 32 | 0.0002 |
| ASIANNH | 106 | 120 | 0.0075 | 32 | 31 | 0.0003 | 59 | 65 | 0.0013 | 473 | 472 | 0.0003 |
| HPINH | 10 | 10 | 0.0000 | 1 | 0 | 0.0003 | 8 | 3 | 0.0010 | 2 | 2 | 0.0000 |
| OTHERNH | 3 | 9 | 0.0030 | 4 | 2 | 0.0007 | 7 | 17 | 0.0020 | 79 | 90 | 0.0011 |
| MLTMNNH | 18 | 24 | 0.0031 | 7 | 2 | 0.0017 | 13 | 13 | 0.0000 | 85 | 84 | 0.0001 |
| HISP18 | 236 | 220 | 0.0091 | 110 | 111 | 0.0011 | 1,193 | 1,201 | 0.0022 | 783 | 787 | 0.0004 |
| NONHISP18 | 1,326 | 1,329 | 0.0091 | 2,043 | 2,014 | 0.0011 | 2,496 | 2,488 | 0.0022 | 5,921 | 5,929 | 0.0004 |
| WHITENH18 | 1,089 | 1,080 | 0.0000 | 1,931 | 1,913 | 0.0033 | 2,267 | 2,259 | 0.0022 | 2,630 | 2,628 | 0.0010 |
| BLACKNH18 | 129 | 122 | 0.0038 | 40 | 37 | 0.0012 | 149 | 144 | 0.0014 | 2,868 | 2,876 | 0.0004 |
| AIANNH18 | 20 | 18 | 0.0012 | 41 | 41 | 0.0003 | 14 | 6 | 0.0022 | 22 | 26 | 0.0006 |
| ASIANNH18 | 72 | 86 | 0.0094 | 23 | 19 | 0.0017 | 50 | 58 | 0.0022 | 304 | 303 | 0.0002 |
| HPINH18 | 4 | 4 | 0.0000 | 1 | 0 | 0.0005 | 4 | 3 | 0.0003 | 2 | 2 | 0.0000 |
| OTHERNH18 | 2 | 6 | 0.0026 | 3 | 2 | 0.0005 | 5 | 10 | 0.0014 | 43 | 43 | 0.0000 |
| MLTMNNH18 | 10 | 13 | 0.0020 | 4 | 2 | 0.0009 | 7 | 8 | 0.0003 | 52 | 51 | 0.0002 |

Table 2: Characteristics of Twelve Block Groups (continued)
( $C_{T D A}(g)$ counts result from 2020 Census Redistricting Data Production Settings ( $\epsilon=17.14$ for persons) version of TDA.)

| Demographic Group (g) | Block Group 130510107001 (GA) |  |  | Block Group$517100038001 \text { (VA) }$ |  |  | $\begin{gathered} \text { Block Group } \\ 121199112001 \text { (FL) } \end{gathered}$ |  |  | $\begin{gathered} \text { Block Group } \\ 060730187001(\mathrm{CA}) \end{gathered}$ |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $C_{S W A}$ | $C_{T D A}$ | $D R_{g}$ | $C_{S W A}$ | $C_{T D A}$ | $D R_{g}$ | $C_{S W A}$ | $C_{T D A}$ | $D R_{g}$ | $C_{S W A}$ | $C_{T D A}$ | $D R_{g}$ |
| TOTAL | 15,089 | 15,101 | $c$ | 19,506 | 19,512 | $c$ | 29,677 | 29,672 | $c$ | 37,452 | 37,453 | c |
| TOTAL18 | 11,561 | 11,567 | c | 19,486 | 19,460 | $c$ | 29,214 | 29,220 | c | 28,368 | 28,408 |  |
| TOTALHISP | 1,066 | 1,057 ${ }^{3}$ | 0.0007 | 2,599 | $\mathbf{2 , 5 9 2}{ }^{3}$ | 0.0004 | 502 | $501{ }^{2}$ | 0.0000 | 8,192 | $\mathbf{8 , 1 9 0}{ }^{2}$ | 0.0001 |
| TOTALNH | 14,023 | 14,044 | 0.0007 | 16,907 | 16,920 | 0.0004 | 29,175 | 29,171 | 0.0000 | 29,260 | 29,263 | 0.0001 |
| WHITENH | 7,901 | 7,923 ${ }^{1}$ | 0.0010 | 10,579 | 10,590 ${ }^{1}$ | 0.0004 | 28,555 | 28,550 ${ }^{1}$ | 0.0000 | 23,326 | 23,328 ${ }^{1}$ | 0.0000 |
| BLACKNH | 5,281 | $\mathbf{5 , 2 8 4}{ }^{2}$ | 0.0001 | 4,972 | $\mathbf{4 , 9 7 6}{ }^{2}$ | 0.0001 | 276 | $274{ }^{3}$ | 0.0001 | 3,040 | $\mathbf{3 , 0 4 7}{ }^{3}$ | 0.0002 |
| AIANNH | 54 | 51 | 0.0002 | 275 | 284 | 0.0005 | 58 | 62 | 0.0001 | 601 | 601 | 0.0000 |
| ASIANNH | 643 | 649 | 0.0004 | 776 | 782 | 0.0003 | 246 | 244 | 0.0001 | 1,422 | 1,427 | 0.0001 |
| HPINH | 17 | 12 | 0.0003 | 80 | 77 | 0.0002 | 7 | 7 | 0.0000 | 340 | 343 | 0.0001 |
| OTHERNH | 42 | 40 | 0.0001 | 45 | 43 | 0.0001 | 15 | 17 | 0.0001 | 89 | 91 | 0.0001 |
| MLTMNNH | 85 | 85 | 0.0000 | 180 | 168 | 0.0006 | 18 | 17 | 0.0000 | 442 | 426 | 0.0004 |
| HISP18 | 693 | 693 | 0.0000 | 2,597 | 2,583 | 0.0005 | 460 | 460 | 0.0000 | 5,506 | 5,539 | 0.0009 |
| NONHISP18 | 10,868 | 10,874 | 0.0000 | 16,889 | 16,877 | 0.0005 | 28,754 | 28,760 | 0.0000 | 22,862 | 22,869 | 0.0009 |
| WHITENH18 | 6,404 | 6,409 | 0.0001 | 10,562 | 10,562 | 0.0007 | 28,186 | 28,189 | 0.0001 | 18,751 | 18,745 | 0.0011 |
| BLACKNH18 | 3,849 | 3,860 | 0.0008 | 4,971 | 4,972 | 0.0004 | 247 | 249 | 0.0001 | 2,118 | 2,124 | 0.0001 |
| AIANNH18 | 46 | 40 | 0.0005 | 275 | 284 | 0.0005 | 58 | 56 | 0.0001 | 436 | 439 | 0.0001 |
| ASIANNH18 | 494 | 494 | 0.0000 | 776 | 780 | 0.0003 | 227 | 229 | 0.0001 | 1,032 | 1,036 | 0.0001 |
| HPINH18 | 9 | 10 | 0.0001 | 80 | 77 | 0.0001 | 7 | 7 | 0.0000 | 261 | 261 | 0.0000 |
| OTHERNH18 | 22 | 21 | 0.0001 | 45 | 42 | 0.0002 | 14 | 15 | 0.0000 | 62 | 60 | 0.0001 |
| MLTMNNH18 | 44 | 40 | 0.0003 | 180 | 160 | 0.0010 | 15 | 15 | 0.0000 | 202 | 204 | 0.0001 |

## I.4. THE QUESTION

More focused and concretely, we might proceed as follows to get an answer to our question at the national level (might also look at each state). To be more specific, imagine ordering the 217,740 block groups from smallest to largest $C_{S W A}$ counts for the demographic group TOTAL (Later, we focus only on block groups where $50 \leq C_{S W A} \leq 2,499$ ). To each block group in this ordering, imagine attaching its Table (as given for example in Tables 1a, 1b, or 2) of counts and difference of ratios values for all of the twenty demographic groups. To respond to our question, we seek to determine a value $C_{S W A}^{*}$ for the TOTAL block group such that for block groups whose TOTAL $C_{S W A}$ value is less than $C_{S W A}^{*}$, the differences of ratios of the twenty demographic groups tend to be large, i.e., the counts (or characteristics) are not reliable; also for block groups whose TOTAL $C_{S W A}$ values are greater than $C_{S W A}^{*}$, the differences of ratios of the twenty demographic groups tend to be small. See (2) below. (We use a similar ordering for MCDs and places, as well as for Congressional and state legislative districts.)

$$
\begin{equation*}
C_{S W A(1)} \leq C_{S W A(2)} \leq C_{S W A(3)} \leq \cdots \leq C_{S W A}^{*} \leq \cdots \leq C_{S W A(217,739)} \leq C_{S W A(217,740)}, \tag{2}
\end{equation*}
$$

where the $C_{S W A(i)}$ counts are the counts for the TOTAL block group $i$, for $i=1 ; 2 ; \ldots ; 217,740$.
Table 3 reveals an empirical answer to our question. For each block group, we consider three criteria (others could be considered) for the expression "reliable characteristics" based on the largest demographic group's (LDG) $D R_{g} \leq 0.01$; the largest demographic group's (LDG) $D R_{g} \leq 0.03$; and the largest demographic group's (LDG) $D R_{g} \leq 0.05$. For each criterion (column), Table 3 gives proportions of the number of block groups that satisfy the criterion for different strata of block groups based on TOTAL $C_{S W A}$ counts. For example, consider the 7,356 block groups in the stratum where " $700 \leq C_{S W A} \leq 749$ " for the TOTAL demographic group. We consider three (3) different criteria and present the proportion of block groups that satisfy Criterion I, or Criterion II, or Criterion III. For Criterion I (LDG $D R_{g} \leq 0.01$ ), 0.5007 (or $50.07 \%$ ) of the 7,356 block groups have $D R_{g} \leq 0.01$ for LDG counts. Because the proportions tend to increase as one goes down the Criterion I column, it seems that for each stratum below the stratum $700 \leq C_{S W A} \leq 749$ (i.e., those strata with larger block group TOTAL counts), one also tends to see that at least 0.5007 of
the block groups have $D R_{g} \leq 0.01$ for LDG counts. We observe a similar trend for the other two Criterion columns. For Criterion III (LDG $D R_{g} \leq 0.05$ ), 0.9826 (or $98.26 \%$ ) of the 7,356 block groups have $D R_{g} \leq 0.05$ for the block group's largest demographic group among TOTALHISP, WHITENH, BLACKNH, AIANNH, ASIANNH, and HPINH groups. We do not consider any block groups where the $C_{S W A}$ count for TOTAL block group is less than 50 or greater than 2,499. (Table $3^{\prime}$ of APPENDIX C gives analogous results as Table 3 for the 18 years and over population.)

Table 3: Proportion of Block Groups in Each Stratum for Three Criteria
(Proportion computations result from 2020 Census Redistricting Data Production Settings ( $\epsilon=17.14$ for persons) version of TDA.)
Population: United States (50 States \& DC)

| Stratum for Block Groups Using $C_{S W A}$ for TOTAL | Number of Block Groups | Reliable Characteristics Criteria |  |  |
| :---: | :---: | :---: | :---: | :---: |
|  |  | Criterion I <br> LDG $D R_{g} \leq 0.01$ | $\begin{array}{r} \text { Criterion II } \\ \text { LDG } D R_{g} \leq 0.03 \end{array}$ | $\begin{array}{r} \text { Criterion III } \\ \text { LDG } D R_{g} \leq 0.05 \end{array}$ |
| $50 \leq C_{S W A} \leq 99$ | 128 | 0.1250 | 0.3594 | 0.5156 |
| $100 \leq C_{S W A} \leq 149$ | 99 | 0.1818 | 0.5253 | 0.7071 |
| $150 \leq C_{S W A} \leq 199$ | 124 | 0.1694 | 0.5565 | 0.7581 |
| $200 \leq C_{S W A} \leq 249$ | 154 | 0.2662 | 0.6234 | 0.7922 |
| $250 \leq C_{S W A} \leq 299$ | 209 | 0.2919 | 0.6459 | 0.8565 |
| $300 \leq C_{S W A} \leq 349$ | 264 | 0.3636 | 0.7348 | 0.8902 |
| $350 \leq C_{S W A} \leq 399$ | 407 | 0.3366 | 0.7346 | 0.8698 |
| $400 \leq C_{S W A} \leq 449$ | 569 | 0.4077 | 0.7750 | 0.9315 |
| $450 \leq C_{S W A} \leq 499$ | 915 | 0.4087 | 0.8284 | 0.9552 |
| $500 \leq C_{S W A} \leq 549$ | 1,699 | 0.4197 | 0.8458 | 0.9588 |
| $550 \leq C_{S W A} \leq 599$ | 3,238 | 0.4546 | 0.8684 | 0.9654 |
| $600 \leq C_{S W A} \leq 649$ | 5,131 | 0.4578 | 0.8827 | 0.9751 |
| $650 \leq C_{S W A} \leq 699$ | 6,683 | 0.4718 | 0.8927 | 0.9753 |
| $700 \leq C_{S W A} \leq 749$ | 7,356 | 0.5007 | 0.9082 | 0.9826 |
| $750 \leq C_{S W A} \leq 799$ | 8,170 | 0.5160 | 0.9100 | 0.9845 |
| $800 \leq C_{S W A} \leq 849$ | 8,213 | 0.5268 | 0.9293 | 0.9897 |
| $850 \leq C_{S W A} \leq 899$ | 8,441 | 0.5517 | 0.9371 | 0.9914 |
| $900 \leq C_{S W A} \leq 949$ | 8,657 | 0.5557 | 0.9409 | 0.9920 |
| $950 \leq C_{S W A} \leq 999$ | 8,723 | 0.5849 | 0.9512 | 0.9952 |
| $1,000 \leq C_{S W A} \leq 1,049$ | 8,398 | 0.6044 | 0.9582 | 0.9952 |
| $1,050 \leq C_{S W A} \leq 1,099$ | 8,345 | 0.6192 | 0.9646 | 0.9965 |
| $1,100 \leq C_{S W A} \leq 1,149$ | 7,950 | 0.6244 | 0.9701 | 0.9972 |
| $1,150 \leq C_{S W A} \leq 1,199$ | 7,860 | 0.6422 | 0.9763 | 0.9977 |
| $1,200 \leq C_{S W A} \leq 1,249$ | 7,451 | 0.6515 | 0.9757 | 0.9988 |
| $1,250 \leq C_{S W A} \leq 1,299$ | 7,124 | 0.6645 | 0.9749 | 0.9978 |
| $1,300 \leq C_{S W A} \leq 1,349$ | 6,714 | 0.6822 | 0.9812 | 0.9988 |
| $1,350 \leq C_{S W A} \leq 1,399$ | 6,507 | 0.6859 | 0.9866 | 0.9989 |
| $1,400 \leq C_{S W A} \leq 1,449$ | 5,911 | 0.7090 | 0.9866 | 0.9992 |
| $1,450 \leq C_{S W A} \leq 1,499$ | 5,617 | 0.7002 | 0.9858 | 0.9995 |
| $1,500 \leq C_{S W A} \leq 1,549$ | 5,390 | 0.7330 | 0.9900 | 0.9994 |
| $1,550 \leq C_{S W A} \leq 1,599$ | 4,856 | 0.7341 | 0.9866 | 0.9994 |
| $1,600 \leq C_{S W A} \leq 1,649$ | 4,508 | 0.7420 | 0.9918 | 0.9998 |
| $1,650 \leq C_{S W A} \leq 1,699$ | 4,325 | 0.7489 | 0.9908 | 0.9998 |
| $1,700 \leq C_{S W A} \leq 1,749$ | 4,093 | 0.7669 | 0.9922 | 1.0000 |
| $1,750 \leq C_{S W A} \leq 1,799$ | 3,689 | 0.7650 | 0.9938 | 0.9997 |
| $1,800 \leq C_{S W A} \leq 1,849$ | 3,469 | 0.7530 | 0.9925 | 1.0000 |
| $1,850 \leq C_{S W A} \leq 1,899$ | 3,252 | 0.7811 | 0.9945 | 0.9997 |
| $1,900 \leq C_{S W A} \leq 1,949$ | 3,008 | 0.7793 | 0.9947 | 1.0000 |
| $1,950 \leq C_{S W A} \leq 1,999$ | 2,832 | 0.7970 | 0.9965 | 1.0000 |
| $2,000 \leq C_{S W A} \leq 2,049$ | 2,573 | 0.8022 | 0.9965 | 1.0000 |
| $2,050 \leq C_{S W A} \leq 2,099$ | 2,356 | 0.7975 | 0.9966 | 1.0000 |
| $2,100 \leq C_{S W A} \leq 2,149$ | 2,307 | 0.8331 | 0.9957 | 1.0000 |
| $2,150 \leq C_{S W A} \leq 2,199$ | 2,033 | 0.8170 | 0.9975 | 1.0000 |
| $2,200 \leq C_{S W A} \leq 2,249$ | 1,999 | 0.8354 | 0.9990 | 1.0000 |
| $2,250 \leq C_{S W A} \leq 2,299$ | 1,892 | 0.8494 | 0.9984 | 1.0000 |
| $2,300 \leq C_{S W A} \leq 2,349$ | 1,666 | 0.8331 | 0.9982 | 1.0000 |
| $2,350 \leq C_{S W A} \leq 2,399$ | 1,622 | 0.8453 | 0.9994 | 1.0000 |
| $2,400 \leq C_{S W A} \leq 2,449$ | 1,421 | 0.8621 | 0.9993 | 1.0000 |
| $2,450 \leq C_{S W A} \leq 2,499$ | 1,350 | 0.8600 | 1.0000 | 1.0000 |
| Total | 199,698 |  |  |  |

Using Criterion II and searching from top to bottom for the first stratum whose proportion is at least 0.9500: From Table 3, take $C_{S W A}^{*}$ to be between 950 and 999. For block groups whose TOTAL $C_{S W A}$ count is at least 999, the difference of ratios between the $C_{T D A}$ and $C_{S W A}$ ratios for the LDG will tend to be less than or equal to $3 \%$ (using our data).

Using Criterion III and searching from top to bottom for the first stratum whose proportion is at least 0.9500: From Table 3, take $C_{S W A}^{*}$ to be between 450 and 499. For block groups whose TOTAL $C_{S W A}$ count is at least 499, the difference of ratios between the $C_{T D A}$ and $C_{S W A}$ ratios for the LDG will tend to be less than or equal to $5 \%$ (using our data).

Using data to be released to the public (one run of the 2020 Census redistricting data production settings version of $T D A$ ), we might say, empirically based on the data for the block groups used in our study, that
"for any block group with a TOTAL count between 450 and 499 people, the difference between the TDA ratio of the largest demographic group (LDG) and the corresponding $S W A$ ratio for the $L D G$ is less than or equal to 5 percentage points at least $95 \%$ of the time".

The same version of the TDA was applied to the same underlying CEF data 25 independent times, i.e., for 25 additional runs. For each run, the stratum where we first observed that 0.9500 was exceeded is given in Table 4 for each run. (Table 3a' of APPENDIX $C$ gives analogous results as Table 3a for the 18 years and over population.)

Table 3a: For Each Run, the Stratum and Stratum Proportion When 0.9500 First Exceeded (Proportion computations result from 2020 Census Redistricting Data Production Settings ( $\epsilon=17.14$ for persons) version of TDA.)

Population: United States (50 States \& DC)

|  |  | $\begin{gathered} \text { Criterion III } \\ \text { LDG } D R_{g} \leq 0.05 \end{gathered}$ |
| :---: | :---: | :---: |
| $T D A$ Run | Stratum for Block Groups | Proportion When 0.9500 First Exceeded |
| 1 | $450 \leq C_{S W A} \leq 499$ | 0.9716 |
| 2 | $400 \leq C_{S W A} \leq 449$ | 0.9596 |
| 3 | $450 \leq C_{S W A} \leq 499$ | 0.9661 |
| 4 | $400 \leq C_{S W A} \leq 449$ | 0.9543 |
| 5 | $400 \leq C_{S W A} \leq 449$ | 0.9561 |
| 6 | $400 \leq C_{S W A} \leq 449$ | 0.9508 |
| 7 | $350 \leq C_{S W A} \leq 399$ | 0.9509 |
| 8 | $450 \leq C_{S W A} \leq 499$ | 0.9541 |
| 9 | $450 \leq C_{S W A} \leq 499$ | 0.9617 |
| 10 | $450 \leq C_{S W A} \leq 499$ | 0.9661 |
| 11 | $450 \leq C_{S W A} \leq 499$ | 0.9596 |
| 12 | $450 \leq C_{S W A} \leq 499$ | 0.9683 |
| 13 | $400 \leq C_{S W A} \leq 449$ | 0.9525 |
| 14 | $400 \leq C_{S W A} \leq 449$ | 0.9543 |
| 15 | $350 \leq C_{S W A} \leq 399$ | 0.9558 |
| 16 | $450 \leq C_{S W A} \leq 499$ | 0.9650 |
| 17 | $450 \leq C_{S W A} \leq 499$ | 0.9607 |
| 18 | $400 \leq C_{S W A} \leq 449$ | 0.9596 |
| 19 | $450 \leq C_{S W A} \leq 499$ | 0.9727 |
| 20 | $350 \leq C_{S W A} \leq 399$ | 0.9582 |
| 21 | $450 \leq C_{S W A} \leq 499$ | 0.9617 |
| 22 | $450 \leq C_{S W A} \leq 499$ | 0.9683 |
| 23 | $350 \leq C_{S W A} \leq 399$ | 0.9558 |
| 24 | $450 \leq C_{S W A} \leq 499$ | 0.9628 |
| 25 | $450 \leq C_{S W A} \leq 499$ | 0.9519 |

Each "block group" represents a type of defined geography used by the Census Bureau which is among a series of statistical and legal geographic entities that have a nesting relationship with
each other including: nation, state, county, tract, block group, and block. Many Census Bureau data products provide access to information about such nested geographies.

