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A Set of End Weights to End All End Weights

Given a moving average used to calculate a seasonal factor curve or a trend-cycle curve, the problem is to derive a set of moving average weights for the end terms which will minimize revisions between preliminary and final estimates of the seasonal or trend-cycle. As an example, consider the end weights for the 5-term moving average as a seasonal factor curve. End weights for other seasonal curves of different length and trend-cycle curves may be derived in a similar manner.

Let X_1, \dots, X_5 represent the latest five S-I ratios (where X_5 represents the end year value). Similarly, let S_1, \dots, S_5 represent the "true" seasonal factors corresponding to X_1, \dots, X_5 . Assume that $X_i = S_i + I_i$, where I_i is the irregular corresponding to X_i . Further assume that the I_i 's are independent with equal variance σ^2 .

Define:

$$S'_4 = W_1 X_2 + W_2 X_3 + W_3 X_4 + W_2 X_5 + W_1 X_6,$$
$$S'_5 = W_1 X_3 + W_2 X_4 + W_3 X_5 + W_2 X_6 + W_1 X_7,$$

where S'_4 and S'_5 are the ultimate estimates of S_4 and S_5 . As the preliminary estimates of S_4 and S_5 when data is available through X_5 , define:

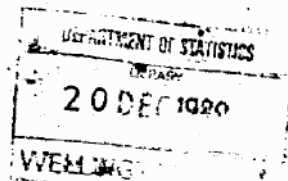
$$S_4'' = U_1 X_2 + U_2 X_3 + U_3 X_4 + U_4 X_5,$$
$$S_5'' = V_1 X_3 + V_2 X_4 + V_3 X_5.$$

We want to minimize

$$R_1 = E(S_4'' - S'_4)^2 \text{ and } R_2 = E(S_5'' - S'_5)^2$$

with respect to the U_i and V_i subject to the constraints

$$\sum_1^4 U_i = \sum_1^3 V_i = 1. \text{ (Also, } \sum_1^5 W_i = 1.)$$



$$\begin{aligned} \text{Now } R_1 &= \sigma \left(S_4^{n2} - S_4' \right) + B \left(S_4^{n2} - S_4' \right) \\ &= \left[(U_1 - W_1)^2 + (U_2 - W_2)^2 + (U_3 - W_3)^2 + (U_4 - W_2)^2 + W_1^2 \right] \sigma I^2 \\ &+ \left[(U_1 - W_1)S_2 + (U_2 - W_2)S_3 + (U_3 - W_3)S_4 + (U_4 - W_2)S_5 - W_1 S_6 \right]^2, \\ \text{where } B \left(S_4^{n2} - S_4' \right) &= E(S_4^{n2}) - E(S_4'). \end{aligned}$$

Clearly, it is an easy matter to minimize $\sigma S_4^{n2} - S_4'$ with respect to the U_i and obtain estimates of the U_i in terms of the W_i . To minimize $B S_4^{n2} - S_4'$, however, it is necessary to make the further assumption that $S_{t+1} - S_t = \Delta_S$, where Δ_S is a constant. Hence, define $S_2 = S$, $S_3 = S + \Delta_S$, ..., $S_6 = S + 4\Delta_S$.

To minimize R_1 , form $F = R_1 - \lambda \left[\sum_{i=1}^4 U_i - 1 \right]$, where λ is a Lagrange multiplier. Differentiating, we have:

$$(1) \frac{\partial F}{\partial U_1} = 2(U_1 - W_1) \sigma I^2 - \lambda + 2S_2 B = 0,$$

$$(2) \frac{\partial F}{\partial U_2} = 2(U_2 - W_2) \sigma I^2 - \lambda + 2S_3 B = 0,$$

$$(3) \frac{\partial F}{\partial U_3} = 2(U_3 - W_3) \sigma I^2 - \lambda + 2S_4 B = 0,$$