There are other types of defined geographies that are not a part of this nesting. These geographies (e.g., places, school districts, minor civil divisions,..) do not provide a complete national coverage and we consider them in this study as proxies for the yet to be defined electoral geography such as congressional, state legislative, and other electoral districts. [A Census Bureau designated place (CDP) is a statistical entity (geography) that is typically an unincorporated community, a concentration of population, housing, and commercial structures, identifiable by name, but not within an incorparated place. A Census Bureau incorporated place is a legally bounded entity, typically includes cities, towns (except in some states), villages, boroughs (except in New York and Alaska). A minor civil division (MCD) is a legally defined county subdivision. MCDs are the primary divisions of a county. They comprise both governmentally functioning entities-that is, those with elected or appointed officials who provide services and raise revenues-and nonfunctioning entities that exist primarily for administrative purposes, such as election districts. Source: Census Bureau]

## Analysis of Places and MCDs

As with Tables 3 and 3a for block groups, we present analagous results in Tables 4 and 4a using results from a single run and an additional 25 runs for all "places and MCDs". Altogether, we make use of 21,591 places and minor civil divisions (including $6,607,533$ blocks). Concerning the distribution of these places and MCDs using TOTAL counts, we note: Min $=0 ; 25^{t h}$ percentile $=547 ; 50^{t h}$ percentile $=2,065$; mean $=11,743 ; 75^{t h}$ percentile $=7,695$; Max $=3,796,060$. Again using Criterion III for all places and minor civil divisions in the United States, the stratum where we first observed that 0.9500 was exceeded is given in Table 4 a for each run. Also, see details of a single run in Table 4.

Table 4: Proportion of Places and MCDs in Each Stratum for Three Criteria
(Proportion computations result from 2020 Census Redistricting Data Production Settings ( $\epsilon=17.14$ for persons) version of TDA.)
Population: United States (50 States \& DC)

|  |  | Reliable Characteristics Criteria |  |  |
| :---: | :---: | :---: | :---: | :---: |
| Stratum for Places and MCDs Using $C_{S W A}$ for TOTAL | Number of Places and MCDs | $\begin{array}{r} \text { Criterion I } \\ \text { LDG } D R_{g} \leq 0.01 \end{array}$ | Criterion II <br> LDG $D R_{g} \leq 0.03$ | Criterion III $\text { LDG } D R_{g} \leq 0.05$ |
| $50 \leq C_{S W A} \leq 99$ | 573 | 0.2182 | 0.5969 | 0.7923 |
| $100 \leq C_{S W A} \leq 149$ | 622 | 0.4051 | 0.7540 | 0.9116 |
| $150 \leq C_{S W A} \leq 199$ | 645 | 0.3442 | 0.8109 | 0.9473 |
| $200 \leq C_{S W A} \leq 249$ | 598 | 0.4197 | 0.8361 | 0.9632 |
| $250 \leq C_{S W A} \leq 299$ | 500 | 0.4860 | 0.9000 | 0.9760 |
| $300 \leq C_{S W A} \leq 349$ | 448 | 0.5379 | 0.9152 | 0.9844 |
| $350 \leq C_{S W A} \leq 399$ | 417 | 0.5731 | 0.9233 | 0.9808 |
| $400 \leq C_{S W A} \leq 449$ | 399 | 0.6416 | 0.9449 | 0.9975 |
| $450 \leq C_{S W A} \leq 499$ | 344 | 0.6424 | 0.9680 | 0.9913 |
| $500 \leq C_{S W A} \leq 549$ | 341 | 0.6716 | 0.9765 | 0.9971 |
| $550 \leq C_{S W A} \leq 599$ | 291 | 0.7113 | 0.9691 | 0.9966 |
| $600 \leq C_{S W A} \leq 649$ | 277 | 0.6859 | 0.9783 | 1.0000 |
| $650 \leq C_{S W A} \leq 699$ | 306 | 0.7157 | 0.9902 | 1.0000 |
| $700 \leq C_{S W A} \leq 749$ | 254 | 0.7165 | 0.9843 | 1.0000 |
| $750 \leq C_{S W A} \leq 799$ | 233 | 0.7425 | 0.9914 | 1.0000 |
| $800 \leq C_{S W A} \leq 849$ | 255 | 0.7569 | 0.9608 | 1.0000 |
| $850 \leq C_{S W A} \leq 899$ | 222 | 0.7162 | 0.9955 | 1.0000 |
| $900 \leq C_{S W A} \leq 949$ | 201 | 0.7562 | 0.9851 | 1.0000 |
| $950 \leq C_{S W A} \leq 999$ | 210 | 0.7571 | 0.9952 | 1.0000 |
| $1,000 \leq C_{S W A} \leq 1,049$ | 223 | 0.7982 | 0.9955 | 1.0000 |
| $1,050 \leq C_{S W A} \leq 1,099$ | 157 | 0.8153 | 0.9873 | 1.0000 |
| $1,100 \leq C_{S W A} \leq 1,149$ | 194 | 0.7423 | 0.9897 | 0.9948 |
| $1,150 \leq C_{S W A} \leq 1,199$ | 178 | 0.8596 | 1.0000 | 1.0000 |
| $1,200 \leq C_{S W A} \leq 1,249$ | 162 | 0.8395 | 1.0000 | 1.0000 |
| $1,250 \leq C_{S W A} \leq 1,299$ | 174 | 0.8563 | 0.9885 | 1.0000 |
| $1,300 \leq C_{S W A} \leq 1,349$ | 164 | 0.8659 | 0.9939 | 1.0000 |
| $1,350 \leq C_{S W A} \leq 1,399$ | 166 | 0.8614 | 1.0000 | 1.0000 |
| $1,400 \leq C_{S W A} \leq 1,449$ | 134 | 0.9030 | 0.9851 | 0.9925 |
| $1,450 \leq C_{S W A} \leq 1,499$ | 153 | 0.8562 | 0.9935 | 1.0000 |
| $1,500 \leq C_{S W A} \leq 1,549$ | 147 | 0.9320 | 1.0000 | 1.0000 |
| $1,550 \leq C_{S W A} \leq 1,599$ | 135 | 0.8741 | 1.0000 | 1.0000 |
| $1,600 \leq C_{S W A} \leq 1,649$ | 124 | 0.9516 | 1.0000 | 1.0000 |
| $1,650 \leq C_{S W A} \leq 1,699$ | 139 | 0.9137 | 1.0000 | 1.0000 |
| $1,700 \leq C_{S W A} \leq 1,749$ | 141 | 0.8794 | 1.0000 | 1.0000 |
| $1,750 \leq C_{S W A} \leq 1,799$ | 127 | 0.8740 | 1.0000 | 1.0000 |
| $1,800 \leq C_{S W A} \leq 1,849$ | 134 | 0.8881 | 1.0000 | 1.0000 |
| $1,850 \leq C_{S W A} \leq 1,899$ | 117 | 0.8803 | 1.0000 | 1.0000 |
| $1,900 \leq C_{S W A} \leq 1,949$ | 108 | 0.9259 | 0.9907 | 1.0000 |
| $1,950 \leq C_{S W A} \leq 1,999$ | 120 | 0.9000 | 1.0000 | 1.0000 |
| $2,000 \leq C_{S W A} \leq 2,049$ | 106 | 0.9340 | 1.0000 | 1.0000 |
| $2,050 \leq C_{S W A} \leq 2,099$ | 100 | 0.8700 | 0.9900 | 1.0000 |
| $2,100 \leq C_{S W A} \leq 2,149$ | 110 | 0.9000 | 1.0000 | 1.0000 |
| $2,150 \leq C_{S W A} \leq 2,199$ | 105 | 0.9429 | 1.0000 | 1.0000 |
| $2,200 \leq C_{S W A} \leq 2,249$ | 95 | 0.9474 | 1.0000 | 1.0000 |
| $2,250 \leq C_{S W A} \leq 2,299$ | 77 | 0.9351 | 1.0000 | 1.0000 |
| $2,300 \leq C_{S W A} \leq 2,349$ | 111 | 0.8919 | 1.0000 | 1.0000 |
| $2,350 \leq C_{S W A} \leq 2,399$ | 109 | 0.9450 | 1.0000 | 1.0000 |
| $2,400 \leq C_{S W A} \leq 2,449$ | 83 | 0.9398 | 1.0000 | 1.0000 |
| $2,450 \leq C_{S W A} \leq 2,499$ | 94 | 0.9149 | 1.0000 | 1.0000 |
| Total | 199,698 |  |  |  |

Table 4a: For Each Run, the Stratum and Stratum Proportion When 0.9500 First Exceeded
(Proportion computations result from 2020 Census Redistricting Data Production Settings ( $\epsilon=17.14$ for persons) version of $T D A$.)
Population: United States (50 States \& DC)

|  |  | Criterion III $\text { LDG } D R_{g} \leq 0.05$ |
| :---: | :---: | :---: |
| $T D A$ Run | Stratum for Places \& MCDs | Proportion When 0.9500 First Exceeded |
| 1 | $150 \leq C_{S W A} \leq 199$ | 0.9504 |
| 2 | $200 \leq C_{S W A} \leq 249$ | 0.9548 |
| 3 | $150 \leq C_{S W A} \leq 199$ | 0.9566 |
| 4 | $150 \leq C_{S W A} \leq 199$ | 0.9504 |
| 5 | $200 \leq C_{S W A} \leq 249$ | 0.9632 |
| 6 | $200 \leq C_{S W A} \leq 249$ | 0.9632 |
| 7 | $200 \leq C_{S W A} \leq 249$ | 0.9615 |
| 8 | $200 \leq C_{S W A} \leq 249$ | 0.9615 |
| 9 | $200 \leq C_{S W A} \leq 249$ | 0.9582 |
| 10 | $150 \leq C_{S W A} \leq 199$ | 0.9643 |
| 11 | $150 \leq C_{S W A} \leq 199$ | 0.9566 |
| 12 | $150 \leq C_{S W A} \leq 199$ | 0.9504 |
| 13 | $200 \leq C_{S W A} \leq 249$ | 0.9615 |
| 14 | $150 \leq C_{S W A} \leq 199$ | 0.9550 |
| 15 | $200 \leq C_{S W A} \leq 249$ | 0.9565 |
| 16 | $150 \leq C_{S W A} \leq 199$ | 0.9550 |
| 17 | $200 \leq C_{S W A} \leq 249$ | 0.9515 |
| 18 | $150 \leq C_{S W A} \leq 199$ | 0.9519 |
| 19 | $150 \leq C_{S W A} \leq 199$ | 0.9504 |
| 20 | $200 \leq C_{S W A} \leq 249$ | 0.9532 |
| 21 | $200 \leq C_{S W A} \leq 249$ | 0.9615 |
| 22 | $200 \leq C_{S W A} \leq 249$ | 0.9548 |
| 23 | $150 \leq C_{S W A} \leq 199$ | 0.9566 |
| 24 | $150 \leq C_{S W A} \leq 199$ | 0.9550 |
| 25 | $150 \leq C_{S W A} \leq 199$ | 0.9519 |

Using the data that will be released to the public (one run of the 2020 Census Production Settings version of $T D A$ ), we might say (as we did with block groups), empirically based on the data for the MCDs and places used in our study, that
"for any MCD or place with a TOTAL count between 200 and 249 people, the difference between the TDA ratio of the largest demographic group ( $L D G$ ) and the corresponding $S W A$ ratio for the $L D G$ is less than or equal to 5 percentage points at least $95 \%$ of the time".

## Analysis of Congressional and State Legislative Districts

Another type of defined geography that is not a part of this nesting includes Congressional districts and state legislative districts. As we will see with Rhode Island in Part II of this study report, each state has Congressional district(s) (CD), state legislative districts in an upper chamber (SLDU), and state legislative districts in a lower chamber (SLDL).

As with the summary display in Table 3a for block groups and the summary display in Table 4a for places and MCDs, we use results from the 25 runs for all "Congressional and state legislative districts". Altogether, we make use of all $7,167(=436+1,946+4,785)$ Congressional and state legislative districts in the United States. The Table below gives a few parameters for the national accounting of these districts.

|  | CD | SLDU | SLDL |
| :---: | ---: | :---: | :---: |
| Number of Districts | 436 | 1,946 | 4,785 |
| Min Population | 526,283 | 13,629 | 3,173 |
| Median Population | 705,831 | 121,212 | 41,713 |
| Mean Population | 708,132 | 158,656 | 64,016 |
| Max Population | 989,415 | 940,612 | 470,325 |

Again using Criterion III for all Congressional and state legislative districts in the United States, the stratum, where we first observed that 0.9500 was exceeded is given in Table 5 for each run. We display the entire table to emphasize that for each and every one of these districts, the size is sufficiently large to believe that the $T D A$ counts are reliable for the largest demographic group (LDG) "all" of the time (based on our data).

Using the data that will be released to the public (one run of the 2020 Census redistricting data production settings version of $T D A$ ), we might say (as we did with block groups, also with MCDs and places) based on Table 5, that
"for all Congressional and state legislative districts, the difference between the TDA ratio of the largest demographic group ( $L D G$ ) and the corresponding $S W A$ ratio for the $L D G$ is less than or equal to 5 percentage points $100 \%$ of the time".

Table 5: For Each Run, the Stratum and Stratum Proportion When 0.9500 First Exceeded (Proportion computations result from 2020 Census Redistricting Data Production Settings ( $\epsilon=17.14$ for persons) version of $T D A$.) Population: United States (50 States \& DC)

|  |  | Criterion III LDG $D R_{g} \leq 0.05$ |
| :---: | :---: | :---: |
| $T D A$ Run | Stratum for Congressional \& State Legislative Districts | Proportion When 0.9500 First Exceeded |
| 1 | $3,150 \leq C_{S W A} \leq 3,199$ | 1.0000 |
| 2 | $3,150 \leq C_{S W A} \leq 3,199$ | 1.0000 |
| 3 | $3,150 \leq C_{S W A} \leq 3,199$ | 1.0000 |
| 4 | $3,150 \leq C_{S W A} \leq 3,199$ | 1.0000 |
| 5 | $3,150 \leq C_{S W A} \leq 3,199$ | 1.0000 |
| 6 | $3,150 \leq C_{S W A} \leq 3,199$ | 1.0000 |
| 7 | $3,150 \leq C_{S W A} \leq 3,199$ | 1.0000 |
| 8 | $3,150 \leq C_{S W A} \leq 3,199$ | 1.0000 |
| 9 | $3,150 \leq C_{S W A} \leq 3,199$ | 1.0000 |
| 10 | $3,150 \leq C_{S W A} \leq 3,199$ | 1.0000 |
| 11 | $3,150 \leq C_{S W A} \leq 3,199$ | 1.0000 |
| 12 | $3,150 \leq C_{S W A} \leq 3,199$ | 1.0000 |
| 13 | $3,150 \leq C_{S W A} \leq 3,199$ | 1.0000 |
| 14 | $3,150 \leq C_{S W A} \leq 3,199$ | 1.0000 |
| 15 | $3,150 \leq C_{S W A} \leq 3,199$ | 1.0000 |
| 16 | $3,150 \leq C_{S W A} \leq 3,199$ | 1.0000 |
| 17 | $3,150 \leq C_{S W A} \leq 3,199$ | 1.0000 |
| 18 | $3,150 \leq C_{S W A} \leq 3,199$ | 1.0000 |
| 19 | $3,150 \leq C_{S W A} \leq 3,199$ | 1.0000 |
| 20 | $3,150 \leq C_{S W A} \leq 3,199$ | 1.0000 |
| 21 | $3,150 \leq C_{S W A} \leq 3,199$ | 1.0000 |
| 22 | $3,150 \leq C_{S W A} \leq 3,199$ | 1.0000 |
| 23 | $3,150 \leq C_{S W A} \leq 3,199$ | 1.0000 |
| 24 | $3,150 \leq C_{S W A} \leq 3,199$ | 1.0000 |
| 25 | $3,150 \leq C_{S W A} \leq 3,199$ | 1.0000 |

## I.5. CONCLUDING REMARKS FOR PART I

Remark 1: Within each of the criterion columns across Table 3, the values of the proportions tend to increase (though not always) as we go from the stratum with the smallest block groups to the stratum with the largest block groups using the $S W A$ TOTAL counts. Also, the values of the proportions within a stratum (row) do increase as we go from Criterion I to Criterion III. From Table 3, we believe that a value for $C_{S W A}^{*}$ can be produced (which is based on the data used in this study). This $C_{S W A}^{*}$ is an empirical result. We can make similar statements relating to MCDs and places using Table 4, as well statements relating to Congressional and state legislative districts using Table 5.

Remark 2: Much of our focus in Part I has been in the context of the total population characteristics for block groups, MCDs and places, and Congressional and state legislative districts. In Table $3^{\prime}$ of APPENDIX C, we performed an analysis for the over 18 years and over population characteristics for block groups similar to what was done in Table 3 for the total population characteristics. We observed that the 5 percentage point criterion is reached $95 \%$ of the time for TOTAL18 in block groups whose size range between 500 and 549 people.

Remark 3: While small demographic groups are important, in the context of redistricting, it is the largest among the demographic groups that have the potential to form electoral districts where sufficiently large (and compact) minority groups have the opportunity "to elect representatives of their choice". We believe that support for consideration of the largest demographic group(s) is as noted in Section 2 of the Voting Rights Act of 1965 (as amended) and is called for by one of the
three Gingles Requirements in the U.S. Supreme Court case Thornburg v. Gingles (1986) when establishing a violation of Section 2.

We understand that the potential for creating an electoral district that provides minority citizens with the opportunity to elect candidates of their choice is not necessarily limited to those block groups in which that group is the "largest demographic group". For example, a demographic group could comprise the second largest population group in two or more contiguous, randomly-created block groups. A different configuration of constituent blocks could result in that group being the basis of a district that affords the requisite opportunity to elect a candidate of their choice.

## Part II <br> VARIABILITY ASSESSMENT OF DATA TREATED BY THE TOPDOWN ALGORITHM

## II.1. INTRODUCTION

Part II is an update of our results in [6] where $\epsilon=10.3$ and the 2021-04-28 version of TDA was used; whereas, throughout this study, $\epsilon=17.14$ and further advances have been made resulting in the 2020 Census redistricting data production settings version of TDA. We reuse wording from [6] in many places; we do this in an attempt to repeat some of what we feel is important and in making this a more complete document. Of course, specific data results will differ.

As in [6], the specific focus of Part II is whether the explicitly acknowledged randomness used in the TDA for disclosure avoidance in the 2020 Census delivers official data that are fit for the development and analysis of redistricting plans. That randomness is characterized in this paper by measures of the variability observed in 25 runs of the same version of the TDA using the same allocation of the privacy-loss budget in each run $(\epsilon=17.14)$. The variability inherent in the official 2010 PL-94-171 redistricting data resulted primarily from disclosure avoidance via household swapping. The parameters defining the rule(s) used in swapping that resulted in the official 2010 redistricting data are confidential and no estimates of the resulting variability have ever been published, including in this paper. Our approach (in the rest of this study as was the case in our earlier study [6]) has two parts: (1) to report observations on variability of results among 25 runs of the TDA [1] for Rhode Island and (2) to report observations on variability between the results among the 25 runs of the TDA and the published 2010 Census Public Law 94-171 data for Rhode Island. In Part II, we also repeat these two-part analyses for three specific cases provided by the DOJ.

## 2010 Census Data for Rhode Island

The TDA was applied to data in the 2010 CEF for Rhode Island twenty-five different times, which we refer to as twenty-five runs of the TDA. For each run and for each of the 25,181 blocks in Rhode Island in the 2010 Census, various demographic variables report counts of various combinations of race, ethnicity (Hispanic or not Hispanic), and age.

Rhode Island has two (2) Congressional districts (CD), 38 state legislative districts (SLDU) in its upper legislative chamber, and 75 state legislative districts (SLDL) in its lower legislative chamber. These form the foundation of our case study for Rhode Island.

## 2010 Census Data for Three Cases Provided by DOJ

For three cases (jurisdictions) provided by DOJ, we conduct similar analyses of data in Section II. 6 as just described for Rhode Island. The three cases are Panola County, Mississippi (MS) (2,180 blocks); Tate County (School District), MS (784 blocks); and Tylertown (Walthall County), MS (136 blocks).

## Overview of Part II

An overview of Part II follows. In Section II. 2 of this report, we present data for the two Congressional districts of Rhode Island and using formatted data tables as shown in Table 6. Section II. 3 visually compares 2010 CEF data treated by the disclosure avoidance method (swapping [6]) with randomly selected runs of the same 2010 CEF data treated by the TDA method (i.e., differential privacy) being planned for use by the 2020 Census. Section II. 4 is similar to Section II. 3 except the visual comparisons are for four of Rhode Island's Upper Chamber Districts. Section II. 5 is similar to Sections II. 3 and II. 4 except the visual comparisons are for four of Rhode Island's Lower Chamber Districts. Section II. 6 investigates three cases provided by DOJ using varying (mainly smaller) total population and varying group composition selected for comparisons similar to those of previous Sections for CDs, SLDUs, and SLDLs. Section II. 7 defines and looks at variablility among the $25 T D A$ runs of Rhode Island data using the planned $T D A$ method of 2020, and it also
looks at variability among the $25 T D A$ runs in comparison with the public data for Rhode Island from 2010 (this section also presents similar tables for the three cases provided by DOJ). The insert following Table 6 gives a suggestion for reviewing the tables of counts and percentages. The key empirical message on variability is given in the last paragraph of Section II.7. Section II. 8 provides some concluding remarks based on the tables. The APPENDICES follow Section II.8.

## II.2. FORMAT OF COUNTS \& PERCENTAGES TABLES USED IN OUR STUDY

Table 6 shows the redistricting plan (POST-2010) adopted by Panola County, Mississippi. Panola County, with five (5) districts, has an overall population (TOTAL) of 34,707 people based on the 2010 Census. The average population per district (IDEAL POPULATION) is $34,707 / 5=$ 6,941 people. Using the POST-2010 plan, the deviations from the IDEAL POPULATION for each of the 5 districts (DEV) are 33, -392, 133, 164, and 64, respectively; and the corresponding percent deviations $(\mathrm{DEV}=\mathrm{DEV} / 6941) \times 100 \%$ are respectively: $0.48 \%,-5.65 \%,-1.92 \%, 2.36 \%$, and $0.92 \%$. From Table 6, it is noteworthy that the demographic group of WHITENH has 16,981 people which is WHITENHP $=48.93 \%$ of the county's population while the demographic group BLACKNH has 16,899 people which is BLACKNHP $=48.69 \%$ of the county's population. Other demographic group characteristics in Table 6 are given for the 18 years and over population (TOTAL18).

Table 6. POST-2010 Census Demographics, Counts, \& Percentages: Panola County, Mississippi

| Demographics |  | Counts \& Percentages by District (POST-2010) |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| DIST-ID | Panola | 1 | 2 | 3 | 4 | 5 |
| TOTAL | 34,707 | 6,974 | 6,549 | 7,074 | 7,105 | 7,005 |
| DEV |  | 33 | -392 | 133 | 164 | 64 |
| DEVP |  | 0.48 | -5.65 | 1.92 | 2.36 | 0.92 |
| TOTAL18 | 25,363 | 5,214 | 4,732 | 5,171 | 5,345 | 4,901 |
| TOTALHISP | 494 | 66 | 75 | 85 | 120 | 148 |
| TOTALHISPP | 1.42 | 0.95 | 1.15 | 1.20 | 1.69 | 2.11 |
| TOTALNH | 34,213 | 6,908 | 6,474 | 6,989 | 6,985 | 6,857 |
| TOTALNHP | 98.58 | 99.05 | 98.85 | 98.8 | 98.31 | 97.89 |
| WHITENH | 16,981 | 2,419 | 2,096 | 4,030 | 5,250 | 3,186 |
| WHITENHP | 48.93 | 34.69 | 32.00 | 56.97 | 73.89 | 45.48 |
| BLACKNH | 16,899 | 4,427 | 4,332 | 2,925 | 1,658 | 3,557 |
| BLACKNHP | 48.69 | 63.48 | 66.15 | 41.35 | 23.34 | 50.78 |
| AIANNH | 148 | 26 | 20 | 15 | 38 | 49 |
| AIANNHP | 0.43 | 0.37 | 0.31 | 0.21 | 0.53 | 0.70 |
| ASIANNH | 89 | 8 | 7 | 5 | 17 | 52 |
| ASIANNHP | 0.26 | 0.11 | 0.11 | 0.07 | 0.24 | 0.74 |
| HPINH | 4 | 0 | 0 | 0 | 2 | 2 |
| HPINHP | 0.01 | 0.00 | 0.00 | 0.00 | 0.03 | 0.03 |
| OTHERNH | 19 | 7 | 5 | 1 | 3 | 3 |
| OTHERNHP | 0.05 | 0.10 | 0.08 | 0.01 | 0.04 | 0.04 |
| MLTMNNH | 73 | 21 | 14 | 13 | 17 | 8 |
| MLTMNNHP | 0.21 | 0.30 | 0.21 | 0.18 | 0.24 | 0.11 |
| HISP18 | 298 | 44 | 44 | 52 | 63 | 95 |
| HISP18P | 1.17 | 0.84 | 0.93 | 1.01 | 1.18 | 1.94 |
| NONHISP18 | 25,065 | 5,170 | 4,688 | 5,119 | 5,282 | 4,806 |
| NONHISP18P | 98.83 | 99.16 | 99.07 | 98.99 | 98.82 | 98.06 |
| WHITENH18 | 13,455 | 2,025 | 1,732 | 3,072 | 4,115 | 2,511 |
| WHITENH18P | 53.05 | 38.84 | 36.6 | 59.41 | 76.99 | 51.23 |
| BLACKNH18 | 11,394 | 3,099 | 2,928 | 2,024 | 1,118 | 2,225 |
| BLACKNH18P | 44.92 | 59.44 | 61.88 | 39.14 | 20.92 | 45.40 |
| AIANNH18 | 115 | 21 | 16 | 11 | 29 | 38 |
| AIANNH18P | 0.45 | 0.40 | 0.34 | 0.21 | 0.54 | 0.78 |
| ASIANNH18 | 54 | 8 | 5 | 2 | 12 | 27 |
| ASIANNH18P | 0.21 | 0.15 | 0.11 | 0.04 | 0.22 | 0.55 |
| HPINH18 | 2 | 0 | 0 | 0 | 1 | 1 |
| HPINH18P | 0.01 | 0.00 | 0.00 | 0.00 | 0.02 | 0.02 |
| OTHERNH18 | 5 | 1 | 0 | 1 | 2 | 1 |
| OTHERNH18P | 0.02 | 0.02 | 0.00 | 0.02 | 0.04 | 0.02 |
| MLTMNH18 | 40 | 16 | 7 | 9 | 5 | 3 |
| MLTMNH18P | 0.16 | 0.31 | 0.15 | 0.17 | 0.09 | 0.06 |

Source: U.S. Department of Justice, Washington, D.C.

## A Suggestion from the Authors for Reviewing Each Table

When we inspect the various tables that follow in this study, we first look at the column of overall counts and percentages for the various demographic groups in a jurisdiction (e.g., state or county or school district) and then ask how these counts and percentages are distributed over the various districts.

## II.3. EXAMINATION OF RHODE ISLAND CONGRESSIONAL DISTRICT DATA

Table 7 shows results from three randomly chosen runs of the twenty-five runs of the $T D A$ for Congressional Districts CD-01 and CD-02 for Rhode Island (last six columns) and displays them with the counts from the 2010 Census (alternately referred to as swapping or Summary File 1 (SF1) in this part of our study) relative to the boundaries for the $113^{\text {th }}$ Congress. These three runs provide a taste of what variability might be expected among the various runs of the TDA. Throughout this report, we use the same value of $\epsilon=17.14$, and exactly the same implementation code and parameters, for all discussed runs of the TDA.

In Table 7, we also compare the results for CD-01 and CD-02 from each of the three TDA runs with the corresponding published results (2010 Census, SF1) for CD-01 and CD-02.

From Table 7, while the corresponding counts for each demographic group (on each row) vary among the runs as well as relative to the released 2010 Census counts, the corresponding percentages displayed differ by less than 0.5 of a percentage point for all demographic groups. The fact that the DEV values for the three runs differ from -0.5 and 0.5 should be of no concern because the 2020 Congressional redistricting would use the noise-infused block level counts to create Congressional districts where the DEV values differ by no more than 1 person. In general, state legislative districts are allowed to deviate by more than 1 person.

In Table 7, note that CD-01 has smaller counts for WHITENH than CD-02 using the 2010 Census counts. As a consequence, CD-01 has comparatively larger counts for most minority demographic groups than CD-02. This observation is true for the total population group counts as well as for the 18 and older population groups. This observation tends to also hold for each of the three TDA runs. (The same holds true for WHITENH18 and most minority groups in the 18 and older population.)

## II.4. EXAMINATION OF RHODE ISLAND's 38 UPPER CHAMBER DISTRICTS

There are 38 districts with one legislator each in Rhode Island's Upper Chamber. Therefore, the IDEAL POPULATION for each State Upper Chamber District is $\frac{1,052,567}{38}=27,699.1$. Columns 2-5 of Table 8 give 2010 Census counts and percentages for the State Upper Chamber Districts (SLDU) 01, 02, 03, and 04. Columns 6-9 of Table 8 give corresponding counts and percentages from the same TDA Run A noted in Table 7.

For the 2010 Census counts as well as the counts for the TDA Run A, SLDU-02 has relatively high percentages for both TOTALHISPP and HISP18P. Similarly, for the 2010 Census counts as well as for the TDA Run A, SLDU-03 and SLDU-04 each has relatively high percentages for both WHITENHP and WHITENH18P. SLDU-01 has a relatively high percentage total for TOTALHISPP and BLACKNHP. The same holds true in SLDU-01 for HISP18P and BLACKNH18P.

Table 7. Rhode Island: Three of Twenty-five Runs of the TDA by Congressional Districts (CDs) for the $113^{t h}$ Congress
( $C_{T D A}(g)$ counts result from 2020 Census Redistricting Data Production Settings ( $\epsilon=17.14$ for persons) version of TDA.)

|  |  | $\begin{gathered} 2010 \mathrm{C} \\ (P L 94 \\ \text { Counts \& } \\ \text { POST- } \end{gathered}$ | ensus, SF1 <br> -171)(2013) <br> \& Percentages <br> 2010 Plan | Counts \& Percentages, $113^{\text {th }}$ Congress 3 Out of 25 Runs of the TDA |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Demographics |  | $113{ }^{\text {th }}$ | Congress | TDA-Run A |  | $T D A$-Run B |  | TDA-Run C |  |
| DIST-ID | Rhode Island | CD-01 | CD-02 | CD-01 | CD-02 | CD-01 | CD-02 | CD-01 | CD-02 |
| TOTAL | 1,052,567 | 526,283 | 526,284 | 526,125 | 526,442 | 526,120 | 526,447 | 526,140 | 526,427 |
| DEV |  | -0.5 | 0.5 | -158.5 | 158.5 | -163.5 | 163.5 | -143.5 | 143.5 |
| DEVP |  | 0.00 | 0.00 | -0.03 | 0.03 | -0.03 | 0.03 | -0.03 | 0.03 |
| TOTAL18 | 828,611 | 412,778 | 415,833 | 412,604 | 416,009 | 412,633 | 415,956 | 412,648 | 415,951 |
| TOTALHISP | 130,655 | 76,100 | 54,555 | 76,050 | 54,607 | 76,081 | 54,575 | 76,070 | 54,603 |
| TOTALHISPP | 12.41 | 14.46 | 10.37 | 14.45 | 10.37 | 14.46 | 10.37 | 14.46 | 10.37 |
| TOTALNH | 921,912 | 450,183 | 471,729 | 450,075 | 471,835 | 450,039 | 471,872 | 450,070 | 471,824 |
| TOTALNHP | 87.59 | 85.54 | 89.63 | 85.55 | 89.63 | 85.54 | 89.63 | 85.54 | 89.63 |
| WHITENH | 803,685 | 377,109 | 426,576 | 377,055 | 426,626 | 377,035 | 426,649 | 377,057 | 426,628 |
| WHITENHP | 76.35 | 71.66 | 81.05 | 71.67 | 81.04 | 71.66 | 81.04 | 71.66 | 81.04 |
| BLACKNH | 57,927 | 37,627 | 20,300 | 37,600 | 20,340 | 37,678 | 20,255 | 37,619 | 20,301 |
| BLACKNHP | 5.50 | 7.15 | 3.86 | 7.15 | 3.86 | 7.16 | 3.85 | 7.15 | 3.86 |
| AIANNH | 6,839 | 3,142 | 3,697 | 3,161 | 3,694 | 3,119 | 3,739 | 3,084 | 3,759 |
| AIANNHP | 0.65 | 0.60 | 0.70 | 0.60 | 0.70 | 0.59 | 0.71 | 0.59 | 0.71 |
| ASIANNH | 34,194 | 17,705 | 16,489 | 17,664 | 16,518 | 17,676 | 16,516 | 17,722 | 16,473 |
| ASIANNHP | 3.25 | 3.36 | 3.13 | 3.36 | 3.14 | 3.36 | 3.14 | 3.37 | 3.13 |
| HPINH | 655 | 383 | 272 | 392 | 263 | 411 | 258 | 361 | 294 |
| HPINHP | 0.06 | 0.07 | 0.05 | 0.07 | 0.05 | 0.08 | 0.05 | 0.07 | 0.06 |
| OTHERNH | 10,296 | 8,492 | 1,804 | 8,484 | 1,818 | 8,487 | 1,802 | 8,517 | 1,763 |
| OTHERNHP | 0.98 | 1.61 | 0.34 | 1.61 | 0.35 | 1.61 | 0.34 | 1.62 | 0.33 |
| MLTMNNH | 8,316 | 5,725 | 2,591 | 5,719 | 2,576 | 5,633 | 2,653 | 5,710 | 2,606 |
| MLTMNNHP | 0.79 | 1.09 | 0.49 | 1.09 | 0.49 | 1.07 | 0.50 | 1.09 | 0.50 |
| HISP18 | 84,715 | 49,303 | 35,412 | 49,146 | 35,562 | 49,240 | 35,469 | 49,237 | 35,487 |
| HISP18P | 10.22 | 11.94 | 8.52 | 11.91 | 8.55 | 11.93 | 8.53 | 11.93 | 8.53 |
| NONHISP18 | 743,896 | 363,475 | 380,421 | 363,458 | 380,447 | 363,393 | 380,487 | 363,411 | 380,464 |
| NONHISP18P | 89.78 | 88.06 | 91.48 | 88.09 | 91.45 | 88.07 | 91.47 | 88.07 | 91.47 |
| WHITENH18 | 660,823 | 312,240 | 348,583 | 312,217 | 348,607 | 312,174 | 348,648 | 312,219 | 348,605 |
| WHITENH18P | 79.75 | 75.64 | 83.83 | 75.67 | 83.80 | 75.65 | 83.82 | 75.66 | 83.81 |
| BLACKNH18 | 39,485 | 25,402 | 14,083 | 25,392 | 14,105 | 25,458 | 14,029 | 25,405 | 14,081 |
| BLACKNH18P | 4.77 | 6.15 | 3.39 | 6.15 | 3.39 | 6.17 | 3.37 | 6.16 | 3.39 |
| AIANNH18 | 4,963 | 2,332 | 2,631 | 2,344 | 2,626 | 2,298 | 2,684 | 2,306 | 2,657 |
| AIANNH18P | 0.60 | 0.56 | 0.63 | 0.57 | 0.63 | 0.56 | 0.65 | 0.56 | 0.64 |
| ASIANNH18 | 25,333 | 13,276 | 12,057 | 13,252 | 12,072 | 13,235 | 12,090 | 13,280 | 12,051 |
| ASIANNH18P | 3.06 | 3.22 | 2.90 | 3.21 | 2.90 | 3.21 | 2.91 | 3.22 | 2.90 |
| HPINH18 | 500 | 307 | 193 | 305 | 189 | 326 | 186 | 302 | 195 |
| HPINH18P | 0.06 | 0.07 | 0.05 | 0.07 | 0.05 | 0.08 | 0.04 | 0.07 | 0.05 |
| OTHERNH18 | 7,290 | 6,061 | 1,229 | 6,069 | 1,227 | 6,113 | 1,177 | 6,070 | 1,208 |
| OTHERNH18P | 0.88 | 1.47 | 0.30 | 1.47 | 0.29 | 1.48 | 0.28 | 1.47 | 0.29 |
| MLTMNH18 | 5,502 | 3,857 | 1,645 | 3,879 | 1,621 | 3,789 | 1,673 | 3,829 | 1,667 |
| MLTMNH18P | 0.66 | 0.93 | 0.40 | 0.94 | 0.39 | 0.92 | 0.40 | 0.93 | 0.40 |

Source: Data from 3 Runs of the TDA, U. S. Bureau of the Census, Washington, D.C.
Selected observations for Table 7:
1: Corresponding percentages between the 2010 Census data and the TDA data on each row displayed in Table 7 differ by less than 0.5 of a percentage point for all demographic groups.
2: CD-01 has lower counts for WHITENH (also WHITENH18) than CD-02 when using the 2010 Census counts. As a consequence, CD-01 has comparatively larger counts for most minority demographic groups than CD-02. The same relationships between the CD-01 and CD-02 data hold for these demographic groups within the 18 and older population groups. This observation also tends to hold for each of the three TDA runs.