$$(4) \frac{\partial F}{\partial U_4} = 2(U_4 - W_2) \sigma I^2 - \lambda + 2S_5 B = 0.$$

$$\text{Summing: } 2 \left[1 - (1 - W_1) \right] \sigma I^2 - 4\lambda + 2B \sum_{i=1}^5 S_i = 0;$$

$$\lambda = \frac{W_1 \sigma I^2}{2} + \frac{B \sum_{i=1}^5 S_i}{2}.$$

$$\begin{aligned} \text{Now } U_1 &= W_1 + \frac{\lambda}{2\sigma I^2} - \frac{S_2 B}{\sigma I^2} \\ &= W_1 + \frac{W_1}{4} + \frac{B \sum_{i=1}^5 S_i}{4\sigma I^2} - \frac{B S_2}{\sigma I^2} \\ &= W_1 + \frac{W_1}{4} + \frac{B}{4\sigma I^2} \left[4S + 6\Delta_S - 4S \right]. \end{aligned}$$

$$\begin{aligned}
 (1') \quad U_1 &= W_1 + \frac{W_1}{4} + \frac{3\Delta S}{2\sigma I^2} \quad B \\
 &= W_1 + \frac{W_1}{4} + \frac{3\Delta S}{2\sigma I^2} \left[\sigma \left(\frac{4}{1} U_1 - \frac{5}{1} W_1 \right) + \Delta S \left[U_2 + 2U_3 + 3U_4 - \right. \right. \\
 &\quad \left. \left. (W_1 + 2W_3 + 3W_2 + 4W_1) \right] \right] \\
 &= W_1 + \frac{W_1}{4} + \frac{3\Delta S^2}{2\sigma I^2} \left[2 - 2U_1 - U_2 + U_4 - 2 \right] \\
 &= W_1 + \frac{W_1}{4} + \frac{3\Delta S^2}{2\sigma I^2} \left[-2U_1 - U_2 + U_4 \right]
 \end{aligned}$$

$$(1'') \quad \left(1 + 3 \frac{\Delta S^2}{\sigma I^2} \right) U_1 + \frac{3}{2} \frac{\Delta S^2}{\sigma I^2} U_2 + 0 \cdot U_3 - \frac{3}{2} \frac{\Delta S^2}{\sigma I^2} U_4 = W_1 + \frac{W_1}{4}.$$

It is necessary at this point to derive a functional relationship between $\frac{\Delta S^2}{\sigma I^2}$ and \bar{I}'/\bar{S}' , since the W_1 are based on the size of \bar{I}'/\bar{S}' .

$$\begin{aligned}
 \bar{I}'^2 &= \frac{2}{\pi} \sigma \delta I^2 \left(\delta I = \frac{I_{t+1} - I_t}{I_t} \text{ -- this approximation due to Rosenblatt} \right) \\
 &= \frac{2}{\pi} \left[\frac{\sigma I_{t+1}^2}{E(I_{t+1})^2} + \frac{\sigma I_t^2}{E(I_t)^2} - \frac{2 \sigma I_{t+1} I_t}{E(I_{t+1})E(I_t)} \right] \\
 &= \frac{2}{\pi} \left[\frac{2 \sigma I^2}{(100)^2} \right] \\
 &= \frac{4}{\pi} \left[\frac{\sigma I^2}{(100)^2} \right] \\
 \sigma I^2 &= \frac{\pi}{4} (100)^2 \bar{I}'^2 \\
 \bar{S}' &= \frac{1}{n-1} \sum_1^{n-1} \left| \frac{S_{t+1} - S_t}{S_t} \right| \quad (n = \text{number of years}) \\
 &= \frac{\Delta S}{n-1} \sum_1^{n-1} \frac{1}{S_t} \\
 \bar{S}'^2 &= \Delta S^2 \left[\frac{1}{n-1} \sum_1^{n-1} \frac{1}{S_t} \right]^2
 \end{aligned}$$