Table 8. Rhode Island Run A of Twenty-five Runs of the TDA for State Upper Chamber Districts (SLDU) 01, 02, 03, and 04 (4 of 38 Districts)
( $C_{T D A}(g)$ counts result from 2020 Census Redistricting Data Production Settings ( $\epsilon=17.14$ for persons) version of TDA.)

|  | 2010 Census, SF1 <br> (PL 94-171) (2013) <br> Counts \& Percentages <br> POST-2010 Plan |  |  |  | Counts \& Percentages, 2013 Run A of the $T D A$ |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Demographics |  |  |  |  |  |  |  |  |
| DIST-ID | SLDU-01 | SLDU-02 | SLDU-03 | SLDU-04 | SLDU-01 | SLDU-02 | SLDU-03 | SLDU-04 |
| TOTAL | 28,161 | 28,079 | 28,398 | 28,201 | 28,095 | 28,151 | 28,614 | 28,196 |
| DEV | 461.9 | 379.9 | 698.9 | 501.9 | 395.9 | 451.9 | 914.9 | 496.9 |
| DEVP | 1.64 | 1.35 | 2.46 | 1.78 | 1.41 | 1.61 | 3.20 | 1.76 |
| TOTAL18 | 20,914 | 19,846 | 25,361 | 23,599 | 20,876 | 19,853 | 25,435 | 23,579 |
| TOTALHISP | 10,282 | 16,288 | 1,409 | 3,217 | 10,254 | 16,312 | 1,510 | 3,177 |
| TOTALHISPP | 36.51 | 58.01 | 4.96 | 11.41 | 36.50 | 57.94 | 5.28 | 11.27 |
| TOTALNH | 17,879 | 11,791 | 26,989 | 24,984 | 17,841 | 11,839 | 27,104 | 25,019 |
| TOTALNHP | 63.49 | 41.99 | 95.04 | 88.59 | 63.50 | 42.06 | 94.72 | 88.73 |
| WHITENH | 10,222 | 3,553 | 22,028 | 21,210 | 10,192 | 3,510 | 22,018 | 21,247 |
| WHITENHP | 36.30 | 12.65 | 77.57 | 75.21 | 36.28 | 12.47 | 76.95 | 75.35 |
| BLACKNH | 4,862 | 4,332 | 1,124 | 2,348 | 4,888 | 4,337 | 1,175 | 2,345 |
| BLACKNHP | 17.27 | 15.43 | 3.96 | 8.33 | 17.40 | 15.41 | 4.11 | 8.32 |
| AIANNH | 283 | 216 | 135 | 172 | 267 | 213 | 137 | 201 |
| AIANNHP | 1.00 | 0.77 | 0.48 | 0.61 | 0.95 | 0.76 | 0.48 | 0.71 |
| ASIANNH | 1,526 | 3,032 | 3,262 | 826 | 1,541 | 3,071 | 3,273 | 806 |
| ASIANNHP | 5.42 | 10.80 | 11.49 | 2.93 | 5.48 | 10.91 | 11.44 | 2.86 |
| HPINH | 25 | 11 | 16 | 14 | 19 | 18 | 19 | 12 |
| HPINHP | 0.09 | 0.04 | 0.06 | 0.05 | 0.07 | 0.06 | 0.07 | 0.04 |
| OTHERNH | 457 | 189 | 224 | 241 | 454 | 208 | 243 | 236 |
| OTHERNHP | 1.62 | 0.67 | 0.79 | 0.85 | 1.62 | 0.74 | 0.85 | 0.84 |
| MLTMNNH | 504 | 458 | 200 | 173 | 480 | 482 | 239 | 172 |
| MLTMNNHP | 1.79 | 1.63 | 0.70 | 0.61 | 1.71 | 1.71 | 0.84 | 0.61 |
| HISP18 | 6,458 | 11,014 | 1,241 | 2,097 | 6,430 | 10,991 | 1,244 | 2,067 |
| HISP18P | 30.88 | 55.50 | 4.89 | 8.89 | 30.80 | 55.36 | 4.89 | 8.77 |
| NONHISP18 | 14,456 | 8,832 | 24,120 | 21,502 | 14,446 | 8,862 | 24,191 | 21,512 |
| NONHISP18P | 69.12 | 44.50 | 95.11 | 91.11 | 69.20 | 44.64 | 95.11 | 91.23 |
| WHITENH18 | 9,131 | 3,062 | 19,682 | 18,839 | 9,111 | 3,055 | 19,679 | 18,850 |
| WHITENH18P | 43.66 | 15.43 | 77.61 | 79.83 | 43.64 | 15.39 | 77.37 | 79.94 |
| BLACKNH18 | 3,309 | 3,027 | 973 | 1,599 | 3,335 | 3,008 | 1,005 | 1,612 |
| BLACKNH18P | 15.82 | 15.25 | 3.84 | 6.78 | 15.98 | 15.15 | 3.95 | 6.84 |
| AIANNH18 | 197 | 154 | 110 | 136 | 185 | 146 | 107 | 151 |
| AIANNH18P | 0.94 | 0.78 | 0.43 | 0.58 | 0.89 | 0.74 | 0.42 | 0.64 |
| ASIANNH18 | 1,170 | 2,135 | 2,989 | 611 | 1,163 | 2,156 | 3,001 | 600 |
| ASIANNH18P | 5.59 | 10.76 | 11.79 | 2.59 | 5.57 | 10.86 | 11.80 | 2.51 |
| HPINH18 | 20 | 11 | 14 | 13 | 17 | 18 | 16 | 9 |
| HPINH18P | 0.10 | 0.06 | 0.06 | 0.06 | 0.08 | 0.09 | 0.06 | 0.04 |
| OTHERNH18 | 326 | 125 | 186 | 178 | 332 | 134 | 203 | 167 |
| OTHERNH18P | 1.56 | 0.63 | 0.73 | 0.75 | 1.59 | 0.67 | 0.80 | 0.71 |
| MLTMNH18 | 303 | 318 | 166 | 126 | 303 | 345 | 180 | 123 |
| MLTMNH18P | 1.45 | 1.60 | 0.65 | 0.53 | 1.45 | 1.74 | 0.71 | 0.52 |

Source: Data from Run A of the $T D A$, U. S. Bureau of the Census, Washington, D.C.
Selected observations for Table 8:
1: SLDU-01 has percentage total $\geq 50 \%$ for TOTALHISPP and BLACKNHP (also HISP18P and BLACKNH18P) for 2010 Census and the TDA run.
2: SLDU-02 has percentages $\geq 50 \%$ for both TOTALHISPP and HISP18P for 2010 Census and the $T D A$ run.
3: SLDU-03 and SLDU-04 each has a percentage $\geq 50 \%$ for both WHITENHP and WHITENH18P for the 2010 Census and the TDA run.

## II.5. EXAMINATION OF RHODE ISLAND's 75 LOWER CHAMBER DISTRICTS

There are 75 districts with one legislator each in Rhode Island's Lower Chamber. Therefore, the IDEAL POPULATION for each State Lower Chamber District (SLDL) is $\frac{1,052,567}{75}=14,034.2$. As with Table 8 for Rhode Island's Upper Chamber, Columns 2-5 of Table 9 give 2010 Census counts and percentages for the State Lower Chamber Districts 01, 02, 03, and 04. Columns 6-9 of Table 9 give corresponding counts and percentages from the same TDA Run A noted in Table 7.

For the 2010 Census counts as well as for the $T D A$ Run A, note the SLDL-03 has a relatively high percentage total for TOTALHISPP and BLACKNHP as well as a high percentage total for HISP18P and BLACKNH18P. Similarly for the 2010 Census counts as well as for the TDA Run A, note that SLDL-01, SLDL-02, and SLDL-04 each has relatively high percentages for both WHITENHP and WHITENH18P.

Unlike in Table 7 for the congressional districts, the corresponding percentages for the demographic groups in the Lower Chamber Districts differ by approximately 1 percentage point. Thus we see more variability for lower levels of geography.

## II.6. EXAMINATION OF THREE CASES PROVIDED BY DOJ

To examine variability for each of the cases provided by DOJ, we proceed for each as we did with Rhode Island. A high level overview of the three cases follows

|  | Jurisdiction | 2010 Census <br> Population | Number of <br> Districts | Number of <br> Blocks Overall | Number of Blocks <br> by Districts |
| :--- | :--- | ---: | :---: | ---: | :--- |
| 1. | Panola County, MS | 34,707 | 5 | 2,180 | $(458 ; 492 ; 413 ; 443 ; 374)$ |
| 2. | Tate County, MS <br> (School District) | 18,823 | 5 | 784 | $(168 ; 204 ; 139 ; 178 ; 95)$ |
| 3. | Tylertown, MS <br> (Walthall County) | 1,609 | 4 | 136 | $(35 ; 42 ; 42 ; 17)$ |

Panola County, MS: In Table 10, the 2010 Census data show, WHITENHP $=48.93 \%$ and BLACKNHP $=48.69 \%$ for the overall county as noted earlier in Table 6. For the same data, and for districts 01,02 , and 05 , we see BLACKNHP values of $63.48 \%, 66.15 \%$, and $50.78 \%$, respectively; for districts 03 and 04 , we see WHITENHP values of $56.97 \%$ and $73.89 \%$, respectively. We see similar corresponding percentages for the results from the TDA.

Tate County (School District), MS: In Table 11, the 2010 Census data show WHITENHP $=68.22 \%$ and BLACKNHP $=28.63 \%$ for the overall county. In addition, the 2010 Census data for districts $01,03,04$, and 05 show WHITENHP values of $86.31 \%, 78.04 \%, 62.02 \%$, and $73.40 \%$, respectively; for district 02, we see BLACKNHP $=54.94 \%$. We see similar corresponding percentages for the results from the TDA.

Tylertown (Walthall County), MS: In Table 12, the 2010 Census data show WHITENHP = $53.45 \%$ and BLACKNHP $=42.20 \%$ for Tylertown (the county seat of Walthall County) overall. For the same data, and for districts 01,02 , and 03 , we see WHITENHP values of $91.60 \%, 53.88 \%$, and $62.92 \%$, respectively; for district 04, we see BLACKNHP $=89.13 \%$. We see less similar corresponding percentages for the results from the TDA for Tylertown than we see for Panola and Tate.

Table 9. Rhode Island Run A of Twenty-five Runs of the TDA for State Lower Chamber Districts (SLDL) 01, 02, 03, and 04 (4 of 75 Districts)
( $C_{T D A}(g)$ counts result from 2020 Census Production Redistricting Data Settings ( $\epsilon=17.14$ for persons) version of $T D A$.)

|  | 2010 Census, SF1 (PL 94-171) (2013) Counts \& Percentages POST-2010 Plan |  |  |  | Counts \& Percentages, 2013 Run A of the $T D A$ |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Demographics |  |  |  |  |  |  |  |  |
| DIST-ID | SLDL-01 | SLDL-02 | SLDL-03 | SLDL-04 | SLDL-01 | SLDL-02 | SLDL-03 | SLDL-04 |
| TOTAL | 13,881 | 13,821 | 13,949 | 13,713 | 14,056 | 13,729 | 13,790 | 13,617 |
| DEV | -153.2 | -213.2 | -85.2 | -321.2 | 21.8 | -305.2 | -244.2 | -417.2 |
| DEVP | -1.10 | -1.54 | -0.61 | -2.34 | 0.15 | -2.22 | -1.77 | -3.06 |
| TOTAL18 | 12,835 | 12,800 | 9,607 | 11,205 | 12,874 | 12,712 | 9,589 | 11,157 |
| TOTALHISP | 1,002 | 1,768 | 5,905 | 1,049 | 1,068 | 1,715 | 5,887 | 1,041 |
| TOTALHISPP | 7.22 | 12.79 | 42.33 | 7.65 | 7.60 | 12.49 | 42.69 | 7.64 |
| TOTALNH | 12,879 | 12,053 | 8,044 | 12,664 | 12,988 | 12,014 | 7,903 | 12,576 |
| TOTALNHP | 92.78 | 87.21 | 57.67 | 92.35 | 92.40 | 87.51 | 57.31 | 92.36 |
| WHITENH | 9,922 | 8,714 | 3,465 | 9,539 | 9,922 | 8,709 | 3,428 | 9,529 |
| WHITENHP | 71.48 | 63.05 | 24.84 | 69.56 | 70.59 | 63.44 | 24.86 | 69.98 |
| BLACKNH | 581 | 1,125 | 3,015 | 1,495 | 605 | 1,124 | 3,001 | 1,480 |
| BLACKNHP | 4.19 | 8.14 | 21.61 | 10.90 | 4.30 | 8.19 | 21.76 | 10.87 |
| AIANNH | 46 | 104 | 189 | 126 | 60 | 89 | 158 | 112 |
| AIANNHP | 0.33 | 0.75 | 1.35 | 0.92 | 0.43 | 0.65 | 1.15 | 0.82 |
| ASIANNH | 2,175 | 1,776 | 794 | 792 | 2,204 | 1,763 | 788 | 782 |
| ASIANNHP | 15.67 | 12.85 | 5.69 | 5.78 | 15.68 | 12.84 | 5.71 | 5.74 |
| HPINH | 12 | 16 | 12 | 1 | 11 | 10 | 8 | 3 |
| HPINHP | 0.09 | 0.12 | 0.09 | 0.01 | 0.08 | 0.07 | 0.06 | 0.02 |
| OTHERNH | 57 | 148 | 257 | 396 | 74 | 143 | 248 | 382 |
| OTHERNHP | 0.41 | 1.07 | 1.84 | 2.89 | 0.53 | 1.04 | 1.80 | 2.81 |
| MLTMNNH | 86 | 170 | 312 | 315 | 112 | 176 | 272 | 288 |
| MLTMNNHP | 0.62 | 1.23 | 2.24 | 2.30 | 0.80 | 1.28 | 1.97 | 2.12 |
| HISP18 | 951 | 1,475 | 3,518 | 693 | 939 | 1,423 | 3,524 | 685 |
| HISP18P | 7.41 | 11.52 | 36.62 | 6.18 | 7.29 | 11.19 | 36.75 | 6.14 |
| NONHISP18 | 11,884 | 11,325 | 6,089 | 10,512 | 11,935 | 11,289 | 6,065 | 10,472 |
| NONHISP18P | 92.59 | 88.48 | 63.38 | 93.82 | 92.71 | 88.81 | 63.25 | 93.86 |
| WHITENH18 | 9,081 | 8,339 | 3,040 | 8,119 | 9,059 | 8,351 | 3,029 | 8,117 |
| WHITENH18P | 70.75 | 65.15 | 31.64 | 72.46 | 70.37 | 65.69 | 31.59 | 72.75 |
| BLACKNH18 | 560 | 972 | 1,971 | 1,144 | 576 | 958 | 1,982 | 1,141 |
| BLACKNH18P | 4.36 | 7.59 | 20.52 | 10.21 | 4.47 | 7.54 | 20.67 | 10.23 |
| AIANNH18 | 45 | 82 | 129 | 101 | 52 | 69 | 114 | 96 |
| AIANNH18P | 0.35 | 0.64 | 1.34 | 0.90 | 0.40 | 0.54 | 1.19 | 0.86 |
| ASIANNH18 | 2,052 | 1,655 | 575 | 635 | 2,082 | 1,649 | 567 | 631 |
| ASIANNH18P | 15.99 | 12.93 | 5.99 | 5.67 | 16.17 | 12.97 | 5.91 | 5.66 |
| HPINH18 | 10 | 14 | 11 | 1 | 11 | 7 | 7 | 3 |
| HPINH18P | 0.08 | 0.11 | 0.11 | 0.01 | 0.09 | 0.06 | 0.07 | 0.03 |
| OTHERNH18 | 51 | 126 | 190 | 280 | 65 | 122 | 193 | 268 |
| OTHERNH18P | 0.40 | 0.98 | 1.98 | 2.50 | 0.50 | 0.96 | 2.01 | 2.40 |
| MLTMNH18 | 85 | 137 | 173 | 232 | 90 | 133 | 173 | 216 |
| MLTMNH18P | 0.66 | 1.07 | 1.80 | 2.07 | 0.70 | 1.05 | 1.80 | 1.94 |

Source: Data from Run A of the TDA, U. S. Bureau of the Census, Washington, D.C.
Selected observations for Table 9:
1: SLDL-01, SLDL-02, and SLDL-04 each has a percentage $\geq 50 \%$ for both WHITENHP and WHITENH18P for 2010 Census and the TDA run.
2: SLDL-03 has a percentage total $\geq 50 \%$ for TOTALHISPP and BLACKNHP jointly, as well as a percentage total $\geq 50 \%$ for HISP18P and BLACKNH18P jointly for 2010 Census and the TDA run.

Table 10. Panola County, MS Run A of Twenty-five Runs of the TDA for County Districts 01, 02, 03, 04, and 05
$\left(C_{T D A}(g)\right.$ counts result from 2020 Census Redistricting Data Production Settings ( $\epsilon=17.14$ for persons) version of TDA.)
2010 Census IDEAL POPULATION $=\frac{34,707}{5}=6,941.4 \quad$ TDA IDEAL POPULATION $=\frac{34,710}{5}=6,942.0$


Source: Data from Run A of the TDA, U.S. Bureau of the Census, Washington, D.C.
Selected observations for Table 10:
1: Panola has WHITENHP $=48.93 \%$ and BLACKNHP $=48.69 \%$ for the 2010 Census; and WHITENHP $=48.94 \%$ and BLACKNHP $=48.65 \%$ for the $T D A$ run. For $18^{+}$population, WHITENH18P $\geq 50.00 \%$ for the 2010 Census and for the TDA run.
2: Districts 01 and 02 each has a percentage $\geq 50 \%$ for BLACKNHP (also BLACKNH18P) for both the 2010 Census and the TDA run. District 05 has a BLACKNHP (also BLACKNH18P) percentage close to $50.00 \%$ for both the 2010 Census and the TDA run.

Table 11. Tate County School Districts (SD), MS Run A of Twenty-five Runs of the TDA for School Districts 01, 02, 03, 04, and 05
( $C_{T D A}(g)$ counts result from 2020 Census Redistricting Data Production Settings ( $\epsilon=17.14$ for persons) version of TDA.)
2010 Census IDEAL POPULATION $=\frac{18,823}{5}=3,764.6 \quad$ TDA IDEAL POPULATION $=\frac{18,813}{5}=3,762.6$


Source: Data from Run A of the TDA, U.S. Bureau of the Census, Washington, D.C.
Selected observations for Table 11:
1: Tate Schools has WHITENHP $=68.22 \%$ and BLACKNHP $=28.63 \%$ for the 2010 Census; and WHITENHP $=68.03 \%$ and BLACKNHP $=28.75 \%$ for the TDA run. Similar results for $18^{+}$ population.
2: School District 02 is the only district with a WHITENHP (also WHITENH18P) percentage lower than $50.00 \%$ in both the 2010 Census and the TDA run.

Table 12. Tylertown (Walthall County), MS Run A of Twenty-five Runs of the TDA for Districts 01, 02, 03, and 04
( $C_{T D A}(g)$ counts result from 2020 Census Redistricting Data Production Settings ( $\epsilon=17.14$ for persons) version of TDA.)
2010 Census IDEAL POPULATION $=\frac{1,609}{4}=402.25 \quad$ TDA IDEAL POPULATION $=\frac{1,614}{4}=403.50$

|  | ```2010 Census, SF1 (PL 94-171) Counts \& Percentages POST-2010 Plan``` |  |  |  |  | Counts \& Percentages Run A of the TDA |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Demographics |  |  |  |  |  |  |  |  |  |  |
| DIST-ID | Tylertown | 01 | 02 | 03 | 04 | Tylertown | 01 | 02 | 03 | 04 |
| TOTAL | 1,609 | 405 | 399 | 391 | 414 | 1,614 | 405 | 417 | 396 | 396 |
| DEV |  | 2.8 | -3.2 | -11.2 | 11.8 |  | 1.5 | 13.5 | -7.5 | -7.5 |
| DEVP |  | 0.68 | -0.81 | -2.88 | 2.84 |  | 0.37 | 3.24 | -1.89 | -1.89 |
| TOTAL18 | 1,233 | 327 | 320 | 313 | 273 | 1,226 | 327 | 327 | 301 | 271 |
| TOTALHISP | 42 | 12 | 7 | 9 | 14 | 42 | 5 | 11 | 18 | 8 |
| TOTALHISPP | 2.61 | 2.96 | 1.75 | 2.30 | 3.38 | 2.60 | 1.23 | 2.64 | 4.55 | 2.02 |
| TOTALNH | 1,567 | 393 | 392 | 382 | 400 | 1,572 | 400 | 406 | 378 | 388 |
| TOTALNHP | 97.39 | 97.04 | 98.25 | 97.70 | 96.62 | 97.40 | 98.77 | 97.36 | 95.45 | 97.98 |
| WHITENH | 860 | 371 | 215 | 246 | 28 | 850 | 370 | 217 | 236 | 27 |
| WHITENHP | 53.45 | 91.60 | 53.88 | 62.92 | 6.76 | 52.66 | 91.36 | 52.04 | 59.60 | 6.82 |
| BLACKNH | 679 | 17 | 174 | 119 | 369 | 682 | 17 | 183 | 128 | 354 |
| BLACKNHP | 42.20 | 4.20 | 43.61 | 30.43 | 89.13 | 42.26 | 4.20 | 43.88 | 32.32 | 89.39 |
| AIANNH | 14 | 5 | 3 | 3 | 3 | 14 | 4 | 2 | 4 | 4 |
| AIANNHP | 0.87 | 1.23 | 0.75 | 0.77 | 0.72 | 0.87 | 0.99 | 0.48 | 1.01 | 1.01 |
| ASIANNH | 12 | 0 | 0 | 12 | 0 | 17 | 6 | 2 | 7 | 2 |
| ASIANNHP | 0.75 | 0.00 | 0.00 | 3.07 | 0.00 | 1.05 | 1.48 | 0.48 | 1.77 | 0.51 |
| HPINH | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 1 | 0 | 0 |
| HPINHP | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.06 | 0.00 | 0.24 | 0.00 | 0.00 |
| OTHERNH | 0 | 0 | 0 | 0 | 0 | 4 | 2 | 0 | 1 | 1 |
| OTHERNHP | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.25 | 0.49 | 0.00 | 0.25 | 0.25 |
| MLTMNNH | 2 | 0 | 0 | 2 | 0 | 4 | 1 | 1 | 2 | 0 |
| MLTMNNHP | 0.12 | 0.00 | 0.00 | 0.51 | 0.00 | 0.25 | 0.25 | 0.24 | 0.51 | 0.00 |
| HISP18 | 27 | 7 | 4 | 8 | 8 | 25 | 2 | 5 | 12 | 6 |
| HISP18P | 2.19 | 2.14 | 1.25 | 2.56 | 2.93 | 2.04 | 0.61 | 1.53 | 3.99 | 2.21 |
| NONHISP18 | 1,206 | 320 | 316 | 305 | 265 | 1,201 | 325 | 322 | 289 | 265 |
| NONHISP18P | 97.81 | 97.86 | 98.75 | 97.44 | 97.07 | 97.96 | 99.39 | 98.47 | 96.01 | 97.79 |
| WHITENH18 | 723 | 302 | 188 | 210 | 23 | 713 | 303 | 191 | 196 | 23 |
| WHITENH18P | 58.64 | 92.35 | 58.75 | 67.09 | 8.42 | 58.16 | 92.66 | 58.41 | 65.12 | 8.49 |
| BLACKNH18 | 462 | 14 | 127 | 81 | 240 | 467 | 10 | 130 | 90 | 237 |
| BLACKNH18P | 37.47 | 4.28 | 39.69 | 25.88 | 87.91 | 38.09 | 3.06 | 39.76 | 29.90 | 87.45 |
| AIANNH18 | 10 | 4 | 1 | 3 | 2 | 10 | 3 | 1 | 2 | 4 |
| AIANNH18P | 0.81 | 1.22 | 0.31 | 0.96 | 0.73 | 0.82 | 0.92 | 0.31 | 0.66 | 1.48 |
| ASIANNH18 | 10 | 0 | 0 | 10 | 0 | 7 | 6 | 0 | 1 | 0 |
| ASIANNH18P | 0.81 | 0.00 | 0.00 | 3.19 | 0.00 | 0.57 | 1.83 | 0.00 | 0.33 | 0.00 |
| HPINH18 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| HPINH18P | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 |
| OTHERNH18 | 0 | 0 | 0 | 0 | 0 | 3 | 2 | 0 | 0 | 1 |
| OTHERNH18P | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.24 | 0.61 | 0.00 | 0.00 | 0.37 |
| MLTMNH18 | 1 | 0 | 0 | 1 | 0 | 1 | 1 | 0 | 0 | 0 |
| MLTMNH18P | 0.08 | 0.00 | 0.00 | 0.32 | 0.00 | 0.08 | 0.31 | 0.00 | 0.00 | 0.00 |

Source: Data from Run A of the TDA, U.S. Bureau of the Census, Washington, D.C.
Selected observations for Table 12:
1: Tylertown has WHITENHP $=53.45 \%$ and BLACKNHP $=42.20 \%$ for the 2010 Census; and WHITENHP $=52.66 \%$ and BLACKNHP $=42.26 \%$ for the $T D A$ run. Similar results hold for $18^{+}$population.
2: District 04 has a BLACKNHP (also BLACKNH18P) percentage $\geq 75 \%$ in both the 2010 Census and the TDA run.
3: District 02 has WHITENHP $=53.88 \%$ in the 2010 Census and WHITENHP $=52.04 \%$ for the TDA run.

## II.7. VARIATION DUE TO THE TopDown ALGORITHM

Definitions of Redistricting Measures of Variation. The measures defined here are all for a specific $\epsilon$. Henceforth, and to simplify notation, we use $S$ for $S W A$ and $T$ for TDA. Let

$$
\begin{aligned}
G & \equiv \text { the number of demographic groups; } \\
C_{S}(g) & \equiv \text { the population of group } g(2010 \text { Census, SF1), for } g=1, \ldots, G \text {; and } \\
C_{T i}(g) & \equiv \text { the population of group } g \text { resulting from the } i^{t h} T D A \text { run, for } i=1, \ldots, 25 .
\end{aligned}
$$

We have the following measures including two types of variation among the $25 T D A$ runs within group $g$ : one relative to $\bar{C}_{T}(g)$ (see below) and another relative to $C_{S}(g)$.
(i) The average population of group $g$ over the $25 T D A$ runs is

$$
\bar{C}_{T}(g) \equiv \frac{C_{T 1}(g)+C_{T 2}(g)+\cdots+C_{T, 25}(g)}{25}
$$

(ii) The variation(1) among the population of group $g$ over the $25 T D A$ runs is

$$
V(1)_{g} \equiv \frac{\left[C_{T 1}(g)-\bar{C}_{T}(g)\right]^{2}+\left[C_{T 2}(g)-\bar{C}_{T}(g)\right]^{2}+\cdots+\left[C_{T, 25}(g)-\bar{C}_{T}(g)\right]^{2}}{25}
$$

(iii) The relative variation(1) among the population of group $g$ over the $25 T D A$ runs is

$$
R V(1)_{g} \equiv \frac{\sqrt{V(1)_{g}}}{\bar{C}_{T}(g)}
$$

(iv) The average relative variation(1) among the population over the $G$ groups (essentially a coefficient of variation) is

$$
A V E R V(1) \equiv \frac{R V(1)_{1}+R V(1)_{2}+\cdots+R V(1)_{G}}{G}
$$

(v) Denote the median relative variation(1) among the population over the $G$ groups by $M E D R V(1)$.
(vi) The variation(2) among the population of group $g$ over the $25 T D A$ runs is

$$
V(2)_{g} \equiv \frac{\left[C_{T 1}(g)-C_{S}(g)\right]^{2}+\left[C_{T 2}(g)-C_{S}(g)\right]^{2}+\cdots+\left[C_{T, 25}(g)-C_{S}(g)\right]^{2}}{25}
$$

(vii) The relative variation(2) among the population of group $g$ over the $25 T D A$ runs is

$$
R V(2)_{g} \equiv \frac{\sqrt{V(2)_{g}}}{C_{S}(g)}
$$

(viii) The average relative variation(2) among the population over the $G$ groups is

$$
A V E R V(2) \equiv \frac{R V(2)_{1}+R V(2)_{2}+\cdots+R V(2)_{G}}{G}
$$

(ix) Denote the median relative variation(2) among the population over the $G$ groups by $M E D R V(2)$.
$V(1)_{g}$ is an empirical variance measuring variation among the $25 T D A$ runs for group $g$; and $V(2)_{g}$ is an empirical mean square error measuring variation and any potential bias (i.e., (bias ${ }^{2}$ ) relative to $C_{S}(g)$ for the $25 T D A$ runs for group $g$.

Tables $7 \mathrm{~V}, 8 \mathrm{~V}, 9 \mathrm{~V}, 10 \mathrm{~V}, 11 \mathrm{~V}$, and 12 V are companion tables for Tables $7,8,9,10,11$, and 12 respectively. The formats among the Tables $7 \mathrm{~V}, 8 \mathrm{~V}, 9 \mathrm{~V}, 10 \mathrm{~V}, 11 \mathrm{~V}$, and 12 V are the same, so we make a few comments about Table 7 V which also hold for the others. For each demographic group $g$ in each district (Rhode Island and CD for Table 7V; SLDU for Table 8V; SLDL for Table 9V; etc.), we provide two sets of three quantities. The first set of quantities gives the average count $\left(C_{T}(g)\right)$ over the $25 T D A$ runs and two associated measures of variation $\left(\sqrt{V(1)_{g}}\right.$ and $\left.R V(1)_{g}\right)$ relative to $\bar{C}_{T}(g)$, while the second set of quantities gives the 2010 Census (swapping) count $\left(C_{S}(g)\right)$ and two associated measures of variation $\left(\sqrt{V(2)_{g}}\right.$ and $\left.R V(2)_{g}\right)$ relative to $C_{S}(g)$. It is worth noting that $\sqrt{V(2)_{g}}$ and $R V(2)_{g}$ are not measures of variability in the swapped data. It is also worth noting that the unit is "persons" for each of the quantities $\bar{C}_{T}(g), \sqrt{V(1)_{g}}, C_{S}(g)$, and $\sqrt{V(2)_{g}}$, while the quantities $R V(1)_{g}$ and $R V(2)_{g}$ are unitless. So for example, we consider the demographic group $g=$ ASIANNH of CD-01 in Table 7V. We observe: $\bar{C}_{T}(g)=17,680$ persons; $\sqrt{V(1)_{g}}=20$ persons; and $R V(1)_{g}=0.001$. We also observe: $C_{S}(g)=17,705$ persons; $\sqrt{V(2)_{g}}=32$ persons; and $R V(2)_{g}=0.002$. The detailed computations for these quantities are illustrated in APPENDIX B. In the tables to follow, a few presented results are rounded. In such cases, especially when there is division, one may not be able to obtain other related presented results exactly.

Selected observations for Table 7V:

1: $\quad R V(1)_{g}$ and $R V(2)_{g}$ are largest for the groups $g=$ HPINH and HPINH18 which have the smallest counts. In general, groups with smaller counts tend to have more relative variation.
2: For a given group $g$, there is a tendency for $R V(2)_{g} \geq R V(1)_{g}$. While this may not be surprising given the definitions of the two measures of variation, this inequality need not hold in all cases, as standardized measures of variation insert different measures of total in the denominator.
3: We observe that $R V(1)_{g}$ and $R V(2)_{g}$ for counts of groups in CD- 02 tend to be larger than for corresponding groups in CD-01. This may be because the districts formed in 2013 resulted in fewer members of minority groups being included in CD-02 than in the corresponding groups in CD-01.

Notice that the computations for $A V E R V(1)$ and $A V E R V(2)$ each only average over the relative variations for the counts in a column. Similarly, $\operatorname{MEDRV}(1)$ and $M E D R V(2)$ are each the median over the relative variations for the counts in a column.

## The Key Empirical Message on Variability

The two measures $\operatorname{AVERV}(\cdot)$ and $M E D R V(\cdot)$ summarize the key single empirical message of this study ( $\epsilon=17.14$ ):

Relative variability in the TDA increases as we consider smaller pieces of geography and population - from state (RI POP $=1,052,567$ ); to Congressional district (RICD IDEAL POP = 526,283.5); to upper chamber district (RI-SLDU IDEAL POP = 27,699.1); to lower chamber district (RI-SLDL IDEAL POP = 14,034.2); to Panola County, MS (DISTRICT IDEAL POP = 6,941.4); to Tate County, MS (SCHOOL DISTRICT IDEAL POP = 3,764.6); and finally to Tylertown (Walthall County), MS (DISTRICT IDEAL POP $=402.25$ ).

To see this empirical evidence, sequentially observe the values for $\operatorname{AVERV}(\cdot)$ and $M E D R V(\cdot)$ on the last two rows of Tables $7 \mathrm{~V} ; 8 \mathrm{~V} ; 9 \mathrm{~V} ; 10 \mathrm{~V} ; 11 \mathrm{~V} ; 12 \mathrm{~V}$. We highlight some of this using DISTRICT IDEAL POPULATION and AVERV (1) in Figure 1.

Figure 1

| Jurisdiction | District | IDEAL POPULATION | AVERV $(1)$ |
| :--- | :---: | ---: | ---: |
| Rhode Island | CD-01 | $526,283.50$ | 0.006 |
| Rhode Island | CD-02 | $526,283.50$ | 0.008 |
|  |  |  |  |
| Rhode Island | SLDU-01 | $27,699.10$ | 0.036 |
| Rhode Island | SLDU-02 | $27,699.10$ | 0.051 |
| Rhode Island | SLDU-03 | $27,699.10$ | 0.045 |
| Rhode Island | SLDU-04 | $27,699.10$ | 0.048 |
|  |  |  |  |
| Rhode Island | SLDL-01 | $14,034.20$ | 0.067 |
| Rhode Island | SLDL-02 | $14,034.20$ | 0.069 |
| Rhode Island | SLDL-03 | $14,034.20$ | 0.057 |
| Rhode Island | SLDL-04 | $14,034.20$ | 0.104 |
|  |  |  |  |
| Panola County, MS | D-01 | $6,941.40$ | 0.362 |
| Panola County, MS | D-02 | $6,941.40$ | 0.353 |
| Panola County, MS | D-03 | $6,941.40$ | 0.344 |
| Panola County, MS | D-04 | $6,941.40$ | 0.369 |
| Panola County, MS | D-05 | $6,941.40$ | 0.280 |
|  |  |  |  |
| Tate County Schools, MS | D-01 | $3,764.60$ | 0.335 |
| Tate County Schools, MS | D-02 | $3,764.60$ | 0.355 |
| Tate County Schools, MS | D-03 | $3,764.60$ | 0.493 |
| Tate County Schools, MS | D-04 | $3,764.60$ | 0.449 |
| Tate County Schools, MS | D-05 | $3,764.60$ | 0.376 |
| Tylertown, MS |  |  |  |
| Tylertown, MS | D-01 | D-02 | 402.25 |
| Tylertown, MS | D-03 | 402.25 | 0.748 |
| Tylertown, MS | D-04 | 402.25 | 0.622 |
|  |  | 402.25 | 1.265 |

Plot of AVERV(1) for IDEAL POPULATION Values Noted Above


Table 7 V . Counts \& Measures of Variation for Rhode Island Twenty-five Runs of the TDA
$\left(C_{T}(g)\right.$ counts result from 2020 Census Redistricting Data Production Settings ( $\epsilon=17.14$ for persons) version of $T D A$.)

|  | (Counts \& Measures of Variation) (2013) |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| DIST-ID | Rhode Island | Rhode Island | CD-01 | CD-01 | CD-02 | CD-02 |
| Demographic (g) | $\bar{C}_{T}(\mathrm{~g})$ | $C_{S}(g)$ | $\bar{C}_{T}(\mathrm{~g})$ | $C_{S}(g)$ | $\bar{C}_{T}(\mathrm{~g})$ | $C_{S}(\mathrm{~g})$ |
|  | $\sqrt{V(1)_{g}}$ | $\sqrt{V(2) g}$ | $\sqrt{V(1) g}$ | $\sqrt{V(2) g}$ | $\sqrt{V(1)_{g}}$ | $\sqrt{V(2) g}$ |
|  | $R V(1)_{g}$ | $R V(2){ }_{g}$ | $R V(1) g_{g}$ | $R V(2) g^{\prime}$ | $R V(1){ }_{g}$ | $R V(2) g_{g}$ |
| TOTAL | 1,052,567 | 1,052,567 | 526,149 | 526,283 | 526,418 | 526,284 |
|  |  | 0 | 43 | 141 | 43 | 141 |
|  | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 |
| TOTAL18 | 828,610 | 828,611 | 412,690 | 412,778 | 415,920 | 415,833 |
|  | 15 | 15 | 85 | 122 | 80 | 119 |
|  | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 |
| TOTALHISP | 130,660 | 130,655 | 76,098 | 76,100 | 54,562 | 54,555 |
|  |  | 14 | 51 | 51 | 57 | 58 |
|  | 0.000 | 0.000 | 0.001 | 0.001 | 0.001 | 0.001 |
| TOTALNH | 921,907 | 921,912 | 450,051 | 450,183 | 471,856 | 471,729 |
|  |  | 14 | 39 | 137 | 48 | 136 |
|  | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 |
| WHITENH | 803,685 | 803,685 | 377,045 | 377,109 | 426,640 | 426,576 |
|  | 8 | 8 | 22 | 68 | 20 | 68 |
|  | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 |
| BLACKNH | 57,929 | 57,927 | 37,641 | 37,627 | 20,288 | 20,300 |
|  | 10 | 10 | 18 | 23 | 19 | 22 |
|  | 0.000 | 0.000 | 0.000 | 0.001 | 0.001 | 0.001 |
| AIANNH | 6,848 | 6,839 | 3,113 | 3,142 | 3,734 | 3,697 |
|  |  | 12 | 19 | 35 | 18 | 41 |
|  | 0.001 | 0.002 | 0.006 | 0.011 | 0.005 | 0.011 |
| ASIANNH | 34,193 | 34,194 | 17,680 | 17,705 | 16,513 | 16,489 |
|  | 9 | 9 | 20 | 32 | 18 | 30 |
|  | 0.000 | 0.000 | 0.001 | 0.002 | 0.001 | 0.002 |
| HPINH | 654 | 655 | 383 | 383 | 271 | 272 |
|  | 11 | 11 | 13 | 13 | 11 | 11 |
|  | 0.016 | 0.016 | 0.033 | 0.033 | 0.041 | 0.042 |
| OTHERNH | 10,292 | 10,296 | 8,503 | 8,492 | 1,789 | 1,804 |
|  |  | 13 | 19 | 22 | 18 | ${ }^{23}$ |
|  | 0.001 | 0.001 | 0.002 | 0.003 | 0.010 | 0.013 |
| MLTMNNH | 8,306 | 8,316 | 5,686 | 5,725 | 2,620 | 2,591 |
|  |  | 21 | 30 | 49 | 29 | 41 |
|  | 0.002 | 0.003 | 0.005 | 0.009 | 0.011 | 0.016 |
| HISP18 | 84,717 | 84,715 | 49,283 | 49,303 | 35,434 | 35,412 |
|  | 14 | 14 | 52 | 56 | 51 | 55 |
|  | 0.000 | 0.000 | 0.001 | 0.001 | 0.001 | 0.002 |
| NONHISP18 | 743,894 | 743,896 | 363,407 | 363,475 | 380,486 | 380,421 |
|  | 12 | 12 | 69 | 96 | 65 | 92 |
|  | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 |
| WHITENH18 | 660,821 | 660,823 | 312,208 | 312,240 | 348,616 | 348,583 |
|  |  | 8 | 29 | 43 | 28 | 44 |
|  | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 |
| BLACKNH18 | 39,485 | 39,485 | 25,403 | 25,402 | 14,081 | 14,083 |
|  |  | 9 | 27 | 27 | 27 | 27 |
|  | 0.000 | 0.000 | 0.001 | 0.001 | 0.002 | 0.002 |
| AIANNH18 | 4,965 | 4,963 | 2,313 | 2,332 | 2,652 | 2,631 |
|  |  | 9 | 23 | 30 | 24 | 32 |
|  | 0.002 | 0.002 | 0.010 | 0.013 | 0.009 | 0.012 |
| ASIANNH18 | 25,330 | 25,333 | 13,257 | 13,276 | 12,073 | 12,057 |
|  |  | 7 | 22 | 29 | 22 | 27 |
|  | 0.000 | 0.000 | 0.002 | 0.002 | 0.002 | 0.002 |
| HPINH18 | 499 | 500 | 303 | 307 | 196 | 193 |
|  |  | 8 | 12 | 12 | 8 | 9 |
|  | 0.017 | 0.017 | 0.039 | 0.040 | 0.043 | 0.047 |
| OTHERNH18 | 7,285 | 7,290 | 6,079 | 6,061 | 1,207 | 1,229 |
|  |  | 12 | 19 | 26 | 18 | 28 |
|  | 0.001 | 0.002 | 0.003 | 0.004 | 0.015 | 0.023 |
| MLTMNH18 | 5,505 | 5,502 | 3,844 | 3,857 | 1,661 | 1,645 |
|  |  | 15 | 29 | 32 | 29 | 33 |
|  | 0.003 | 0.003 | 0.008 | 0.008 | 0.018 | 0.020 |
| AVERV( $\cdot$ ) | 0.002 | 0.002 | 0.006 | 0.006 | 0.008 | 0.010 |
| $M E D R V(\cdot)$ | 0.000 | 0.000 | 0.001 | 0.001 | 0.002 | 0.002 |

Table 8 V . Counts \& Measures of Variation for Rhode Island Twenty-five Runs of the TDA for State Upper Chamber Districts (SLDU) 01, 02, 03, and 04 (4 of 38 Districts, 2013)
$\left(C_{T}(g)\right.$ counts result from 2020 Census Redistricting Data Production Settings ( $\epsilon=17.14$ for persons) version of $T D A$.)