Clearly, the above analysis holds only if Δ_S is constant for the entire historical period. Let $\tilde{S} = \frac{1}{n-1} \sum_{t=1}^{n-1} S_t$. If all $S_t = \hat{S}$, $\tilde{S} = \frac{n-1}{\hat{S}}$.

$$\text{Hence, } \bar{S}'^2 = \Delta_S^2 \left(\frac{\tilde{S}}{n-1}\right)^2 \text{ and } \Delta_S^2 = \frac{(n-1)^2 \tilde{S}^2}{\tilde{S}^2}. \text{ When all } S_t = \hat{S}, \\ \Delta_S^2 = \bar{S}'^2 \cdot \tilde{S}^2.$$

$$\text{Then } \frac{\Delta_S^2}{\sigma_I^2} = \frac{(n-1)^2 \bar{S}'^2 / \tilde{S}^2}{\pi/4 (100)^2 \bar{I}'^2} \\ = \frac{4}{\pi} \frac{(n-1)^2}{(100)^2 \tilde{S}^2} \left(\frac{\bar{S}'}{\bar{I}'}\right)^2 = D.$$

In the special case where all $S_t = \hat{S}$,

$$D = \frac{\Delta_S^2}{\sigma_I^2} = \frac{4}{\pi} \frac{\hat{S}^2}{(100)^2} \left(\frac{\bar{S}'}{\bar{I}'}\right)^2$$

In the special case where $\frac{\tilde{S}}{n-1} = \frac{1}{100}$,

$$D = \frac{\Delta_S^2}{\sigma_I^2} = \frac{4}{\pi} \left(\frac{\bar{S}'}{\bar{I}'}\right)^2.$$

Then we have:

$$(1') \quad U_1 = W_1 + \frac{W_1}{4} + \frac{3}{2} \frac{D}{\Delta_S} \cdot B,$$

$$(2') \quad U_2 = W_2 + \frac{W_1}{4} + \frac{1}{2} \frac{D}{\Delta_S} \cdot B,$$

$$(3') \quad U_3 = W_3 + \frac{W_1}{4} - \frac{1}{2} \frac{D}{\Delta_S} \cdot B,$$

$$(4') \quad U_4 = W_4 + \frac{W_1}{4} - \frac{3}{2} \frac{D}{\Delta_S} \cdot B;$$

or

$$(1'') \quad (1 + 3D) U_1 + \frac{3}{2} D U_2 - \frac{3}{2} D U_4 = W_1 + \frac{W_1}{4},$$

$$(2'') \quad D U_1 + (1 + \frac{D}{2}) U_2 - \frac{D}{2} U_4 = W_2 + \frac{W_1}{4},$$

$$(3^n) \quad -D U_1 - \frac{D}{2} U_2 + U_3 + \frac{D}{2} U_4 = W_2 + \frac{W_1}{4},$$

$$(4^n) \quad -3D U_1 - \frac{3}{2} D U_2 + (1 + \frac{3}{2} D) U_4 = W_2 + \frac{W_1}{4}.$$

Using matrix notation, $Au = X$; $u = A^{-1}X$, where

$$Au = \begin{bmatrix} (1 + 3D) & \frac{3}{2} D & 0 & -\frac{3}{2} D \\ D & (1 + \frac{D}{2}) & 0 & -\frac{D}{2} \\ -D & -\frac{D}{2} & 1 & \frac{D}{2} \\ -3D & -\frac{3}{2} D & 0 & (1 + \frac{3}{2} D) \end{bmatrix} \begin{bmatrix} U_1 \\ U_2 \\ U_3 \\ U_4 \end{bmatrix} = \begin{bmatrix} W_1 + \frac{W_1}{4} \\ W_2 + \frac{W_1}{4} \\ W_3 + \frac{W_1}{4} \\ W_4 + \frac{W_1}{4} \end{bmatrix} = X.$$

To solve for A^{-1} and U , use the DAM regression program.