| DIST-ID | (Measures of Variation) (2013) |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | SLDU-01 | SLDU-01 | SLDU-02 | SLDU-02 | SLDU-03 | SLDU-03 | SLDU-04 | SLDU-04 |
|  | $\bar{C}_{T}(g)$ | $C_{S}(g)$ | $\bar{C}_{T}(\mathrm{~g})$ | $C_{S}(g)$ | $\bar{C}_{T}(g)$ | $C_{S}(g)$ | $\bar{C}_{T}(\mathrm{~g})$ | $C_{S}(g)$ |
|  | $\sqrt{V(1)_{g}}$ | $\sqrt{V(2)_{g}}$ | $\sqrt{V(1)_{g}}$ | $\sqrt{V(2)_{g}}$ | $\sqrt{V(1)_{g}}$ | $\sqrt{V(2)_{g}}$ | $\sqrt{V(1)_{g}}$ | $\sqrt{V(2)_{g}}$ |
| Demographic (g) | $R V(1) g$ | $R V(2) g$ | $R V(1){ }_{g}$ | $R V(2) g$ | $R V(1) g$ | $R V(2) g$ | $R V(1){ }_{g}$ | $R V(2) g$ |
| TOTAL | 28,109 | 28,161 | 28,147 | 28,079 | 28,600 | 28,398 | 28,228 | 28,201 |
|  |  |  | 26 | 73 | 26 | 204 | 47 | 54 |
|  | 0.001 | 0.002 | 0.001 | 0.003 | 0.001 | 0.007 | 0.002 | 0.002 |
| TOTAL18 | 20,906 | 20,914 | 19,892 | 19,846 | 25,431 | 25,361 | 23,600 | 23,599 |
|  |  |  |  |  | 31 | 76 | 45 | 45 |
|  | 0.001 | 0.001 | 0.002 | 0.003 | 0.001 | 0.003 | 0.002 | 0.002 |
| TOTALHISP | 10,229 | 10,282 | 16,341 | 16,288 | 1,494 | 1,409 | 3,172 | 3,217 |
|  | 30 | 61 | 28 | 60 | 18 | 87 | 27 | 52 |
|  | 0.003 | 0.006 | 0.002 | 0.004 | 0.012 | 0.062 | 0.008 | 0.016 |
| TOTALNH | 17,881 | 17,879 | 11,806 | 11,791 | 27,106 | 26,989 | 25,055 | 24,984 |
|  | 34 | 34 | 23 | 28 | 28 | 120 | 35 | 79 |
|  | 0.002 | 0.002 | 0.002 | 0.002 | 0.001 | 0.004 | 0.001 | 0.003 |
| WHITENH | 10,212 | 10,222 | 3,517 | 3,553 | 22,023 | 22,028 | 21,260 | 21,210 |
|  |  |  | 14 | 39 | 15 | 15 | 30 | 58 |
|  | 0.002 | 0.002 | 0.004 | 0.011 | 0.001 | 0.001 | 0.001 | 0.003 |
| BLACKNH | 4,886 | 4,862 | 4,339 | 4,332 | 1,153 | 1,124 | 2,346 | 2,348 |
|  | 28 | 37 | 13 | 15 | 15 | 33 | 17 | 18 |
|  | 0.006 | 0.008 | 0.003 | 0.003 | 0.013 | 0.029 | 0.007 | 0.007 |
| AIANNH | 266 | 283 | 209 | 216 | 148 | 135 | 195 | 172 |
|  | 10 | 20 | 9 | 12 | 6 | 15 | 10 | 25 |
|  | 0.039 | 0.072 | 0.044 | 0.054 | 0.042 | 0.108 | 0.052 | 0.148 |
| ASIANNH | 1,554 | 1,526 | 3,055 | 3,032 | 3,269 | 3,262 | 805 | 826 |
|  | 14 | 31 | 12 | 26 | 8 | 10 | 16 | 27 |
|  | 0.009 | 0.021 | 0.004 | 0.008 | 0.002 | 0.003 | 0.020 | 0.032 |
| HPINH | 25 | 25 | 14 | 11 | 15 | 16 | 13 | 14 |
|  | 5 | 5 | 4 | 5 | 4 | 4 | 3 | 3 |
|  | 0.208 | 0.210 | 0.282 | 0.475 | 0.234 | 0.228 | 0.231 | 0.229 |
| OTHERNH | 456 | 457 | 200 | 189 | 253 | 224 | 240 | 241 |
|  | 10 | 10 | 7 | 13 | 9 | 30 | 10 | 10 |
|  | 0.022 | 0.022 | 0.033 | 0.068 | 0.035 | 0.136 | 0.041 | 0.041 |
| MLTMNNH | 482 | 504 | 473 | 458 | 244 | 200 | 197 | 173 |
|  | 13 | 26 | 17 | 22 | 15 | 47 | 12 | 27 |
|  | 0.028 | 0.052 | 0.035 | 0.049 | 0.062 | 0.235 | 0.060 | 0.156 |
| HISP18 | 6,431 | 6,458 | 11,043 | 11,014 | 1,249 | 1,241 | 2,068 | 2,097 |
|  | 27 |  | 31 |  | 17 | 19 | 24 | 38 |
|  | 0.004 | 0.006 | 0.003 | 0.004 | 0.014 | 0.015 | 0.012 | 0.018 |
| NONHISP18 | 14,474 | 14,456 | 8,849 | 8,832 | 24,182 | 24,120 | 21,532 | 21,502 |
|  | 30 | 35 | 19 | 26 | 29 | 68 | 30 | 43 |
|  | 0.002 | 0.002 | 0.002 | 0.003 | 0.001 | 0.003 | 0.001 | 0.002 |
| WHITENH18 | 9,137 | 9,131 | 3,046 | 3,062 | 19,682 | 19,682 | 18,864 | 18,839 |
|  | 16 | 17 | 11 | 19 | 13 | 13 | 25 | 35 |
|  | 0.002 | 0.002 | 0.004 | 0.006 | 0.001 | 0.001 | 0.001 | 0.002 |
| BLACKNH18 | 3,312 | 3,309 | 3,030 | 3,027 | 986 | 973 | 1,598 | 1,599 |
|  |  |  |  | 15 | 13 | 18 | 18 | 18 |
|  | 0.008 | 0.008 | 0.005 | 0.005 | 0.013 | 0.019 | 0.011 | 0.011 |
| AIANNH18 | 187 | 197 | 150 | 154 | 117 | 110 | 146 | 136 |
|  | 9 | 14 | 10 | 11 | 10 | 12 | 10 | 14 |
|  | 0.050 | 0.069 | 0.068 | 0.071 | 0.085 | 0.111 | 0.067 | 0.101 |
| ASIANNH18 | 1,183 | 1,170 | 2,149 | 2,135 | 3,001 | 2,989 | 600 | 611 |
|  | 14 | 19 | 12 | 19 | 9 | 15 | 15 | 19 |
|  | 0.012 | 0.016 | 0.006 | 0.009 | 0.003 | 0.005 | 0.025 | 0.030 |
| HPINH18 | 20 | 20 | 11 | 11 | 13 | 14 | 10 | 13 |
|  | 5 | 5 | 5 | 5 | 3 | 4 | 3 | 4 |
|  | 0.237 | 0.237 | 0.430 | 0.445 | 0.262 | 0.254 | 0.297 | 0.313 |
| OTHERNH18 | 327 | 326 | 134 | 125 | 198 | 186 | 175 | 178 |
|  | 12 | 12 | 7 | 11 | 9 | 15 | 8 | 9 |
|  | 0.038 | 0.038 | 0.053 | 0.090 | 0.045 | 0.080 | 0.047 | 0.050 |
| MLTMNH18 | 307 | 303 | 329 | 318 | 185 | 166 | 139 | 126 |
|  | 14 | 15 | 14 | 18 | 14 | 23 | 11 | 17 |
|  | 0.045 | 0.048 | 0.043 | 0.055 | 0.076 | 0.140 | 0.079 | 0.136 |
| $A V E R V(\cdot)$ | 0.036 | 0.041 | 0.051 | 0.068 | 0.045 | 0.072 | 0.048 | 0.065 |
| MEDRV( $\cdot$ ) | 0.008 | 0.012 | 0.004 | 0.009 | 0.013 | 0.024 | 0.016 | 0.024 |

Source: Data from 25 Runs of the TDA, U. S. Bureau of the Census, Washington, D.C.

Table 9V. Counts \& Measures of Variation for Rhode Island Twenty-five Runs of the TDA for State Lower Chamber Districts (SLDL) 01, 02, 03, and 04 (4 of 75 Districts, 2013)
$\left(C_{T}(g)\right.$ counts result from 2020 Census Redistricting Data Production Settings ( $\epsilon=17.14$ for persons) version of $T D A$.)

| DIST-ID | (Measures of Variation) (2013) |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | SLDL-01 | SLDL-01 | SLDL-02 | SLDL-02 | SLDL-03 | SLDL-03 | SLDL-04 | SLDL-04 |
|  | $\bar{C}_{T}(\mathrm{~g})$ | $C_{S}(g)$ | $\bar{C}_{T}(\mathrm{~g})$ | $C_{S}(g)$ | $\bar{C}_{T}(\mathrm{~g})$ | $C_{S}(\mathrm{~g})$ | $\bar{C}_{T}(\mathrm{~g})$ | $C_{S}(\mathrm{~g})$ |
|  | $\sqrt{V(1)_{g}}$ | $\sqrt{V(2)_{g}}$ | $\sqrt{V(1)_{g}}$ | $\sqrt{V(2) g}$ | $\sqrt{V(1)_{g}}$ | $\sqrt{V(2)_{g}}$ | $\sqrt{V(1) g}$ | $\sqrt{V(2)_{g}}$ |
| Demographic (g) | $R V(1) g^{\prime}$ | $R V(2) g_{g}$ | $R V(1){ }_{g}$ | $R V(2) g$ | $R V(1) g_{g}$ | $R V(2) g$ | $R V(1) g^{\prime}$ | $R V(2){ }_{g}$ |
| TOTAL | 14,098 | 13,881 | 13,721 | 13,821 | 13,776 | 13,949 | 13,594 | 13,713 |
|  |  | 237 |  | 109 |  | 180 | 37 | 125 |
|  | 0.005 | 0.016 | 0.003 | 0.008 | 0.003 | 0.013 | 0.003 | 0.009 |
| TOTAL18 | 12,939 | 12,835 | 12,648 | 12,800 | 9,562 | 9,607 | 11,127 | 11,205 |
|  | 54 | 118 | 36 | 156 | 34 | 56 | 32 | 84 |
|  | 0.004 | 0.009 | 0.003 | 0.012 | 0.004 | 0.006 | 0.003 | 0.008 |
| TOTALHISP | 1,098 | 1,002 | 1,723 | 1,768 | 5,844 | 5,905 | 1,023 | 1,049 |
|  |  | 100 | 27 | 53 | 35 | 70 | 11 | 28 |
|  | 0.028 | 0.100 | 0.016 | 0.030 | 0.006 | 0.012 | 0.011 | 0.027 |
| TOTALNH | 13,001 | 12,879 | 11,999 | 12,053 | 7,932 | 8,044 | 12,571 | 12,664 |
|  | 45 | 130 | 32 | 63 | 40 | 119 | 35 | 99 |
|  | 0.003 | 0.010 | 0.003 | 0.005 | 0.005 | 0.015 | 0.003 | 0.008 |
| WHITENH | 9,916 | 9,922 | 8,693 | 8,714 | 3,452 | 3,465 | 9,546 | 9,539 |
|  | 33 | 33 | 22 | 30 | 20 | 24 | 26 | 27 |
|  | 0.003 | 0.003 | 0.002 | 0.003 | 0.006 | 0.007 | 0.003 | 0.003 |
| BLACKNH | 616 | 581 | 1,133 | 1,125 | 2,979 | 3,015 | 1,472 | 1,495 |
|  | 13 | 37 | 19 | 20 | 22 | 42 | 9 | 25 |
|  | 0.021 | 0.064 | 0.016 | 0.018 | 0.007 | 0.014 | 0.006 | 0.017 |
| AIANNH | 62 | 46 | 98 | 104 | 156 | 189 | 112 | 126 |
|  | 6 | 17 | 7 | 9 | 10 | 34 | 5 | 15 |
|  | 0.095 | 0.369 | 0.073 | 0.088 | 0.061 | 0.180 | 0.048 | 0.121 |
| ASIANNH | 2,186 | 2,175 | 1,767 | 1,776 | 802 | 794 | 785 | 792 |
|  |  |  | 22 | 24 | 14 | 16 | 14 | 15 |
|  | 0.008 | 0.010 | 0.012 | 0.013 | 0.017 | 0.020 | 0.017 | 0.019 |
| HPINH | 10 | 12 | 12 | 16 | 13 | 12 | 3 | 1 |
|  | 3 | 3 | 4 | 6 | 4 | 4 | 3 | 3 |
|  | 0.257 | 0.261 | 0.384 | 0.394 | 0.324 | 0.349 | 0.801 | 3.429 |
| OTHERNH | 82 | 57 | 136 | 148 | 248 | 257 | 382 | 396 |
|  | 10 | 27 | 8 | 14 | 10 | 13 | 6 | 15 |
|  | 0.121 | 0.477 | 0.061 | 0.098 | 0.038 | 0.051 | 0.016 | 0.039 |
| MLTMNNH | 128 | 86 | 159 | 170 | 281 | 312 | 271 | 315 |
|  | 15 | 45 | 10 | 15 | 12 | 33 | 11 | 46 |
|  | 0.113 | 0.518 | 0.065 | 0.088 | 0.042 | 0.106 | 0.012 | 0.145 |
| HISP18 | 988 | 951 | 1,393 | 1,475 | 3,514 | 3,518 | 671 | 693 |
|  | 27 | 16 | 24 | 86 | 27 | 27 | 17 | 28 |
|  | 0.028 | 0.019 | 0.017 | 0.058 | 0.008 | 0.008 | 0.025 | 0.040 |
| NONHISP18 | 11,951 | 11,884 | 11,256 | 11,325 | 6,048 | 6,089 | 10,456 | 10,512 |
|  | 38 | 77 | 33 | 77 | 29 | 50 | 33 | 65 |
|  | 0.003 | 0.006 | 0.003 | 0.007 | 0.005 | 0.008 | 0.003 | 0.006 |
| WHITENH18 | 9,080 | 9,081 | 8,317 | 8,339 | 3,040 | 3,040 | 8,127 | 8,119 |
|  | 25 | 25 | 23 | 32 | 17 | 17 | 19 | 21 |
|  | 0.003 | 0.003 | 0.003 | 0.004 | 0.006 | 0.006 | 0.002 | 0.003 |
| BLACKNH18 | 575 | 560 | 967 | 972 | 1,950 | 1,971 | 1,139 | 1,144 |
|  | 12 | 19 | 15 | 16 | 19 | 28 | 10 | 11 |
|  | 0.020 | 0.034 | 0.016 | 0.016 | 0.010 | 0.014 | 0.009 | 0.009 |
| AIANNH18 | 53 | 45 | 77 | 82 | 110 | 129 | 93 | 101 |
|  | 6 | 10 | 8 | 10 | 10 | 21 | 6 | 11 |
|  | 0.104 | 0.217 | 0.109 | 0.120 | 0.087 | 0.163 | 0.070 | 0.105 |
| ASIANNH18 | 2,063 | 2,052 | 1,646 | 1,655 | 582 | 575 | 631 | 635 |
|  | 18 | 21 | 19 | 21 | 11 | 13 | 11 | 12 |
|  | 0.009 | 0.010 | 0.011 | 0.012 | 0.019 | 0.023 | 0.017 | 0.018 |
| HPINH18 | 9 | 10 | 9 | 14 | 10 | 11 | 3 | 1 |
|  | 2 | 3 | 4 | 6 | 4 | 4 | 2 | 3 |
|  | 0.260 | 0.251 | 0.462 | 0.454 | 0.362 | 0.346 | 0.906 | 2.742 |
| OTHERNH18 | 65 | 51 | 112 | 126 | 183 | 190 | 268 | 280 |
|  | 9 | 17 | 9 | 16 | 10 | 12 | 8 | 15 |
|  | 0.132 | 0.327 | 0.076 | 0.127 | 0.053 | 0.063 | 0.031 | 0.053 |
| MLTMNH18 | 105 | 85 | 126 | 137 | 172 | 173 | 196 | 232 |
|  | 14 | 24 | 7 | 13 | 15 | 15 | 12 | 38 |
|  | 0.130 | 0.284 | 0.054 | 0.092 | 0.086 | 0.085 | 0.059 | 0.164 |
| AVERV( $\cdot$ ) | 0.067 | 0.151 | 0.069 | 0.082 | 0.057 | 0.074 | 0.104 | 0.348 |
| $M E D R V(\cdot)$ | 0.024 | 0.056 | 0.016 | 0.024 | 0.013 | 0.018 | 0.016 | 0.023 |

[^2]$\left(C_{T}(g)\right.$ counts result from 2020 Census Redistricting Data Production Settings ( $\epsilon=17.14$ for persons) version of $T D A$.)


Source: Data from 25 Runs of the TDA, U. S. Bureau of the Census, Washington, D.C.

Table 11V. Counts \& Measures of Variation for Tate County School Districts, MS Twenty-five Runs of the TDA for County Districts 01, 02, 03, 04, 05

| DIST-ID | (Measures of Variation) |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Tate Schools |  | $\begin{array}{ll} \hline 01 & 01 \\ \hline \end{array}$ |  | 02 | 02 | 0303 |  | 04 | 04 | 05 | 05 |
|  | $\bar{C}_{T}(g)$ | $C_{S}(g)$ | $\bar{C}_{T}(\mathrm{~g})$ | $C_{S}(g)$ | $\bar{C}_{T}(g)$ | $C_{S}(g)$ | $\bar{C}_{T}(\mathrm{~g})$ | $C_{S}(g)$ | $\bar{C}_{T}(g)$ | $C_{S}(g)$ | $\bar{C}_{T}(g)$ | $C_{S}(g)$ |
|  | $\sqrt{V(1) g}$ | $\sqrt{V(2) g}$ | $\sqrt{V(1)_{g}}$ | $\sqrt{V(2)_{g}}$ | $\sqrt{V(1)_{g}}$ | $\sqrt{V(2)_{g}}$ | $\sqrt{V(1)_{g}}$ | $\sqrt{V(2)_{g}}$ | $\sqrt{V(1)_{g}}$ | $\sqrt{V(2)_{g}}$ | $\sqrt{V(1) g}$ | $\sqrt{V(2) g}$ |
| Demographic (g) | $R V(1) g$ | $R V(2) g^{\prime}$ | $R V(1){ }_{g}$ | $R V(2) g^{\prime}$ | $R V(1) g$ | $R V(2) g^{\prime}$ | $R V(1) g_{g}$ | $R V(2) g^{\prime}$ | $R V(1) g$ | $R V(2) g^{\prime}$ | $R V(1) g_{g}$ | $R V(2) g$ |
| TOTAL | 18,815 | 18,823 | 3,916 | 3,914 | 3,885 | 3,893 | 3,644 | 3,665 | 3,714 | 3,697 | 3,657 | 3,654 |
|  |  | 20 | 22 | 22 | 21 | 23 | 20 | 30 | 26 | 31 | 16 | 16 |
|  | 0.001 | 0.001 | 0.006 | 0.006 | 0.005 | 0.006 | 0.006 | 0.008 | 0.007 | 0.008 | 0.004 | 0.004 |
| TOTAL18 | 13,892 | 13,893 | 2,776 | 2,780 | 2,833 | 2,826 | 2,789 | 2,799 | 2,766 | 2,755 | 2,728 | 2,733 |
|  |  | 17 | 20 | 21 | 19 | 20 | 14 | 17 | 23 | 26 | 13 | 14 |
|  | 0.001 | 0.001 | 0.007 | 0.007 | 0.007 | 0.007 | 0.005 | 0.006 | 0.008 | 0.009 | 0.005 | 0.005 |
| TOTALHISP | 423 | 399 | 95 | 87 | 64 | 63 | 106 | 110 | 51 | 32 | 106 | 107 |
|  | 9 | 26 | 6 | 10 | 4 | 4 | 8 | 9 | 6 | 20 | 8 | 8 |
|  | 0.021 | 0.064 | 0.066 | 0.118 | 0.063 | 0.066 | 0.073 | 0.078 | 0.119 | 0.631 | 0.072 | 0.071 |
| TOTALNH | 18,392 | 18,424 | 3,821 | 3,827 | 3,821 | 3,830 | 3,537 | 3,555 | 3,663 | 3,665 | 3,551 | 3,547 |
|  |  | 37 | 22 | 23 | 21 | 23 | 19 | 26 | 24 | 24 | 18 | 18 |
|  | 0.001 | 0.002 | 0.006 | 0.006 | 0.005 | 0.006 | 0.005 | 0.007 | 0.007 | 0.007 | 0.005 | 0.005 |
| WHITENH | 12,805 | 12,841 | 3,387 | 3,378 | 1,613 | 1,628 | 2,833 | 2,860 | 2,276 | 2,293 | 2,696 | 2,682 |
|  |  | 39 | 14 | 17 | 14 | 21 | 14 | 30 | 20 | 26 | 12 | 19 |
|  | 0.001 | 0.003 | 0.004 | 0.005 | 0.009 | 0.013 | 0.005 | 0.011 | 0.009 | 0.011 | 0.005 | 0.007 |
| BLACKNH | 5,394 | 5,389 | 373 | 400 | 2,158 | 2,139 | 678 | 666 | 1,363 | 1,349 | 822 | 835 |
|  | 11 | 12 | 12 | 30 | 10 | 21 | 11 | 16 | 14 | 20 | 15 | 20 |
|  | 0.002 | 0.002 | 0.033 | 0.074 | 0.004 | 0.010 | 0.016 | 0.024 | 0.010 | 0.015 | 0.018 | 0.024 |
| AIANNH | 101 | 103 | 35 | 32 | 23 | 26 | 15 | 19 | 10 | 11 | 17 | 15 |
|  | 6 | 6 | 4 | 5 | 3 | 4 | 3 | 5 | 3 | 3 | 4 | 4 |
|  | 0.056 | 0.059 | 0.110 | 0.160 | 0.142 | 0.167 | 0.181 | 0.265 | 0.294 | 0.284 | 0.207 | 0.274 |
| ASIANNH | 50 | 47 | 18 | 14 | 13 | 16 | 7 | 6 | 7 | 7 | 5 | 4 |
|  | 4 | 5 | 4 | 6 | 3 | 4 | 2 | 2 | 2 | 2 | 2 | 3 |
|  | 0.084 | 0.105 | 0.245 | 0.403 | 0.197 | 0.242 | 0.296 | 0.337 | 0.255 | 0.254 | 0.414 | 0.642 |
| HPINH | 3 | 3 | 1 | 2 | 1 | 0 | 0 | 0 | 0 | 0 | 1 | 1 |
|  | 3 | 3 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 |
|  | 0.764 | 0.841 | 0.952 | 0.714 | 1.894 | Inf | 2.114 | Inf | 1.710 | Inf | 1.362 | 0.980 |
| OTHERNH | 10 | 9 | 2 | 1 | 5 | 5 | 1 | 1 | 1 | 1 | 2 | 1 |
|  | 5 | 5 | 2 | 2 | 2 | 2 | 1 | 1 | 1 | 1 | 2 | 2 |
|  | 0.487 | 0.571 | 0.938 | 1.766 | 0.509 | 0.470 | 1.138 | 1.281 | 1.138 | 1.281 | 0.981 | 1.887 |
| MLTMNNH | 30 | 32 | 4 | 0 | 9 | 16 | 4 | 3 | 5 | 4 | 8 | 9 |
|  | 7 | 7 | 2 | 5 | 3 | 8 | 2 | 2 | 3 | 3 | 3 | 3 |
|  | 0.225 | 0.220 | 0.583 | Inf | 0.355 | 0.475 | 0.563 | 0.727 | 0.472 | 0.709 | 0.404 | 0.375 |
| HISP18 | 227 | 215 | 48 | 47 | 39 | 34 | 59 | 63 | 24 | 16 | 57 | 55 |
|  | 9 | 15 | 6 | 6 | 4 | 6 | 5 | 7 | 5 | 10 | 6 | 6 |
|  | 0.041 | 0.070 | 0.117 | 0.123 | 0.113 | 0.191 | 0.091 | 0.110 | 0.197 | 0.603 | 0.105 | 0.112 |
| NONHISP18 | 13,666 | 13,678 | 2,728 | 2,733 | 2,795 | 2,792 | 2,731 | 2,736 | 2,742 | 2,739 | 2,671 | 2,678 |
|  |  | 19 | 20 | 21 | 18 | 19 | 12 | 13 | 22 | 22 | 15 | 17 |
|  | 0.001 | 0.001 | 0.007 | 0.008 | 0.007 | 0.007 | 0.004 | 0.005 | 0.008 | 0.008 | 0.006 | 0.006 |
| WHITENH18 | 9,735 | 9,747 | 2,444 | 2,438 | 1,274 | 1,278 | 2,204 | 2,219 | 1,747 | 1,755 | 2,066 | 2,057 |
|  |  |  |  | 13 |  | 11 |  | 18 | 15 | 17 | 12 | 15 |
|  | 0.001 | 0.002 | 0.005 | 0.005 | 0.008 | 0.009 | 0.004 | 0.008 | 0.009 | 0.010 | 0.006 | 0.007 |
| BLACKNH18 | 3,793 | 3,790 | 244 | 261 | 1,482 | 1,471 | 506 | 498 | 979 | 965 | 582 | 595 |
|  |  |  | 11 | 20 |  | 15 | 8 | 12 | 12 | 18 | 12 | 18 |
|  | 0.003 | 0.003 | 0.045 | 0.078 | 0.006 | 0.010 | 0.016 | 0.024 | 0.013 | 0.019 | 0.021 | 0.030 |
| AIANNH18 | 74 | 79 | 23 | 23 | 19 | 21 | 12 | 13 | 7 | 9 | 12 | 13 |
|  | 5 | 7 | 5 | 5 | 3 | 4 | 3 | 3 | 3 | 3 | 3 | 3 |
|  | 0.071 | 0.094 | 0.204 | 0.204 | 0.161 | 0.176 | 0.223 | 0.218 | 0.400 | 0.379 | 0.238 | 0.232 |
| ASIANNH18 |  | 35 | 12 | 8 | 11 | 13 | 5 | 4 | 5 | 6 | 4 | 4 |
|  | 4 |  | 4 | 5 | 3 | 4 | 2 | 2 | 2 | 2 | 2 | 2 |
|  | 0.098 | 0.114 | 0.294 | 0.667 | 0.284 | 0.282 | 0.437 | 0.561 | 0.422 | 0.389 | 0.447 | 0.427 |
| HPINH18 | 3 | 3 | 1 | 2 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 1 |
|  | 3 | 3 | 1 | 2 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 |
|  | 0.919 | 0.894 | 1.106 | 0.775 | 2.051 | Inf | 2.828 | Inf | 1.710 | Inf | 1.491 | 0.980 |
| OTHERNH18 | 5 | 4 | 1 | 1 | 2 | 1 | 1 | 1 | 0 | 0 | 1 | 1 |
|  | 3 | 3 | 1 | 1 | 2 | 2 | 1 | 1 | 1 | 1 | 1 | 1 |
|  | 0.580 | 0.820 | 1.061 | 0.980 | 0.716 | 1.980 | 1.155 | 0.800 | 1.679 | Inf | 1.260 | 1.311 |
| MLTMNH18 |  |  | 3 | 0 | 5 | 8 | 2 | 1 | 3 | 4 | 6 | 7 |
|  | 6 |  | 3 | 4 | 3 | 4 | 2 | 2 | 2 | 2 | 3 | 3 |
|  | 0.322 | 0.310 | 0.921 | Inf | 0.560 | 0.505 | 0.707 | 2.200 | 0.505 | 0.453 | 0.464 | 0.423 |
| AVERV( $\cdot$ ) | 0.184 | 0.209 | 0.335 | Inf | 0.355 | Inf | 0.493 | Inf | 0.449 | Inf | 0.376 | 0.390 |
| MEDRV( $)$ | 0.049 | 0.067 | 0.113 | 0.141 | 0.128 | 0.171 | 0.136 | 0.164 | 0.226 | 0.332 | 0.156 | 0.172 |

Source: Data from 25 Runs of the TDA, U. S. Bureau of the Census, Washington, D.C.

Table 12V. Counts \& Measures of Variation for Tylertown (Walthall County), MS Twenty-five Runs of the TDA for County Districts 01, 02, 03, 04
( $C_{T}(g)$ counts result from 2020 Census Production Redistricting Data Settings ( $\epsilon=17.14$ for persons) version of TDA.)

| DIST-ID | (Measures of Variation) |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Tylertown |  | 01 | 01 | 02 | 02 | 03 | 03 | 04 | 04 |
|  | $\bar{C}_{T}(\mathrm{~g})$ | $C_{S}(g)$ | $\bar{C}_{T}(g)$ | $C_{S}(g)$ | $\bar{C}_{T}(g)$ | $C_{S}(g)$ | $\bar{C}_{T}(g)$ | $C_{S}(g)$ | $\bar{C}_{T}(\mathrm{~g})$ | $C_{S}(g)$ |
|  | $\sqrt{V(1) g}$ | $\sqrt{V(2)_{g}}$ | $\sqrt{V(1) g}$ | $\sqrt{V(2)_{g}}$ | $\sqrt{V(1) g}$ | $\sqrt{V(2)_{g}}$ | $\sqrt{V(1)_{g}}$ | $\sqrt{V(2) g}$ | $\sqrt{V(1)_{g}}$ | $\sqrt{V(2)_{g}}$ |
| Demographic (g) | $R V(1) g_{g}$ | $R V(2){ }_{g}$ | $R V(1) g_{g}$ | $R V(2) g^{\prime}$ | $R V(1) g$ | $R V(2) g$ | $R V(1) g_{g}$ | $R V(2) g^{\prime}$ | $R V(1)_{g}$ | $R V(2) g_{g}$ |
| TOTAL | 1,611 | 1,609 | 402 | 405 | 396 | 399 | 417 | 391 | 397 | 414 |
|  | , |  | 14 | 14 | 13 | 13 | 16 | 30 | 8 | 19 |
|  | 0.001 | 0.002 | 0.035 | 0.036 | 0.033 | 0.033 | 0.037 | 0.077 | 0.021 | 0.047 |
| TOTAL18 | 1,239 | 1,233 | 327 | 327 | 320 | 320 | 326 | 313 | 266 | 273 |
|  | 6 | 9 | 10 | 10 | 12 | 12 | 15 | 20 | 10 | 12 |
|  | 0.005 | 0.007 | 0.031 | 0.031 | 0.038 | 0.038 | 0.045 | 0.062 | 0.036 | 0.043 |
| TOTALHISP | 42 | 42 | 9 | 12 | 10 | 7 | 18 | 9 | 5 | 14 |
|  | 5 | 5 | 3 | 4 | 3 | 4 | 5 | 10 | 3 | 10 |
|  | 0.124 | 0.123 | 0.336 | 0.346 | 0.314 | 0.594 | 0.303 | 1.138 | 0.560 | 0.682 |
| TOTALNH | 1,570 | 1,567 | 393 | 393 | 386 | 392 | 399 | 382 | 392 | 400 |
|  |  |  | 13 | 13 | 12 | 13 | 15 | 23 | 9 | 12 |
|  | 0.003 | 0.004 | 0.034 | 0.034 | 0.031 | 0.034 | 0.039 | 0.060 | 0.022 | 0.030 |
| WHITENH | 848 | 860 | 360 | 371 | 205 | 215 | 249 | 246 | 33 | 28 |
|  | 3 | 13 | 10 | 15 | 9 | 14 | 12 | 13 | 7 | 9 |
|  | 0.003 | 0.015 | 0.028 | 0.039 | 0.045 | 0.063 | 0.050 | 0.052 | 0.216 | 0.307 |
| BLACKNH | 681 | 679 | 20 | 17 | 174 | 174 | 132 | 119 | 356 | 369 |
|  | 3 | 4 | 6 | 7 | 7 | 7 | 7 | 15 | 8 | 15 |
|  | 0.005 | 0.006 | 0.320 | 0.417 | 0.041 | 0.041 | 0.056 | 0.123 | 0.022 | 0.042 |
| AIANNH | 15 | 14 | 5 | 5 | 4 | 3 | 5 | 3 | 1 | 3 |
|  | 4 | 4 | 3 | 3 | 2 | 2 | 2 | 3 | 1 | 2 |
|  | 0.257 | 0.293 | 0.510 | 0.514 | 0.545 | 0.673 | 0.439 | 1.056 | 1.041 | 0.721 |
| ASIANNH | 15 | 12 | 5 | 0 | 1 | 0 | 8 | 12 | 1 | 0 |
|  | 3 | 4 | 3 | 6 | 1 | 2 | 3 | 5 | 1 | 1 |
|  | 0.201 | 0.348 | 0.543 | Inf | 1.025 | Inf | 0.314 | 0.383 | 1.344 | Inf |
| HPINH | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
|  | 1 | 1 | 0 | 0 | 0 | 1 | 1 | 1 | 0 | 0 |
|  | 1.061 | Inf | 2.708 | Inf | 1.458 | Inf | 1.827 | Inf | 4.899 | Inf |
| OTHERNH | 2 | 0 | 1 | 0 | 1 | 0 | 1 | 0 | 0 | 0 |
|  | 2 | 3 | 1 | 1 | 1 | 1 | 1 | 2 | 0 | 0 |
|  | 0.731 | Inf | 1.550 | Inf | 1.033 | Inf | 1.406 | Inf | 2.291 | Inf |
| MLTMNNH | 7 | 2 | 2 | 0 | 2 | 0 | 3 | 2 | 1 | 0 |
|  | 4 | 7 | 1 | 2 | 2 | 3 | 2 | 2 | 1 | 1 |
|  | 0.534 | 3.367 | 0.695 | Inf | 1.016 | Inf | 0.699 | 1.140 | 1.356 | Inf |
| HISP18 | 28 | 27 | 6 | 7 | 6 | 4 | 13 | 8 | 3 | 8 |
|  | 5 | 5 | 2 | 3 | 3 | 3 | 5 | 8 | 2 | 5 |
|  | 0.185 | 0.200 | 0.397 | 0.369 | 0.547 | 0.846 | 0.399 | 0.940 | 0.557 | 0.629 |
| NONHISP18 | 1,210 | 1,206 | 321 | 320 | 314 | 316 | 313 | 305 | 263 | 265 |
|  | 7 | 8 | 9 | 9 | 11 | 11 | 14 | 16 | 9 | 9 |
|  | 0.006 | 0.007 | 0.029 | 0.029 | 0.035 | 0.036 | 0.044 | 0.051 | 0.035 | 0.036 |
| WHITENH18 | 716 | 723 | 297 | 302 | 183 | 188 | 211 | 210 | 25 | 23 |
|  | 4 | 8 | 8 | 10 | 8 | 10 | 11 | 11 | 7 | 7 |
|  | 0.006 | 0.012 | 0.028 | 0.032 | 0.045 | 0.052 | 0.051 | 0.052 | 0.267 | 0.304 |
| BLACKNH18 | 463 | 462 | 13 | 14 | 126 | 127 | 88 | 81 | 236 | 240 |
|  |  | 5 | 4 | 4 | 6 | 6 | 6 | 9 | 6 | 8 |
|  | 0.009 | 0.010 | 0.289 | 0.279 | 0.051 | 0.051 | 0.068 | 0.114 | 0.027 | 0.032 |
| AIANNH18 | 11 | 10 | 4 | 4 | 3 | 1 | 4 | 3 | 1 | 2 |
|  | 4 | 4 | 2 | 2 | 2 | 3 | 2 | 2 | 1 | 2 |
|  | 0.365 | 0.439 | 0.517 | 0.522 | 0.734 | 2.683 | 0.658 | 0.792 | 1.169 | 0.781 |
| ASIANNH18 | 13 | 10 | 4 | 0 | 1 | 0 | 7 | 10 | 0 | 0 |
|  |  | 4 | 2 | 5 | 1 | 1 | 3 | 4 | 1 | 1 |
|  | 0.259 | 0.430 | 0.547 | Inf | 1.541 | Inf | 0.390 | 0.387 | 1.705 | Inf |
| HPINH18 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
|  | 1 | 1 | 0 | 0 | 0 | 1 | 1 | 1 | 0 | 0 |
|  | 1.228 | Inf | 3.391 | Inf | 1.604 | Inf | 1.895 | Inf | 4.899 | Inf |
| OTHERNH18 | 1 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 |
|  | 2 | 2 | 1 | 1 | 1 | 1 | 1 | 1 | 0 | 0 |
|  | 1.195 | Inf | 2.134 | Inf | 1.333 | Inf | 2.067 | Inf | 2.708 | Inf |
| MLTMNH18 | 5 | 1 | 1 | 0 | 1 | 0 | 2 | 1 | 0 | 0 |
|  | 3 | 5 | 1 | 2 | 1 | 2 | 2 | 2 | 1 | 1 |
|  | 0.580 | 4.712 | 0.842 | Inf | 0.976 | Inf | 1.126 | 2.236 | 2.121 | Inf |
| $A V E R V(\cdot)$ | 0.338 | Inf | 0.748 | Inf | 0.622 | Inf | 0.596 | Inf | 1.265 | Inf |
| MEDRV( $\cdot$ ) | 0.193 | 0.246 | 0.453 | 0.466 | 0.546 | 0.759 | 0.352 | 0.589 | 0.800 | 0.702 |

[^3]
## II.8. CONCLUDING REMARKS FOR PART II

For completeness, our first general remark is copied from our earlier report [5].

## General Remark 1: Differential Privacy, TDA, and $\epsilon$

Our objective in Part II of this study has been to report on the level of variability in results from the $T D A$ and to reveal any effects on variability given advances with the $T D A$ and an increased $\epsilon$ to 17.14. Our intent has not been to discuss how the TDA is constructed or how it operates. However, we feel compelled to offer a few such comments in this general remark, though our knowledge and understanding about the TDA is limited [1], [2].

The objective of the $T D A$ is to bring privacy protection to respondent data. There are three things to consider: ( $i$ ) a database (i.e., the 2010 CEF ); (ii) a query made to the database (e.g., the number of people with certain characteristics in the database); and (iii) a randomized data protection mechanism that gives differential privacy (i.e., a probability distribution which is a part of the TDA). As Dwork (2014) [2] notes, "On an intuitive level, the goal of differential privacy is to obscure the presence or absence of any individual (in a database), or small group of individuals, while at the same time preserving statistical utility."

With differential privacy, the degree of privacy protection is reported by a positive quantity $\epsilon$. Consider two different values of $\epsilon, \epsilon_{1}$ and $\epsilon_{2}$. If $\epsilon_{1}<\epsilon_{2}$, more privacy is offered with $\epsilon_{1}$ than with $\epsilon_{2}$. While details of the TDA and its foundation based on principles of differential privacy [4] are out-of-scope for this study (whose focus is only observing variability of output from the TDA), we note that the TDA has two components; and we share a little of our limited understanding. For simplicity, assume that an investigator is interested in knowing the count of persons in the 2010 CEF data with certain very specific characteristics. Thus a query is made of the 2010 CEF data (the answer sought should be a nonnegative integer). In the first component (noisy measurement) of the implementation of the $T D A$, random noise is generated and added to the answer from our query of the 2010 CEF data. The source of the random noise is a probability distribution (differentially private random mechanism) with positive probability at each of the integers ...-3, $-2,-1,0,1,2$, $3, \ldots$ Thus the "noisy answer" that is to be returned to the investigator submitting the query is
"noisy answer" $=$ (the query's answer using 2010 CEF data $)+$ (random noise which is an integer).
However, if the random noise is a negative integer whose absolute value is greater than the query's answer using the 2010 CEF data, then our noisy answer would be a negative noisy answer, which is not feasible. Thus, action is needed. This is the purpose of component two (post-processing) of the $T D A$, to ensure that our "final noisy answer" to the query is a nonnegative integer. So some more work is needed before the investigator eventually gets a "final noisy answer" to the original query.

Statistical theory permits deep explicit understanding of the variability caused by generation of the random noise in the first component. In particular, if $\epsilon_{1}<\epsilon_{2}$, the variability in the noise addition with $\epsilon_{1}$ is more than the variability in the noise with $\epsilon_{2}$. The variability and uncertainty due to the activity of the second component is less well understood by us, and we believe it currently contributes more variability and uncertainty than the first component for some queries. We believe that the empirical variability reported in this study is an overall combination of variability and uncertainty from the two components.

Figure 2


## General Remark 2: Effects on Variability Due to Advances with TDA and Higher $\epsilon$.

We observed reductions in variability from the 2021-04-28 version of the TDA with $\epsilon=10.3$ to the 2020 Census redistricting data production settings version of the $T D A$ with $\epsilon=17.14$. We see this visually by comparing Tables $7 \mathrm{~V}, 8 \mathrm{~V}, 9 \mathrm{~V}, 10 \mathrm{~V}, 11 \mathrm{~V}$, and 12 V of this study with Tables 7 V , $8 \mathrm{~V}, 9 \mathrm{~V}, 10 \mathrm{~V}, 11 \mathrm{~V}$, and 12 V of our earlier study [6]. At a high level, Figure 2 shows $\operatorname{AVERV}(1)$
values for each of the districts as shown in Figure 1 using the 2021-04-28 version and the 2020 Census redistricting data production settings version of the TDA. The $\operatorname{AVERV}(1)$ values for the 2020 Census redistricting data production settings version tend to be lower than for the 2021-04-28 version.

## General Remark 3: Repeat of Some Earlier Specific Remarks [5, 6]

In this remark, we repeat two specific remarks (slightly edited) made in our earlier study [6]:
Need for Better Understanding of the TDA: The output of the version of the TDA studied in this paper infuses noise via differentially private mechanisms with a total privacy-loss budget of $\epsilon=17.14$. It then post-processes those noisy estimates into fully consistent non-negative, integervalued data with the same schema as was produced in 2010. The observation that $R V(2)_{g}>R V(1)_{g}$ (also $\sqrt{V(2)_{g}}>\sqrt{\left.V(1)_{g}\right)}$ in the majority of the variation tables may be a reflection of some phenomenon like a bias caused by post-processing. If there is something like bias, it is relative to the official (swapping) counts from the 2010 Census and not necessarily relative to the unknown true counts. A stronger understanding of the cumulative effects of the noise infusion and postprocessing, as they affect jurisdictions with smaller populations, would be beneficial. This is a topic for further study.

Study Limitation: This study is limited in that new data (TDA) was retrofitted into existing redistricting plans developed using similar, but different data (2010 Census) treated by swapping. In practice, redistricting plans would be drawn using one set of data to satisfy desired parameters. In Congressional redistricting, for instance, DEV would not exceed 1 for any district, by design.