Suppose $\Delta S = 0$; i.e., S_t is a stable seasonal. Then

$B = S \left[\sum_1^4 U_i - \sum_1^5 W_i \right] = 0$. Hence, the U_i are unbiased estimates of the W_i and are as follows:

$$\begin{aligned} U_1 &= W_1 + \frac{W_1}{4}, & U_3 &= W_3 + \frac{W_1}{4} \\ U_2 &= W_2 + \frac{W_1}{4}, & U_4 &= W_2 + \frac{W_1}{4}. \end{aligned}$$

The U_i derived from $(1^n) - (4^n)$ depend on:

- the central weights W_i ,
- the \bar{I}/\bar{S} ratio,
- the level of the S_t ,
- the assumption of a seasonal with a linear trend which is the same in the historical and current period.

Similarly, minimize $R_2 = \sqrt{S_5^{n^2} - S_5'} + B^2 S_5^n - S_5'$

$$= \sqrt{(v_1 - w_1)^2 + (v_2 - w_2)^2 + (v_3 - w_3)^2 + w_2^2 + w_1^2} \sigma I^2$$

$$+ \sqrt{(v_1 - w_1) S_3 + (v_2 - w_2) S_4 + (v_3 - w_3) S_5 - w_2 S_6 - w_1 S_7}^2.$$

Let $F = R_2 - \lambda \left[\sum_1^3 v_1 - 1 \right]$. Differentiating,

$$(1) \frac{\partial F}{\partial v_1} = 2(v_1 - w_1) \sigma I^2 - \lambda + 2 S_3 B = 0,$$

$$(2) \frac{\partial F}{\partial v_2} = 2(v_2 - w_2) \sigma I^2 - \lambda + 2 S_4 B = 0,$$

$$(3) \frac{\partial F}{\partial v_3} = 2(v_3 - w_3) \sigma I^2 - \lambda + 2 S_5 B = 0.$$

Then $\lambda = \frac{2 \sigma I^2 (w_1 + w_2)}{3} + \frac{2 B \sum_3^5 S_i}{3}$.

Now $v_1 = w_1 + \frac{\lambda}{2 \sigma I^2} - \frac{S_3 B}{\sigma I^2}$

$$= w_1 + \frac{w_1 + w_2}{3} + \frac{B \sum_3^5 S_i}{3 \sigma I^2} - \frac{S_3 B}{\sigma I^2}$$

$$(1') \quad v_1 = w_1 + \frac{w_1 + w_2}{3} + \frac{\Delta S' B}{\sigma I^2}$$

$$= w_1 + \frac{w_1 + w_2}{3} + \frac{\Delta S}{\sigma I^2} \left[S \left(\frac{3}{1} v_1 - \frac{5}{1} w_1 \right) + \Delta S \left[\sqrt{v_2 + 2v_3 - (w_2 + 2w_3 + 3w_2 + 4w_1)} \right] \right]$$

$$(1'') \quad v_1 = w_1 + \frac{w_1 + w_2}{2} = \frac{\Delta^2 S}{\sigma I^2} (2v_1 + v_2).$$

Then we have:

$$(1') \quad V_1 = W_1 + \frac{W_1 + W_2}{3} + \frac{D}{\Delta s} \cdot B,$$

$$(2') \quad V_2 = W_2 + \frac{W_1 + W_2}{3},$$

$$(3') \quad V_3 = W_3 + \frac{W_1 + W_2}{3} - \frac{D}{\Delta s} \cdot B;$$

or

$$(1'') \quad (1 + 2D) V_1 = -D V_2 + W_1 + \frac{W_1 + W_2}{3},$$

$$(3'') \quad V_3 = 2D V_1 + D V_2 + W_3 + \frac{W_1 + W_2}{3}$$

(2'') is the same as (2'). Note that the middle weight in a pattern of end weights composed of an uneven number of terms depends only on the W_1 .