## REFERENCES

[1] Abowd, J., Ashmead, R., Garfinkel, S., Kifer, D., Leclerc, P., Machanavajjhala, A., Moran, B., Sexton, W., and Zhuravlev, P. (October 2019 Draft) U.S. Bureau of the Census, Washington, D.C. < https://github.com/uscensusbureau/census2020-das-2010ddp/blob/master/doc/20191020_1843_Consistency_ for_Large_Scale_Differentially_Private_Histograms.pdf >
[2] Dwork, C. (2014). "Differential Privacy: A Cryptographic Approach to Private Data Analysis," in Privacy, Big data, and the Public Good, (Editors: J. Lane, V. Stodden, S. Bender, and H. Nissenbaum), New York, NY: Cambridge University Press, 296-322.
[3] Table P2 HISPANIC OR LATINO, AND NOT HISPANIC OR LATINO BY RACE, Universe: Total population, 2010 Census Redistricting Data (Public Law 94-171) Summary File Also Table P4 HISPANIC OR LATINO, AND NOT HISPANIC OR LATINO BY RACE FOR THE POPULATION 18 YEARS AND OVER, Universe: Total population 18 years and over, 2010 Census Redistricting Data (Public Law 94-171) Summary File, American FactFinder, U. S. Bureau of the Census, Washington, D.C.
[4] Table P9 HISPANIC OR LATINO, AND NOT HISPANIC OR LATINO BY RACE, Universe: Total population, 2010 Census Congressional District Summary File ( $113^{\text {th }}$ Congress) Also Table P11 HISPANIC OR LATINO, AND NOT HISPANIC OR LATINO BY RACE FOR THE POPULATION 18 YEARS AND OVER, Universe: Population 18 years and over, 2010 Census Congressional District Summary File ( $113^{t h}$ Congress), American FactFinder, U. S. Bureau of the Census, Washington, D.C.
[5] Wright, T. and Irimata, K. (2020). "Variability Assessment of Data Treated by the TopDown Algorithm for Redistricting," Study Series (Statistics \#2020-02), Center for Statistical Research and Methodology, U.S. Bureau of the Census, Washington, D.C.
[6] Wright, T. and Irimata, K. (2021). "Empirical Study of Two Aspects of The TopDown Algorithm Output for Redistricting: Reliability \& Variability", Study Series (Statistics \#2021-01), Center for Statistical Research and Methodology, U.S. Bureau of the Census, Washington, D.C.
[7] Zayatz, L. (2007). "Disclosure Avoidance Practices and Research at the U.S. Census Bureau: An Update", Journal of Official Statistics, Vol 21, No. 2, 253-265.
[8] Voting Rights Act of 1965, $110^{\text {th }}$ Public law, $89^{\text {th }}$ Congress, U.S. Statutes at Large, Vol 79, starts p 437.
[9] Thornburg v Gingles (1986), U.S. Supreme Court, Vol 478, U.S. 30.

## APPENDIX A. Data Dictionary for Demographic Groups

| DIST-ID: | Identification for geographical area: e.g., congressional or state legislative, county, or state |
| :--- | :--- |
| TOTAL: | Total population |
| DEV: | Deviation from Ideal $=$ TOTAL - (IDEAL POPULATION) |
| DEVP: | Percent deviation from Ideal $=$ [DEV/(IDEAL POPULATION)] $100 \%$ |
| TOTAL18 | All individuals 18 years of age or older |
| TOTALHISP: | All individuals of any race and who chose Hispanic |
| TOTALHISPP: | [TOTALHISP/TOTAL] $\times 100 \%$ |
| TOTALNH: | All individuals of any race and who chose Not Hispanic |
| TOTALNHP: | [TOTALNH/TOTAL] $\times 100 \%$ |

## APPENDIX B. Computation Illustration for Measures of Variation in Table 7V

For the demographic group $g=A S I A N N H$ of CD-01 in Table 7V, we illustrate the computations for $\bar{C}_{T}(g), \sqrt{V(1)_{g}}, R V(1)_{g}, C_{S}(g), \sqrt{V(2)_{g}}$, and $R V(2)_{g}$ which are all defined in Section II. 7 of this report. The same details follow for all other demographic groups as well as all entries in Tables 7V; 8V; 9V; 10V; 11V; and 12V. From the 2010 Census (swapping), Table 7 gives $C_{S}(g)=17,705$ There are $25 T D A$ runs, and the details for the $i^{\text {th }}$ run are given on row $i$ of the table below for $i=1,2, \ldots, 25$.

| Run $i$ | $C_{T i}(g)$ | $\left(C_{T i}(g)-C_{T}(g)\right)^{2}$ | $\left(C_{T i}(g)-C_{S}(g)\right)^{2}$ |
| ---: | ---: | ---: | ---: |
| 1. | 17,709 | $(17,709-17,680.24)^{2}=827.14$ | $(17,709-17,705)^{2}=16$ |
| 2. | 17,680 | $(17,680-17,680.24)^{2}=0.06$ | $(17,680-17,705)^{2}=625$ |
| 3. | 17,668 | $(17,668-17,680.24)^{2}=149.82$ | $(17,668-17,705)^{2}=1,369$ |
| 4. | 17,678 | $(17,678-17,680.24)^{2}=5.02$ | $(17,678-17,705)^{2}=729$ |
| 5. | 17,693 | $(17,693-17,680.24)^{2}=162.82$ | $(17,693-17,705)^{2}=144$ |
| 6. | 17,671 | $(17,671-17,680.24)^{2}=85.38$ | $(17,671-17,705)^{2}=1,156$ |
| 7. | 17,664 | $(17,664-17,680.24)^{2}=263.74$ | $(17,664-17,705)^{2}=1,681$ |
| 8. | 17,662 | $(17,662-17,680.24)^{2}=332.70$ | $(17,662-17,705)^{2}=1,849$ |
| 9. | 17,692 | $(17,692-17,680.24)^{2}=138.30$ | $(17,692-17,705)^{2}=169$ |
| 10. | 17,690 | $(17,690-17,680.24)^{2}=95.26$ | $(17,690-17,705)^{2}=225$ |
| 11. | 17,646 | $(17,646-17,680.24)^{2}=1,172.38$ | $(17,646-17,705)^{2}=3,481$ |
| 12. | 17,707 | $(17,707-17,680.24)^{2}=716.10$ | $(17,707-17,705)^{2}=4$ |
| 13. | 17,642 | $(17,642-17,680.24)^{2}=1,462.30$ | $(17,642-17,705)^{2}=3,969$ |
| 14. | 17,666 | $(17,666-17,680.24)^{2}=202.78$ | $(17,666-17,705)^{2}=1,521$ |
| 15. | 17,705 | $(17,705-17,680.24)^{2}=613.06$ | $(17,705-17,705)^{2}=0$ |
| 16. | 17,681 | $(17,681-17,680.24)^{2}=0.58$ | $(17,681-17,705)^{2}=576$ |
| 17. | 17,704 | $(17,704-17,680.24)^{2}=564.54$ | $(17,704-17,705)^{2}=1$ |
| 18. | 17,676 | $(17,676-17,680.24)^{2}=17.98$ | $(17,676-17,705)^{2}=841$ |
| 19. | 17,667 | $(17,667-17,680.24)^{2}=175.30$ | $(17,667-17,705)^{2}=1,444$ |
| 20. | 17,690 | $(17,690-17,680.24)^{2}=95.26$ | $(17,690-17,705)^{2}=225$ |
| 21. | 17,645 | $(17,645-17,680.24)^{2}=1,241.86$ | $(17,645-17,705)^{2}=3,600$ |
| 22. | 17,689 | $(17,689-17,680.24)^{2}=76.74$ | $(17,689-17,705)^{2}=256$ |
| 23. | 17,682 | $(17,682-17,680.24)^{2}=3.10$ | $(17,682-17,705)^{2}=529$ |
| 24. | 17,722 | $(17,722-17,680.24)^{2}=1,743.90$ | $(17,722-17,705)^{2}=289$ |
| 25. | 17,677 | $(17,677-17,680.24)^{2}=10.50$ | $(17,677-17,705)^{2}=784$ |
| Totals | 442,006 | $10,156.56$ |  |

Thus we have (compare with corresponding entries of Table 7 V ):

$$
\begin{array}{l|l}
\bar{C}_{T}(g)=\frac{442,006}{25}=17,680.24 \approx \mathbf{1 7}, \mathbf{6 8 0} & C_{S}(g)=\mathbf{1 7}, \mathbf{7 0 5} \\
\sqrt{V(1)_{g}}=\sqrt{\frac{10,156.56}{25}}=20.16 \approx \mathbf{2 0} & \sqrt{V(2)_{g}}=\sqrt{\frac{25,483}{25}}=31.93 \approx \mathbf{3 2} \\
R V(1)_{g}=\frac{\sqrt{V(1)_{g}}}{\bar{C}_{T}(g)}=0.00114 \approx \mathbf{0 . 0 0 1} & R V(2)_{g}=\frac{\sqrt{V(2)_{g}}}{C_{S}(g)}=0.00180 \approx \mathbf{0 . 0 0 2}
\end{array}
$$

## APPENDIX C. Determination of $C_{S W A}^{*}$ Using 18 and Over Characteristics

As an alternative to the results in Table 3, Table $3^{\prime}$ below reveals an empirical answer to our question where we use TOTAL18 demographic groups in place of TOTAL demographic groups. More specifically, we use TOTAL18, HISP18, WHITENH18, BLACKNH18, AIANNH18, ASIANNH18, and HPINH18 in place of TOTAL, HISP18, WHITENH, BLACKNH, AIANNH, ASIANNH, and HPINH, respectively.

Table 3': Proportion of Block Groups in Each Stratum for Three Criteria
(Proportion computations result from 2020 Census Redistricting Data Production Settings ( $\epsilon=17.14$ for persons) version of $T D A$.)
Population: United States (50 States \& DC)

|  |  | Reliable Characteristics Criteria |  |  |
| :---: | :---: | :---: | :---: | :---: |
| Stratum for Block Groups Using $C_{S W A}$ for TOTAL | Number of Block Groups | Criterion I $\text { LDG } D R_{g} \leq 0.01$ | Criterion II $\text { LDG } D R_{g} \leq 0.03$ | Criterion III <br> LDG $D R_{g} \leq 0.05$ |
| $50 \leq C_{S W A} \leq 99$ | 128 | 0.1406 | 0.3984 | 0.5469 |
| $100 \leq C_{S W A} \leq 149$ | 99 | 0.2020 | 0.4747 | 0.6768 |
| $150 \leq C_{S W A} \leq 199$ | 124 | 0.2177 | 0.5081 | 0.7177 |
| $200 \leq C_{S W A} \leq 249$ | 154 | 0.2792 | 0.5974 | 0.8052 |
| $250 \leq C_{S W A} \leq 299$ | 209 | 0.2775 | 0.6316 | 0.8325 |
| $300 \leq C_{S W A} \leq 349$ | 264 | 0.3485 | 0.7652 | 0.9205 |
| $350 \leq C_{S W A} \leq 399$ | 407 | 0.3587 | 0.7838 | 0.9189 |
| $400 \leq C_{S W A} \leq 449$ | 569 | 0.4130 | 0.7926 | 0.9279 |
| $450 \leq C_{S W A} \leq 499$ | 915 | 0.3934 | 0.8372 | 0.9486 |
| $500 \leq C_{S W A} \leq 549$ | 1,699 | 0.4473 | 0.8723 | 0.9670 |
| $550 \leq C_{S W A} \leq 599$ | 3,238 | 0.4682 | 0.8799 | 0.9710 |
| $600 \leq C_{S W A} \leq 649$ | 5,131 | 0.4835 | 0.8953 | 0.9745 |
| $650 \leq C_{S W A} \leq 699$ | 6,683 | 0.4839 | 0.9051 | 0.9791 |
| $700 \leq C_{S W A} \leq 749$ | 7,356 | 0.5103 | 0.9226 | 0.9844 |
| $750 \leq C_{S W A} \leq 799$ | 8,170 | 0.5274 | 0.9263 | 0.9836 |
| $800 \leq C_{S W A} \leq 849$ | 8,213 | 0.5418 | 0.9364 | 0.9864 |
| $850 \leq C_{S W A} \leq 899$ | 8,441 | 0.5664 | 0.9494 | 0.9884 |
| $900 \leq C_{S W A} \leq 949$ | 8,657 | 0.5734 | 0.9485 | 0.9886 |
| $950 \leq C_{S W A} \leq 999$ | 8,723 | 0.5866 | 0.9555 | 0.9883 |
| $1,000 \leq C_{S W A} \leq 1,049$ | 8,398 | 0.6162 | 0.9640 | 0.9887 |
| $1,050 \leq C_{S W A} \leq 1,099$ | 8,345 | 0.6258 | 0.9685 | 0.9887 |
| $1,100 \leq C_{S W A} \leq 1,149$ | 7,950 | 0.6360 | 0.9707 | 0.9896 |
| $1,150 \leq C_{S W A} \leq 1,199$ | 7,860 | 0.6529 | 0.9726 | 0.9898 |
| $1,200 \leq C_{S W A} \leq 1,249$ | 7,451 | 0.6642 | 0.9764 | 0.9906 |
| $1,250 \leq C_{S W A} \leq 1,299$ | 7,124 | 0.6683 | 0.9746 | 0.9883 |
| $1,300 \leq C_{S W A} \leq 1,349$ | 6,714 | 0.6845 | 0.9768 | 0.9885 |
| $1,350 \leq C_{S W A} \leq 1,399$ | 6,507 | 0.6968 | 0.9825 | 0.9900 |
| $1,400 \leq C_{S W A} \leq 1,449$ | 5,911 | 0.7195 | 0.9831 | 0.9888 |
| $1,450 \leq C_{S W A} \leq 1,499$ | 5,617 | 0.7216 | 0.9838 | 0.9909 |
| $1,500 \leq C_{S W A} \leq 1,549$ | 5,390 | 0.7343 | 0.9818 | 0.9866 |
| $1,550 \leq C_{S W A} \leq 1,599$ | 4,856 | 0.7381 | 0.9811 | 0.9889 |
| $1,600 \leq C_{S W A} \leq 1,649$ | 4,508 | 0.7540 | 0.9843 | 0.9891 |
| $1,650 \leq C_{S W A} \leq 1,699$ | 4,325 | 0.7612 | 0.9817 | 0.9875 |
| $1,700 \leq C_{S W A} \leq 1,749$ | 4,093 | 0.7794 | 0.9863 | 0.9895 |
| $1,750 \leq C_{S W A} \leq 1,799$ | 3,689 | 0.7742 | 0.9883 | 0.9905 |
| $1,800 \leq C_{S W A} \leq 1,849$ | 3,469 | 0.7737 | 0.9853 | 0.9888 |
| $1,850 \leq C_{S W A} \leq 1,899$ | 3,252 | 0.7949 | 0.9865 | 0.9883 |
| $1,900 \leq C_{S W A} \leq 1,949$ | 3,008 | 0.7896 | 0.9860 | 0.9880 |
| $1,950 \leq C_{S W A} \leq 1,999$ | 2,832 | 0.8030 | 0.9887 | 0.9894 |
| $2,000 \leq C_{S W A} \leq 2,049$ | 2,573 | 0.8232 | 0.9852 | 0.9872 |
| $2,050 \leq C_{S W A} \leq 2,099$ | 2,356 | 0.8226 | 0.9881 | 0.9907 |
| $2,100 \leq C_{S W A} \leq 2,149$ | 2,307 | 0.8279 | 0.9896 | 0.9918 |
| $2,150 \leq C_{S W A} \leq 2,199$ | 2,033 | 0.8224 | 0.9803 | 0.9813 |
| $2,200 \leq C_{S W A} \leq 2,249$ | 1,999 | 0.8369 | 0.9865 | 0.9870 |
| $2,250 \leq C_{S W A} \leq 2,299$ | 1,892 | 0.8451 | 0.9884 | 0.9884 |
| $2,300 \leq C_{S W A} \leq 2,349$ | 1,666 | 0.8487 | 0.9862 | 0.9862 |
| $2,350 \leq C_{S W A} \leq 2,399$ | 1,622 | 0.8539 | 0.9846 | 0.9846 |
| $2,400 \leq C_{S W A} \leq 2,449$ | 1,421 | 0.8656 | 0.9880 | 0.9887 |
| $2,450 \leq C_{S W A} \leq 2,499$ | 1,350 | 0.8793 | 0.9926 | 0.9926 |
| Total | 199,698 |  |  |  |

Using Criterion II and searching from top to bottom for the first stratum whose proportion is at least 0.9500: From Table $3^{\prime}$, take $C_{S W A}^{*}$ to be between 950 and 999 . For block groups whose TOTAL $C_{S W A}$ count is at least 999, the difference of ratios between the $C_{T D A}$ and $C_{S W A}$ ratios for the LDG will tend to be less than or equal to $3 \%$ (using our data).

Using Criterion III and searching from top to bottom for the first stratum whose proportion is at least 0.9500: From Table $3^{\prime}$, take $C_{S W A}^{*}$ to be between 500 and 549. For block groups whose TOTAL $C_{S W A}$ count is at least 549 , the difference of ratios between the $C_{T D A}$ and $C_{S W A}$ ratios for the LDG will tend to be less than or equal to $5 \%$ (using our data).

Using the data that will be released to the public (one run of the 2020 Census redistricting data production settings version of $T D A$ ), we might say, empirically based on the data for the block groups used in our study, that
"for any block group with a TOTAL count between 500 and 549 people, the difference between the TDA ratio of the largest demographic group ( $L D G$ ) and the corresponding SWA ratio for the $L D G$ among the 18 years and over pouplation is less than or equal to 5 percentage points at least $95 \%$ of the time".
We applied the same version of the TDA to the same underlying CEF data 25 independent times, i.e., for 25 additional runs focusing on the 18 years and over population. The stratum (using $C_{S W A}$ for TOTAL) for each run, where we first observed that 0.9500 was exceeded is given in Table $3 a^{\prime}$ for each run is between 450 and 499 people in 23 of the 25 runs.

Table 3a': For Each Run, the Stratum and Stratum Proportion When 0.9500 First Exceeded (Proportion computations result from 2020 Census Redistricting Data Production Settings ( $\epsilon=17.14$ for persons) version of $T D A$.) Population: United States (50 States \& DC)

|  |  | Criterion III <br> LDG $D R_{g} \leq 0.05$ |
| ---: | :---: | :---: |
|  |  |  |
|  |  |  |
|  | Stratum for | Proportion When |
|  | Block Groups | 0.9500 First Exceeded |
|  |  |  |
| 1 | $450 \leq C_{S W A} \leq 499$ | 0.9552 |
| 2 | $450 \leq C_{S W A} \leq 499$ | 0.9617 |
| 3 | $450 \leq C_{S W A} \leq 499$ | 0.9607 |
| 4 | $450 \leq C_{S W A} \leq 499$ | 0.9650 |
| 5 | $450 \leq C_{S W A} \leq 499$ | 0.9541 |
| 6 | $450 \leq C_{S W A} \leq 499$ | 0.9519 |
| 7 | $450 \leq C_{S W A} \leq 499$ | 0.9639 |
| 8 | $450 \leq C_{S W A} \leq 499$ | 0.9574 |
| 9 | $450 \leq C_{S W A} \leq 499$ | 0.9683 |
| 10 | $400 \leq C_{S W A} \leq 449$ | 0.9525 |
| 11 | $450 \leq C_{S W A} \leq 499$ | 0.9563 |
| 12 | $450 \leq C_{S W A} \leq 499$ | 0.9639 |
| 13 | $450 \leq C_{S W A} \leq 499$ | 0.9650 |
| 14 | $450 \leq C_{S W A} \leq 499$ | 0.9661 |
| 15 | $450 \leq C_{S W A} \leq 499$ | 0.9596 |
| 16 | $450 \leq C_{S W A} \leq 499$ | 0.9519 |
| 17 | $450 \leq C_{S W A} \leq 499$ | 0.9650 |
| 18 | $400 \leq C_{S W A} \leq 449$ | 0.9543 |
| 19 | $450 \leq C_{S W A} \leq 499$ | 0.9672 |
| 20 | $450 \leq C_{S W A} \leq 499$ | 0.9552 |
| 21 | $450 \leq C_{S W A} \leq 499$ | 0.9552 |
| 22 | $450 \leq C_{S W A} \leq 499$ | 0.9661 |
| 23 | $450 \leq C_{S W A} \leq 499$ | 0.9639 |
| 24 | $450 \leq C_{S W A} \leq 499$ | 0.9585 |
| 25 | $450 \leq C_{S W A} \leq 499$ | 0.9639 |
|  |  |  |
|  |  |  |
|  |  |  |
|  |  |  |


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[^1]:    ${ }^{a}$ The ideal population for each of $K$ districts of a jurisdiction is the jurisdiction's total population divided by $K$.

[^2]:    Source: Data from 25 Runs of the TDA, U. S. Bureau of the Census, Washington, D.C.

[^3]:    Source: Data from 25 Runs of the $T D A$, U. S. Bureau of the Census, Washington, D.C.

